A Comparative Study on Inflated and Dispersed Count Data

Monika Arora, Yash Kalyani and Shivam Shanker

Department of Mathematics, Indraprastha Institute of Information Technology, Delhi, India

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Abstract: The availability of zero inflated count data has led to the demonstration of various statistical models and machine learning algorithms to be applied in diverse fields such as healthcare, economics and travel. However, in real life there could be a count k > 0 that is inflated. There are only a few studies on k- inflated count models. To the best of our knowledge, there is no article that demonstrates the machine learning algorithms on such data sets. We apply existing k- inflated count models as well as machine learning algorithms on travel data to compare the prediction and fitness of the models and find the significant covariates. Our study shows that the k- inflated models provide a good fit to the data, however, the predictions from machine learning algorithms are superior. This study can be extended further to include other artificial neural network approaches on a larger data set.

1 INTRODUCTION

In many areas of study, a widely popular data type explored is the count data. The Poisson distribution is commonly used to study equi-dispersed count data. The data is equi-dispersed when mean is equal to the variance. In real life data sets, over-dispersion is a common occurrence. The negative binomial distribution is used when variance is more than mean or over-dispersed in the count data. It is possible that there is inflation at count zero for which the zero inflated distributions are widely used. In seminal paper, (Lambert, 1992) introduced zero inflated Poisson (ZIP) regression model for zero inflated count data. The ZIP regression models are applied in many areas like life science (Ridout et al., 1998), (Hall, 2000), travel (Lord et al., 2005), and economics (Gurmu and Trivedi, 1996), (Cameron and Trivedi, 2013).

In the presence of over-dispersion in zero inflated count data, zero inflated negative binomial (ZINB) model is more appropriate and it was first studied by (Greene, 1994). Instead, of high frequency of zero it is possible there is any count value k > 0 that is inflated. The more appropriate choice for such data sets is k-inflated count models. An extension of ZIP distribution is k-inflated Poisson (kIP) distribution. The kIP model is a mixture of degenerate distribution at k with probability π and Poisson (λ) with probability (1 – π). It is a special case of zero and k inflated Poisson (ZkIP) distribution studied by (Lin and Tsai, 2012) and (Arora, 2018). The kIP model is also a special case of a k-inflated generalized Poisson distribution given by (Bae et al., 2005). The k-inflated analog of negative binomial is k-inflated negative binomial (kINB) distribution. Recently, (Payandeh Najafabadi and MohammadPour, 2018) studied kINB for rate making system. As compared to zero inflated count models, the literature on k-inflated count regression models is not so rich.

The count regression models allow us to study the the relationship between the response variables and covariates. They allows us to find significant covariates and make predictions. However, they have a few limitations. They do not allow to study the non-linear relationships of the covariates and are sensitive to outliers. There are various machine learning algorithms that enable us to build a more robust and complex regression models. They have good predictive abilities and are efficient. Recently, the approaches have been applied on zero inflated count data (see (Lee and Jin, 2006), (Arief and Murfi, 2018),(Alfredo et al., 2018)). Though, we did not come across any work that demonstrates the machine learning algorithms on k-inflated count data.

In this article, we consider the Poisson, NB and their k-inflated analog regression models. We study the fit and prediction of the models using training and test data, respectively. We perform 5-fold validation and implement machine learning algorithms to make the predictions. A comparative study between the

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regression models and machine learning algorithms show that the latter makes better predictions. The article is organized as follows. The Section 2 of the article describes statistical distributions and their corresponding regression models. It also describes the machine learning algorithms applied in the article. We demonstrate the methodology on a real life data set in Section 3. A detailed analysis is performed in Section 4. Lastly, Section 5 concludes the article.

2 METHODOLOGY

In this Section, we describe the methodologies used in this article. We present the statistical distributions and their regression models in Section 2.1 and Section 2.2, respectively. The measures of fit, described in Section 2.3, are used to find the best model. The machine learning algorithms are described in Section 2.4. The measures of prediction are given in Section 2.5.

2.1 Statistical Distributions

The Poisson and negative binomial are most commonly studied for count response variable. For zero inflated count data, their zero inflated analogs, zero inflated Poisson and zero inflated negative binomial are popular. While, for any count k >0 their k-inflated analogs, k-inflated Poisson and k-inflated negative binomial are more appropriate. In this Section we briefly describe Poisson, negative binomial and their k-inflated analogs.

Poisson Distribution. Let a response variable Y follows a Poisson distribution with unknown parameter $\lambda > 0$. The probability mass function (p.m.f.) of random variable Y is given by

$$P(Y = y) = \frac{e^{-\lambda} \lambda^{y}}{y!}; \ y = 0, 1, \dots$$
 (1)

The mean and variance of *Y* is λ . It is the equidispersion property of the distribution that makes it popular. The other commonly used property of the distribution is that it belongs to the exponential family. Therefore, it is easy to find the sufficient statistics and maximum likelihood estimates.

