Developing a Machine Learning Model for Predicting Postnatal Growth in Very Low Birth Weight Infants

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Abstract: Objective of the work is the development of prognostic machine learning models that predict qualitative and quantitative measures of postnatal growth in very low birth weight preterm infants. Observational retrospective data about 964 infants at risk are retrieved from “Fondazione Monza e Brianza per il bambino e la mamma”’s electronic medical record. Both prenatal (gestational, socioeconomic, etc.) and perinatal (nutritional, respiratory assistance, drug prescription and daily growth) data up to a week after birth are the features included. Model’s performances are compared to previous literature and human performance, showing a substantial improvement (in e.g., best regression MAE=0.49, best classification AUC=0.94).

1 MOTIVATIONS AND BACKGROUND

With increasing survival of preterm neonates, optimizing postnatal growth is an essential component in the medical management of this population.

Postnatal growth in premature infants is a strong predictor of outcomes, both as a reflection of concurrent morbidities as well as long-term neurodevelopment. In extremely low birth weight (BW) infants (ELBW; BW<1000g), poor growth during neonatal intensive care unit (NICU) stay is associated with adverse short-term outcomes, as well as it exerts a significant, and independent effect, on neurodevelopment outcome in infancy.

Despite the increasing knowledge about neonatal nutrition and growth, the NICHD Neonatal Research Network found the overall incidence of extrauterine growth restriction (EUGR) to be 79% among extremely preterm infants (Clark et al., 2003, Miller et al., 2014, Ehrenkranz et al., 2006). Therefore, maintaining adequate nutrition and growth in this high-risk population is an important, and difficult to achieve, goal in the NICU. (Lucas et al., 1998, Hayakawa et al., 2003, Stoll et al., 2010)

The prevention of growth impairment and severe nutrient deficits during hospital stay may be achieved through the implementation of the knowledge of macro nutrients/micro nutrients needs, the optimization of nutritional policies and the individualization of the nutritional intervention. Optimizing postnatal growth is in fact an integral component of the management strategies for preterm infants, and an important outcome measure in the NICU. EUGR therefore may be used as one of the objective measures of quality of nutritional care in premature infants - lower EUGR incidence meaning better quality of care.

Objective of the analyses will be to develop and validate a model that correlates prenatal and perinatal features, such as nutritional intakes (both parenteral, e.g. intravenously and enteral, e.g. orally) on postnatal growth and is able to make a prediction on the infant’s growth at discharge. In other words, we would like to investigate: can we predict accurately the postnatal growth, and does the prediction’s interpretation make clinical sense?

The paper is organized as follows: section 2 presents the methods used to collect data and the development of the models; section 3 presents the main results. Finally, section 4 discusses the findings of this work and its limitations.
2 MATERIALS AND METHODS

2.1 Data Available

The observational data has been retrieved from “Fondazione Monza e Brianza per il bambino e la sua mamma”’s (FMBBBM) Neonatal Intensive Care Unit (NICU), one of the largest and most advanced NICUs in Italy. FMBBBM’s electronic registry, Metavision, is property of ‘iMDsoft’.

Infants at risk were included (gestational age<32 weeks OR birth weight<1500g), born from January 2006 to July 2019. The patients come from tertiary care. It is a monocentric study: all the infants were inborn in the San Gerardo hospital, were admitted to the NICU up to 1 day from their birth and had a stay time over 7 days (mean, sd 45±29).

2.2 Outcome

The definitions of EUGR are not consistent in the literature. Several methods have been proposed to quantify the degree of EUGR.

Weight alone shouldn’t be used as an indicator of wellness; a score that considers the infant’s age, sex and weight needs to be used. In 2014, the Intergrowth-21st Consortium published international standards for newborn baby size and postnatal growth (Villar et al., 2014), based on neonates with no major complications or ultrasound evidence of fetal growth restriction (FGR), who were born to healthy mothers without risk factors for FGR.

The Weight Z score was calculated at discharge using the postnatal growth for preterm infants calculator (The International Fetal and Newborn Growth Consortium for the 21st Century, 2019). The Z score at discharge was selected as the quantitative measure of EUGR for each patient, and it is the target of the predictive regression model. For the binary classification model, we defined the EUGR positive class as <10th percentile at discharge (Radmacher et al., 2003).

2.3 Related Work

It is possible to compare our regression model with an existing work from Lin et al (Lin et al., 2015). The inclusion criteria are similar. Lin et al. also use 32 weeks as a threshold for gestational age; however they don’t consider birth weight. They also exclude deaths, unlike in this work. Overall, they include 1,714 infants, compared to our 964. Their outcome was selected as the difference of Weight Z scores from birth to discharge; in this work Z score at discharge alone is considered as the quantitative outcome. Also, the Z scores were calculated using a different method (Fenton, 2003). The model used for prediction is a standard multivariate linear regression. Only two risk factors were considered by the authors: birth weight Z score and completed weeks of gestation. As reported by Lin et al., many important variables are missing from their model, and many of these are implemented in this work as suggested in their discussion, such as Apgar score, gender and ethnicity.

Other works include the one of (Radmacher et al., 2003), who use a logistic regression to predict a binary EUGR class. We will compare our performance to their reported AUC score. (Lee et al., 2018) is also related, however it does not report any performance metric for comparison.

2.4 Features and Preprocessing

We decided to include in the model both prenatal and perinatal features.

- With prenatal features we mean parameters or events occurred before the birth.
- With perinatal features we mean parameters or events occurred after birth (within the first week).

It is thought that most of the main events that influence the outcome results and therefore condition the infant growth happen up to this period.