Negative Binomial Distribution. When the count response is over-dispersed (variance>mean), then negative binomial is a better choice. The p.m.f. of a response variable Y from negative binomial distribution is:

$$P(Y = y) = {\binom{y + \theta - 1}{y}} (1 - p)^{y} p^{\theta}; \ y = 0, 1, \dots,$$
(2)

where $\theta > 0$ and $0 \le p \le 1$.

k–Inflated Poisson Distribution. When zero is in excess in count response, then zero inflated Poisson (ZIP) distribution gives better fit to the data. In real life scenarios, any count k > 0 can be inflated. The analogue distributions for such data sets is k– inflated distributions. The *k*–inflated Poisson distribution (kIP) is a mixture of two distributions. One is degenerate at *k* with probability π and the other is Poisson (λ) with probability $(1 - \pi)$. The p.m.f. of a count response *Y* from kIP distribution is:

$$P(Y = y) = \begin{cases} \pi + (1 - \pi) \frac{e^{-\lambda} \lambda^{k}}{k!}, & y = k \\ (1 - \pi) \frac{e^{-\lambda} \lambda^{y}}{y!}, & y \ge 0, & y \ne k. \end{cases}$$
(3)

where $0 \le \pi \le 1$ and $\lambda > 0$. When $\pi = 0$, (3) reduces to Poisson distribution. The kIP distribution is a special case of zero and *k*-inflated Poisson (ZkIP) distribution studied by (Lin and Tsai, 2012) and (Arora, 2018).

k- Inflated Negative Binomial Distribution. To study the inflation at k > 0, the extension of NB distribution is k- inflated negative binomial (kINB) distribution. The p.m.f. of kINB is given by

$$P(Y=y) = \begin{cases} \pi + (1-\pi) {\binom{k+\theta-1}{k}} (1-p)^k p^{\theta}, & y=k \\ (1-\pi) {\binom{y+\theta-1}{y}} (1-p)^y p^{\theta}, & y \neq k. \end{cases}$$
(4)

On reparameterization, the variance of NB and kINB models in R and SAS is given by $Var(Y) = \lambda + r\lambda^2$; λ is the mean of response and *r* is the dispersion parameter. The high values of *r* corresponds to overdispersion in the data. When *r* is zero or close to zero, it signifies equi- or under-dispersion in the data.

2.2 Count Regression Models

To build the regression model, consider *n* independent count responses $\mathbf{Y} = (Y_1, \ldots, Y_n)$. The regression model corresponding to the Poisson distribution is popularly known as Poisson log linear regression model or simply a Poisson regression model. The Poisson regression model is a generalized linear model with a log link function. The Poisson regression model for the response vector \mathbf{Y} is given by

$$\log(\mathbf{\lambda}) = \beta_0 \mathbf{x_0} + \beta_1 \mathbf{x_1} + \ldots + \beta_p \mathbf{x_p}.$$
 (5)

where $\boldsymbol{\lambda} = (\lambda_1, \dots, \lambda_n)$ is the rate parameter of the model. The unknown parameters corresponding to the vectors of covariates, $\mathbf{x} = (\mathbf{x_0} = \mathbf{1}, \mathbf{x_1}, \dots, \mathbf{x_p})$ is $\boldsymbol{\beta} = (\beta_0, \beta_1, \dots, \beta_p)$. When the response variable follows a NB distribution then the NB regression model is most appropriate. The link function in the NB regression model can be obtained from (5) by replacing $\boldsymbol{\lambda}$ by the mean of NB distribution.

The kIP model is constructed using the ZIP regression model studied by (Lambert, 1992). The kIP model links the covariates to the rate parameter λ . The inflation parameter π is linked to an unknown constant using logit link. The model can be easily extended to associate the desired covariates to the inflation parameter π . The link functions are given by

$$\log(\boldsymbol{\lambda}) = \beta_0 \mathbf{x_0} + \beta_1 \mathbf{x_1} + \ldots + \beta_p \mathbf{x_p},$$

logit (\pi) = \gamma. (6)

Similarly, the kINB regression model uses the same link function as in kIP.

The Poisson, NB models and their k- inflated analogs are implemented in SAS using GENMOD and FMM procedures. The Poisson and NB models are implemented in R using glm and glm.nb functions, respectively. The kIP and kINB models are run in R using optimization routine 'nlminb'. The routine applies Newton-Raphson approach to obtain the estimates. The algorithms in SAS and R use maximum likelihood estimation approach to find the estimates of the unknown parameters. The results obtained in SAS and R are same.

2.3 Hypothesis Testing and Model Selection

In statistical modeling, hypothesis testing allows us to study the significance of the covariates. It also allows to select a good fit model. For the nested models, we compare the models using likelihood ratio test (LRT). Besides LRT, the popularly used measures to select a good model are Akaike Information Criterion (AIC), Bayesian Information Criterion (BIC) and R-squared. The AIC and BIC are based on log-likelihood function of the corresponding model. In this Section, we explain the hypothesis tests used to find significant covariates and unknown parameters in the Poisson, NB, kIP and kINB models. We also describe the thumb rule to choose best fit model using AIC and BIC.

Hypothesis Testing. We use hypothesis testing to study the significance of the parameters. The hypothesis $H_0: \beta_j = 0$ vs. $H_1: \beta_j \neq 0$ tests the significance of the unknown *j*th covariate. The Wald test statistic is $z = \hat{\beta}_j / SE(\hat{\beta}_j)$. For large sample and under the

null hypothesis, the statistic approximately follows a standard normal distribution.

The significance of the dispersion parameters in NB and kINB can also be tested. To test the dispersion parameter *r* in NB and kINB models we use $H_0: r = 0$ vs. $H_1: r \neq 0$. The corresponding test statistic is $z = \hat{r}/SE(\hat{r}) \sim N(0, 1)$.

The inflation parameter in the kIP and kINB models is $0 \le \pi \le 1$. To test the significance of inflation at k, we test $H_0: \pi = 0$ vs. $H_1: \pi > 0$. Under the null, the kIP and kINB models reduce to Poisson and NB, respectively. Hence, the base models (Poisson and NB) are nested in their k-analogs (kIP and kINB). For nested models, we can perform LRT. The test statistic is $-2\log \Lambda = -2(L_0 - L_1)$ where, L_0 is log-likelihood of the model under null while L_1 is log-likelihood under unrestricted. The test statistic follows a mixture of χ^2 distributions (Chant, 1974),(Shapiro, 1985). The distribution of the test statistic changes as $\pi = 0$ is a boundary point. Hence, the regularity conditions are not met.

Akaike Information Criterion (AIC). There are various measures that could be used to select the best model. One of the most popular measures is AIC. It is given by

$$AIC = -2\log L + 2p, \tag{7}$$

where $\log L$ is the log-likelihood of the model and p is the number of parameters in the model. The measure is preferred as it prevents over fitting of the data. To prevent over fitting it penalizes a model on adding more parameters. The model with minimum AIC gives the best fit to the data. The difference between AIC of the various models is subjective. (Burnham and Anderson, 2002) proposed a thumb rule for AIC difference to choose the best model. The AIC difference is given by

$$\Delta_i = AIC_i - min(AIC). \tag{8}$$

Here, AIC_i is AIC of *i*th model. The value of $0 \le \Delta_i \le 2$ provides substantial support, $4 \le \Delta_i \le 7$ implies considerably less support and $\Delta_i > 10$ essentially provides no support for *i*th model.

Bayesian Information Criterion (BIC). Another popular measure is the BIC. It is given by

$$BIC = -2\log L + p\log n, \tag{9}$$

where $\log L$ is the log-likelihood of the model, p is the number of parameters in the model, and n is the number of observations. It is also independent of the prior and penalizes the model on its complexity. The penalty term is larger in BIC than in AIC. While it has been commonly used for model identification in linear regression and time series, it can, however, also be applied to any set of maximum likelihood-based models (Schwarz, 1978). The model with minimum BIC gives the best fit to the data. As in the case of AIC, there also exists a rule of thumb for model selection based on BIC (Kass and Raftery, 1995).

$$\Delta_i = BIC_i - min(BIC). \tag{10}$$

Here, *BIC_i* is BIC of *i*th model. The value of Δ_i can be used as evidence against a candidate model to be the best model. For $0 \le \Delta_i \le 2$, it is not worth more than a bare mention, $2 < \Delta_i \le 6$, the evidence against the model is positive, $6 < \Delta_i \le 10$, the evidence against the model is strong, and $\Delta_i > 10$, the evidence against the model is very strong.

2.4 Machine Learning Algorithms

In this Section, we describe the machine learning algorithms applied in the study. The first algorithm we applied is decision trees. Decision Trees are one of the most commonly used approaches in machine learning. They help solve classification as well as regression problems. The algorithm constructs a tree by breaking the dataset into smaller subsets and incrementally building the tree node by node. They are easy to understand and interpret and require less effort for data preparation. Despite their advantages, they are prone to overfitting and instability because of their non-robustness. For categorical datasets, information gain in decision trees are sometimes biased to the attributes with more levels.

To obtain a better predictive performance than a single model, ensemble methods like bagging and boosting are typically used. Bagging (or Bootstrap Aggregating) involves training the base model on random subsets of the training dataset and then, aggregating their results by either voting or averaging to form a final prediction. An example of this ensemble method is random forests (RF), which uses this technique along with feature bagging. The random feature selection ensures that the trees are independent of each other and provides a better performance due to a better bias-variance trade-off. While bagging involves ensemble voting, boosting involves sequentially building an ensemble by training a new model by re-weighting. The re-weighting involves adjusting the weights of the samples which the previous model has had highest error on. The recent algorithms that demonstrate this technique are AdaBoost and XG-Boost.

Random Forests. The random forests (RF) were proposed by (Tin Kam Ho, 1995) and later extended

by (Breiman, 2001). They are considered to be one of the most robust and accurate learning methods. The technique is frequently used for both classification and regression. Using the concept of bagging, the algorithm builds an ensemble of decision trees. Each decision tree in the random forest provides a classification or prediction. By aggregating over the ensemble, final predictions are then made.

Over the last few years, significant changes have been made to increase the accuracy and overall performance of the algorithm. The growth of these ensembles depends significantly on the random vectors used to grow each individual tree. (Breiman, 996b) suggested the use of bagging where each tree is grown using a random selection (without replacement) from the training set. (Dietterich, 2000) suggested the use of random split selection, in which a random split is selected at each node from among the best m splits. (Tin Kam Ho, 1998) used the random subspace method. It uses a random subset of features to train every decision tree. Finally, (Breiman, 2001) suggested the use of out-of-bag error to estimate the generalization error.

Random Forests are still widely used due to their numerous advantages. They are straightforward to apply and can be easily parallelized. They provide accuracy as good as AdaBoost, even outperforming it in some cases and is relatively robust to outliers and noise. The various implementations of this algorithm in languages like Python and R provide useful estimates of error and variable importance.

Decision Trees. The decision trees are widely known approach, commonly used for supervised learning. They are majorly used for classification problems, but they could be used for regression analysis as well. Similar to a tree, a decision tree comprises of a root node, decision branches, interior nodes and leaf nodes. The root node represents the initial node or the most significant feature. It could be used to split the dataset into smaller subsets based on some decision rules or choices from a number of alternatives that are represented by the branches of the decision tree. The interior nodes and the leaf nodes represent intermediate features used for splitting and the final decision or outcome, respectively. There are various algorithms that could be used to build decision trees such as Chi-squared Automatic Interaction Detection (CHAID), Classification and Regression Trees (CART), C5.0 and Quest, etc. In a decision tree, each node may have two or more branches depending upon the algorithm. The advantage of using decision tree regression is that it requires very little bit of data processing and it can capture nonlinear patterns easily. It is easy to interpret, understand and visualize. In this article, we apply Decision Tree Regression (DTR). Unlike the conventional decision trees which are used for classification, the DTR can predict numerical outcomes of the dependent variables as well.

Adaptive Boosting. Adaptive Boosting or Adaboost is one of the first practical boosting methods introduced by (Freund and Schapire, 1997). This algorithm sequentially boosts weak learners to derive an ensemble of strong learners which gives a better prediction accuracy for the model. Similar to RF, Adaboost is also an ensemble technique whose baseline model is a decision tree. The weak learners are used in ensemble learning because a complex function could be described by simpler general trends which can provide a string approximation of that function. Since, it is often difficult to select an individual optimal learner from a variety of weak learners, boosting comes into play and simplify the selection process.

In Adaboost, weak learners are iteratively trained on the training dataset by continual sampling with replacement. The predictive performance of the weak learners is evaluated and the sampling distribution weights are correspondingly updated for those samples which were incorrectly classified by this weak learner. The next learner is trained on this resampled distribution with updated weights. Adaboost was initially introduced and widely used for classification problems, but it was also extended to regression analysis as well. The accuracy of the algorithm depends majorly on the baseline model's accuracy. It might not give good results if the prediction hypothesis is weak or sufficient data is not available. We aim to investigate the prediction performance of Adaboost for predicting numerical outcomes of the dependent variable in comparison to other models.

Extreme Gradient Boosting. eXtreme Gradient Boosting (or XGBoost) is an end-to-end scalable tree boosting technique used for supervised learning. It is an optimized implementation of gradient boosting machines. It was introduced by (Chen and Guestrin, 2016). Due to its highly optimized algorithm and system features, it is widely used by data scientists around the world. The XGBoost algorithm itself improves upon the gradient boosting machine framework by reducing computational time and efficiently allocating memory resources. It has sparse aware feature that handles missing values in the training data. It also uses LASSO (L1) and Ridge (L2) regularization to penalize the complexity of the model and comes with built-in cross-validation while training. It imple-

ments the distributed weighted quantile sketch algorithm to find the optimum split points. Apart from its algorithmic advantages, it also takes advantage of the system to perform efficiently. It uses parallelization, cache optimization, and out-of-core computing. During training, it makes use of all CPU cores for the construction of trees (parallelization). Its out-ofcore computing optimizes available disk space when handling large datasets that do not fit into memory. Through cache awareness, it efficiently uses hardware resources by allocating internal buffers in each thread to store gradient statistics.

2.5 Prediction Analysis

The model selection is followed by model prediction. The predictive power of the model ensures that the model is robust to new observations. A popular approach to study the model prediction is by constructing training and test data. The training data is used to build the model. The best fit model is found using training data. The test data is used to examine the prediction power of the best fit model. Here, for regression models the ratio of training to test data is 80:20. We apply m-fold cross-validation to the machine learning algorithms. There are various measures to study the predictive performance of the model. We use root mean square error and mean absolute scaled error. The measures are based on comparing the error between observed and predicted response.

m-fold Cross-validation. A m-fold cross validation is a resampling procedure in which the dataset is split into m smaller subsets known as folds. Using the leave-one-out concept, one fold is left out for testing while the model is trained on the remaining m-1 folds in each iteration. This process is repeated for m iterations in total and the aggregate result for the model is derived by averaging over the results of each iteration. The advantage of using m-fold cross-validation is that it gives less biased and yet more accurate estimate by reducing out-of-sample error.

Root Mean Square Error. The root mean squared error (RMSE) is a commonly used measure. It is given by RMSE= $\sqrt{\frac{1}{n} \sum_{i=1}^{n} (y_i - \hat{y}_i)^2}$. Here, y_i is the observed count response and \hat{y}_i is the predicted count re-

sponse for *i*th subject. The minimum value of RMSE on test data indicates good predictive performance of the model or machine learning algorithm under study.

Mean Absolute Scaled Error. When there are zero counts in the data, (Sellers and Shmueli, 2010) recommend to use mean absolute scaled error (MASE). The MASE is given by mean $\left| \frac{y_i - \hat{y}_i}{(1/n)\sum_{i=1}^n |y_i - \bar{y}|} \right|$.

R-squared. The value of R-square (R^2) is a goodness of fit measure commonly used while assessing the fit of models. It is given by

$$R^{2} = \frac{\sum_{i=1}^{n} (\hat{y}_{i} - \bar{y})^{2}}{\sum_{i=1}^{n} (y_{i} - \bar{y})^{2}},$$
(11)

The value of R^2 is bounded between 0 and 1. The R^2 value of 1 indicates that the model perfectly fits the data.

3 DATA

To demonstrate the methodology we use a real life data example. The data set is from Toronto Police Service Public Safety data portal (Police, 2019). The portal provides open access to the information on traffic collision in Toronto. The dataset consists of information on fatal and non-fatal; pedestrian and non pedestrian accidents. It has information on the type of vehicles involved in the accident, namely, automobiles, cyclist, motorcyclists, emergency vehicles and trucks. Similar data sets have been considered in numerous studies (see (Lord et al., 2005),(Lord et al., 2008)).

We considered data on non-fatal accidents of non pedestrians accidents from year 2013-2019. In the database, each accident is identified by a unique accident number (or ACCNUM). The ACCNUM is the same for the people involved in an accident. On the basis of ACCNUM we extracted information on the number of non-fatal injured people. In this study, we predict the number of non-fatal injuries (NoNonFatal) for non pedestrian accidents. The response variable, 'number of non-fatal injuries' (NoNonFatal) is a count response. It takes values from 0 to 13. There are various covariates in the database, but, many of them have missing observations. On the basis of relevance and data availability, we choose the number of people injured or involved in an accident (NoInjured), light, cyclist, automobile, truck, motorcycle, emergency vehicle and transit or city vehicle (Trsncityveh) as covariates in the study. The 'NoInjured' is a discrete count variable. The environmental covariate light has 9 categories so it was treated as a continuous variable. The categories of the light covariate are dark, daylight, dark and artificial, dawn, dusk, dusk

and artificial, daylight and artificial, dawn and artificial, other. The other covariates have binary categories yes or no so they are considered as categorical covariates.

4 ANALYSIS

We demonstrate the methodologies discussed in Section 2 on a real life data set. The data set is briefly described in Section 3. The response variable is number of non-fatal injuries (NoNonFatal) in an accident. After removing possible outliers and missing observations there are 1459 independent subjects and 8 covariates in the study. The mean and variance of the count response are 2.63 and 2.46, respectively. The frequency at count 2 is 646 (44.28%). The preliminary analysis shows that the data is possibly, almost equi-dispersed or slightly under-dispersed and inflated at k = 2. The histogram and boxplot of the response variable are shown in Figures 1 and 2, respectively. The correlation analysis on the covariates did not show any strong correlation. Thus, all the covariates mentioned in Section 3 are included in the study.



Figure 1: Histogram of response.

The count regression models described in Section 2.2 are built to study the relationship between response variable and covariates. The statistical models are also used to make the predictions. To perform the analysis the data is divided into training (80%) and test (20%) data. The training data set is used to construct and find the best model. While, the test data is used to make the predictions. In the training data, the response variable has mean 2.63 and variance is 2.51. The mean and variance of test data are 2.65 and 2.30, respectively. The frequency of 2 in the training data is 499 (42.76%) while the frequency of 2 in the test data is 147 (50.34%). The descriptive analysis shows that

Para	meter	kINB	kIP	NB	Poisson
Inte	ercept	0.4984*	0.4984*	0.4568*	0.4568*
		(0.2894)	(0.2894)	(0.2725)	(0.2725)
NoIı	njured	0.2433*	0.2433*	0.2444*	0.2444*
		(0.0087)	(0.0087)	(0.0083)	(0.0083)
Li	ight	0.0071	0.0071	0.0060	0.0060
		(0.0124)	(0.0124)	(0.0113)	(0.0113)
Су	clist	-0.0623	-0.0623	-0.0475	-0.0475
		(0.0567)	(0.0567)	(0.0487)	(0.0487)
Auto	mobile	-0.2640*	-0.2640*	-0.2006*	-0.2006*
		(0.0920)	(0.0920)	(0.0799)	(0.0799)
Tr	uck	-0.0376	-0.0376	-0.0331	-0.0331
		(0.0879)	(0.0879)	(0.0801)	(0.0801)
Moto	orcycle	-0.0135	-0.0134	-0.0081	-0.0081
		(0.0600)	(0.0600)	(0.0534)	(0.0534)
Emergen	cy Vehicle	-0.1144	-0.1144	-0.1091	-0.1091
		(0.2260)	(0.2260)	(0.2151)	(0.2151)
Trsnc	cityveh	-0.0886	-0.0886	-0.0805	-0.0805
		(0.0885)	(0.0885)	(0.0826)	(0.0826)
	Ŷ	-1.2777	-1.2778	-	-
		(0.1163)	(0.1163)		
	r	< 0.0001	-	< 0.0001	_
		-	· · ·	_	
	π	0.2179*	0.2179*	-	-
		(0.0198)	(0.0198)		
No. of p	arameters	11	10	10	9
Log-lil	kelihood	-1620.70	-1620.70	-1689.06	-1689.06
A	IC	3261.40	3261.40	3398.12	3396.12
В	SIC	3312.00	3312.00	3441.70	3441.70

Table 1: Estimates, standard errors (in parentheses) and model diagnostics (log-likelihood, AIC) for the training data.



Figure 2: Boxplot of the response.

the training and test data has high frequency at 2. A small positive difference between mean and variance is preserved which indicates the existence of almost equi-dispersion or a slight under-dispersion.

The high frequency of 2 in the training data indicates inflation at k = 2. Thus, we construct kINB, kIP, NB and Poisson regression models. The parameter estimates and their standard errors (in parentheses) are given in Table 1. The significant parameters at 10% level of significance are asterisk marked. From Table 1, it could be observed that NoInjured and Automobile are significant covariates in all the models considered under study. The intercept is significant at 10% level while other parameters are significant at 5% level of significance as well.

The Table 1 shows that Poisson is a parsimonious model as it has least number of parameters. While, NB and kIP have equal number of parameters. The kINB has 11 parameters, thus is the most complex model in the study. The kIP and kINB models have significant inflation at 5% level of significance with $\hat{\pi} = 0.2179$. The estimated value of the dispersion parameter (\hat{r}) in NB and kINB models is < 0.0001. When the data is under- or equi-dispersed then the NB model might not converge. On convergence it is reduced to the Poisson model. Here, the negative binomial (NB) model did not converge in SAS and R. Thus, it gives identical results to the Poisson model. While, the kINB model is reduced to the kIP model. The data is almost equi-dispersed, hence, NB and kINB models are reduced to Poisson and kIP, respectively.

To find the best model, we use LRT, AIC and BIC as described in Section 2.3. The AIC and BIC difference between Poisson and kIP model is 134.72

Parameter	kINB	kIP	NB	Poisson
Intercept	0.2219*	0.2218*	0.2131*	0.2131*
	(0.0399)	(0.0399)	(0.0353)	(0.0353)
NoInjured	0.2431*	0.2431*	0.2441*	0.2441*
	(0.0083)	(0.0083)	(0.0079)	(0.0079)
Automobile	-0.2194*	-0.2196 *	-0.1611*	-0.1611*
	(0.0825)	(0.0826)	(0.0699)	(0.0699)
Ŷ	-1.2795*	-1.2794*	-	_
	(0.1164)	(0.1164)		
r	< 0.0001	_	< 0.0001	-
	_		_	
π	0.7824*	0.7824*	_	-
No. of parameters	5	4	4	3
Log-likelihood	-1622.10	-1622.10	-1690.31	-1690.31
AIC	3252.20	3252.20	3388.62	3386.62
BIC	3272.50	3272.50	3408.86	3401.80

Table 2: Comparison of statistical models on training data with significant parameters.

Table 3: Predictive performance of the statistical models on training and test data with all the covariates.

	Train			Test				
Measure	kINB	kIP	NB	Poisson	kINB	kIP	NB	Poisson
RMSE	0.7919	0.7919	0.8578	0.8578	0.6932	0.6932	0.8172	0.8172
MASE	-	-	-	-	0.3840	0.3840	0.3400	0.3400

(>> 10) and 129.7 (>> 10), respectively. The kINB is reduced to the kIP model and has same AIC and BIC as kIP model. The NB model did not converge and the AIC of NB (3398.12) is slightly higher than the Poisson (3396.12) model. Based on the thumb rule given by (Burnham and Anderson, 2002) and (Kass and Raftery, 1995), the kIP model is best fit for the data. The Poisson and kIP models are nested so we can apply the LRT. The hypothesis is $H_0: \pi = 0$ vs. $H_1: \pi > 0$. The LRT statistic is $-2\log \Lambda = 136.72$ with p-value < 0.0001 and H_0 gets rejected. Equivalently, inflation at k = 2 is significant. Here, the NB did not converge in R and SAS. When the data is underor equi-dispersed then the NB is reduced to Poisson irrespective of convergence. On performing LRT on NB and kINB models, we get exactly same results. So, kINB has significant inflation at k=2.

We re-ran the models with only significant covariates, see Table 2. All the parameters are significant at 5% level of significance. We get the similar interpretations. The kIP and kINB models perform better than their base (Poisson and NB) models. The inflation at k=2 is significant in both the kIP and kINB models. We observe from Table 1 and 2 that the covariate NoInjured has a positive impact on the mean response, while, automobile has a negative impact.

We compared the predictive performance of the models on the test data. The RMSE values of the test and training data of kINB, kIP, NB and Poisson models are given in Table 3. The RMSE value of kINB and kIP models is same for the test and training data. Similarly, the RMSE value of NB and Poisson models are 0.8172 and 0.8578 for test and training data, respectively. The RMSE of kINB and kIP models is less than that of NB and Poisson models for both, test and training data. On comparing the NB and Poisson models to their k-inflated analogs we obtain MASE of 0.3840 for kINB and kIP models. While, for NB and Poisson we get MASE equal to 0.3400. The less RMSE and MASE on test data implies less error in prediction and hence better predictive powers. Therefore, kINB and kIP models have slightly less prediction error than their base counterparts.

The data science models like random forests and XGBoost are popularly used for predictions. To make the predictions on the count response, we use techniques described in Section 2.4. We apply decision trees, random forests, Adaboost and XGBoost on the training dataset using cross-validation. In the decision trees we apply decision tree regression (DTR) analysis which essentially uses the CART algorithm. Here, for cross validations we used five-folds. The predictions are made on test data. The RMSE and MASE are calculated to evaluate the performance of the machine learning algorithms. The scikit-learn library of Python is used to perform the analysis.

In the training data, apart from AdaBoost all the algorithms have RMSE ~ 0.34 . The RMSE of de-

Table 4: Predictive performance and goodness of fit of the machine learning algorithms (Decision Trees (DT), Random Forests (RF), Adaptive Boosting (AdaBoost), eXtreme Gradient Boosting (XGBoost) on training and test data with all the covariates.

	Train			Test				
Measure	DT	RF	AdaBoost	XGBoost	DT	RF	AdaBoost	XGBoost
RMSE	0.3406	0.3424	0.3866	0.3420	0.3814	0. 3748	0.4238	0.3698
MASE	_	_	_	_	0.1225	0.1181	0.1578	0.1150
R^2	0.9617	0.9610	0.9410	0.9611	0.9468	0.9489	0.9252	0.9495

Parameter	Decision Trees	Random Forests	AdaBoost	XGBoost
NoInjured	0.9930	0.9906	0.9878	0.9784
Light	0.0031	0.0041	0.0062	0.0029
Cyclist	0.0010	0.0011	0.0013	0.0079
Automobile	0.0005	0.0010	0.0008	0.0018
Truck	0.0006	0.0010	0.0009	0.0032
Motorcycle	0.0005	0.0010	0.0013	0.0027
Emergency Vehicle	< 0.0001	< 0.0001	< 0.0001	0.0004
Trsncityveh	0.0009	0.0009	0.0013	0.0024

Table 5: Comparison of feature importance using machine learning algorithms.

cision trees is minimum (0.3406) and the RMSE of random forests and XGBoost are very close. The AdaBoost has slightly higher RMSE than random forests. Further, the machine learning algorithms are run on the test data. We observe, XGBoost has lowest RMSE (0.3698) while, the AdaBoost has the highest RMSE (0.4238). The random forests perform better than decision trees with a decrease of 1.73% in RMSE. On comparing random forests to XGBoost we observe an improvement of 1.33%. The XGBoost shows a significant decrease of 3.04% when compared to decision trees.

The average MASE is determined by evaluating the average of MASE values obtained using five-fold validation. According to MASE measure, the least value is of the XGBoost (0.1150). While, the highest is of AdaBoost (0.1578). From Table 4, we observe that there is slight improvement in the predictions as we go from decision trees to random forests. Similarly, XGBoost performs better than random forests. To study the fit of the machine learning algorithms we study R^2 values of the training and test data. From Table 4, we observe that the XGBoost provides the best goodness of fit with an R^2 of 0.9611 and 0.9495 for training and test data, respectively. The R^2 of decision trees, random forests and XGBoost are approximately close, while, AdaBoost provides slightly poor fit. Additionly, from Table 5, we observe that 'NoInjured' is the most important feature to predict the response. The feature is significant in the statistical models as well (see Table 1).

In conclusion, from the preliminary analysis we observe that the training data is almost equi-dispersed or slightly under-dispersed and inflated at k = 2. Us-

ing AIC, BIC and LRT, kIP model is the most appropriate model for the data set. On the basis of prediction measures kIP and kINB performs equally well. It is observed that the data is almost equi-dispersed thus kINB is reduced to kIP model. The kIP model is parsimonious. Thus, kIP model is the best choice for the data set. While we have used parametric approaches via statistical modeling to build a model and make predictions, we also aim to study the non-parametric approaches provided by the machine learning algorithms. Using R^2 as the goodness of fit measure and RMSE and MASE as predictive indicators on various machine learning algorithms, we observe that as expected XGBoost performed best. Notably, the decision trees which do not involve any ensemble methods performed poor as compared to the ensemble models but significantly better than the statistical models.

5 SUMMARY

While, there are many studies on zero inflated count data there are only a few applications shown on k-inflated count data. In real life data sets, the occurrence of k > 0 inflated count data is not insignificant. We demonstrate the application of k-inflated count models. When there is inflation at k and data is equidispersed, k-inflated Poisson is the most appropriate model. We observe that the statistical models provide a good fit and decent predictions. However, the predictions obtained using machine learning algorithms are easy to apply and the computational time is very less. The complexity of the machine learning algorithms make them difficult to comprehend. Although, the algorithms provide information on the importance of the covariates; they do not provide any information on the direction (positive or negative) of contribution in the model. Thus, to study k—inflated count data sets, the corresponding regression models are appropriate for interpretations while, machine learning algorithms give superior predictions. So, it is recommended to study both the approaches. Our future work involves studying the approach on a larger data set from a different area. We plan to extend the comparative study by including various artificial neural network approaches.

REFERENCES

- Alfredo, S. G., Dutang, C., and Petrini, L. (2018). Machine learning methods to perform pricing optimization. a comparison with standard glms. *Variance, Casualty Actuarial Society*.
- Arief, F. M. and Murfi, H. (2018). The accuracy of xgboost for insurance claim prediction. *International Journal* of Advances in Soft Computing and Its Applications.
- Arora, M. (2018). Extended Poisson models for count data with inflated frequencies. PhD thesis, Old Dominion University.
- Bae, S., Famoye, F., Wulu, J., Bartolucci, A., and Singh, K. (2005). A rich family of generalized poisson regression models with applications. *Mathematics and Computers in Simulation*, 69(1):4–11. Second Special Issue: Selected Papers of the MSSANZ/IMACS 15th Biennial Conference on Modelling and Simulation.
- Breiman, L. (1996b). Out-of-bag estimation. Technical report, Department of Statistics, University of California, Berkeley.
- Breiman, L. (2001). Random forests. *Machine Learning*, 45(1):5–32.
- Burnham, K. P. and Anderson, D. R. (2002). Model Selection and Multimodel Inference. Springer.
- Cameron, A. C. and Trivedi, P. K. (2013). Regression Analysis of Count Data. Cambridge Press, London, UK.
- Chant, D. (1974). On asymptotic tests of composite hypotheses in nonstandard conditions. *Biometrika*, 61:291–298.
- Chen, T. and Guestrin, C. (2016). Xgboost: A scalable tree boosting system. Proceedings of the 22nd ACM SIGKDD International Conference on Knowledge Discovery and Data Mining.
- Dietterich, T. G. (2000). An experimental comparison of three methods for constructing ensembles of decision trees: Bagging, boosting, and randomization. *Machine Learning*, 40(2):139–157.
- Freund, Y. and Schapire, R. E. (1997). A decision-theoretic generalization of on-line learning and an application to boosting. *Journal of Computer and System Sciences*, 55(1):119–139.

- Greene, W. (1994). Accounting for excess zeros and sample selection in Poisson and negative binomial regression models. Working papers, New York University, Leonard N. Stern School of Business, Department of Economics.
- Gurmu, S. and Trivedi, P. (1996). Excess zeros in count models for recreational trips. *Journal of Business & Economic Statistics*, 14(4):469–477.
- Hall, D. B. (2000). Zero-inflated Poisson and binomial regression with random effects: A case study. *Biometrics*, 56:1030–1039.
- Kass, R. E. and Raftery, A. E. (1995). Bayes factors. Journal of the American Statistical Association, 90(430):773–795.
- Lambert, D. (1992). Zero-inflated Poisson regression, with an application to defects in manufacturing. *Technometrics*, 34:1–14.
- Lee, S.-K. and Jin, S. (2006). Decision tree approaches for zero-inflated count data. *Journal of Applied Statistics*, 33(8):853–865.
- Lin, T. H. and Tsai, M.-H. (2012). Modeling health survey data with excessive zero and *k* responses. *Statistics in Medicine*, 32:1572–1583.
- Lord, D., Guikema, S. D., and Geedipally, S. R. (2008). Application of the Conway-Maxwell-Poisson generalized linear model for analyzing motor vehicle crashes. *Accident Analysis & Prevention*, 40:1123 – 1134.
- Lord, D., Washington, S. P., and Ivan, J. N. (2005). Poisson, poisson-gamma and zero-inflated regression models of motor vehicle crashes: balancing statistical fit and theory. Accident Analysis & Prevention, 37(1):35 – 46.
- Payandeh Najafabadi, A. T. and MohammadPour, S. (2018). A k-inflated negative binomial mixture regression model: Application to rate-making systems. Asia-Pacific Journal of Risk and Insurance, 12(2).
- Police, T. (2019). Toronto police service public safety data portal. website.
- Ridout, M., Demetrio, C., and Hinde, J. (1998). Models for count data with many zeros. In *International Biometric Conference, Cape Town*.
- Schwarz, G. (1978). Estimating the dimension of a model. *The Annals of Statistics*, 6(2):461–464.
- Sellers, K. F. and Shmueli, G. (2010). Predicting censored count data with COM-Poisson regression. Technical report, Robert H. Smith School Research Paper No. RHS-06-129. Available at SSRN: https://ssrn.com/abstract=1702845 or http://dx.doi.org/10.2139/ssrn.1702845.
- Shapiro, A. (1985). Asymptotic distribution of test statistics in the analysis of moment structures under inequality constraints. *Biometrika*, 72:133–144.
- Tin Kam Ho (1995). Random decision forests. In Proceedings of 3rd International Conference on Document Analysis and Recognition, volume 1, pages 278–282 vol.1.
- Tin Kam Ho (1998). The random subspace method for constructing decision forests. *IEEE Transactions on Pattern Analysis and Machine Intelligence*, 20(8):832– 844.