Neonatal characteristics are recorded: gestational age, being appropriate or small for gestational age, birth weight, length, head circumference, Apgar score at 1 and 5 minutes. Additional neonatal features are: twin or simple childbirth, fetal position at birth, amniotic fluid and placenta appearance, fetal distress, PROM duration, GBS screening risk factors and outcome. Centiles and Z scores at birth weight are calculated using the Lancet 21st Intergrowth newborn size for very preterm infants calculator (The International Fetal and Newborn Growth Consortium for the 21st Century, 2019). Infants with birth weight above 10th or below 10th percentile for gestational age, according to the Lancet 21st Intergrowth chart, are classified as appropriate for gestational age (AGA) or small for gestational age (SGA), respectively. Corrected age is calculated from the chronological age adjusting for gestational age. Nutritional intakes, both enteral and parenteral, are retrieved. Cumulative nutritional intakes are calculated as the sum of the daily intakes. Family features are included, such as: alcohol or smoke abuse during pregnancy, pathologies and social problems, whether the course of pregnancy was pathological or physiological, eventual consanguinity of the parents and their profession, level of schooling, ethnicity, age at child’s birth and anamnesis.
Perinatal features were also recorded, up to a week from birth. Average daily hours of mother’s and father’s presence in the NICU are included. For encoding morbidity, therapy was used as a reliable indicator, as database input is always required from an operator before prescribing any medication. 64 different drugs usage were recorded, each one encoded as number of prescription days during the first 7 days after birth. Another 44 additional drugs were present in the database, however, they were not included in this work, as they were never prescribed to any of the infants. For infants with respiratory assistance the most common respiratory assistance mode during day 0 to 6 was included. Weight percentiles after birth were calculated daily using the 21st Intergrowth postnatal growth for preterm infants calculator (The International Fetal and Newborn Growth Consortium for the 21st Century, 2019).

Missing values were imputed using K Nearest Neighbors (KNN) for numeric values (k=3). For categorical values, depending on the nature of the feature, they were imputed as ‘Unknown’ (when such values were already present in the data) or using KNN imputation otherwise.
### Table 1: Descriptive statistics of categorical features. Most common % calculated over non-null values.

<table>
<thead>
<tr>
<th>Feature</th>
<th>Unique values</th>
<th>Most common (MC)</th>
<th>MC %</th>
<th>Miss %</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sex</td>
<td>2</td>
<td>Male</td>
<td>51.5</td>
<td>0.0</td>
</tr>
<tr>
<td>GA Aspect</td>
<td>3</td>
<td>AGA</td>
<td>68.7</td>
<td>19.5</td>
</tr>
<tr>
<td>Social Problems</td>
<td>3</td>
<td>No</td>
<td>57.8</td>
<td>0.0</td>
</tr>
<tr>
<td>Pregnancy Alcohol</td>
<td>3</td>
<td>No</td>
<td>93.0</td>
<td>0.0</td>
</tr>
<tr>
<td>Pregnancy Pathology</td>
<td>26</td>
<td>MPP &lt; 34 GA weeks</td>
<td>26.7</td>
<td>0.0</td>
</tr>
<tr>
<td>Pregnancy Smoke</td>
<td>4</td>
<td>No</td>
<td>83.9</td>
<td>0.0</td>
</tr>
<tr>
<td>Fetal Position</td>
<td>8</td>
<td>Vertex</td>
<td>65.4</td>
<td>0.0</td>
</tr>
<tr>
<td>Birth Mode</td>
<td>6</td>
<td>C-sec, no labor</td>
<td>53.5</td>
<td>0.0</td>
</tr>
<tr>
<td>Birth Twins</td>
<td>2</td>
<td>Simple</td>
<td>75.2</td>
<td>0.0</td>
</tr>
<tr>
<td>Pregnancy Course</td>
<td>3</td>
<td>Pathological</td>
<td>82.0</td>
<td>0.0</td>
</tr>
<tr>
<td>Amniotic Fluid</td>
<td>8</td>
<td>Normal</td>
<td>71.4</td>
<td>0.0</td>
</tr>
<tr>
<td>PROM</td>
<td>4</td>
<td>No</td>
<td>62.6</td>
<td>0.0</td>
</tr>
<tr>
<td>Fetal Distress</td>
<td>5</td>
<td>No</td>
<td>66.9</td>
<td>0.0</td>
</tr>
<tr>
<td>Placenta</td>
<td>10</td>
<td>Normal</td>
<td>46.5</td>
<td>0.0</td>
</tr>
<tr>
<td>Group B Streptococcus (GBS) Screening</td>
<td>5</td>
<td>Not carried out</td>
<td>57.6</td>
<td>0.0</td>
</tr>
<tr>
<td>GBS Risk Factors</td>
<td>6</td>
<td>GA &lt; 37 not induced</td>
<td>35.2</td>
<td>0.0</td>
</tr>
<tr>
<td>GBS Intrapartum</td>
<td>4</td>
<td>Adequate</td>
<td>49.3</td>
<td>0.0</td>
</tr>
<tr>
<td>Parents Consanguinity</td>
<td>5</td>
<td>No</td>
<td>60.3</td>
<td>0.0</td>
</tr>
<tr>
<td>Profession Father</td>
<td>7</td>
<td>Employee: blue collar</td>
<td>35.6</td>
<td>0.0</td>
</tr>
<tr>
<td>Schooling Father</td>
<td>6</td>
<td>Secondary school</td>
<td>37.6</td>
<td>0.0</td>
</tr>
<tr>
<td>Anamnesis Mother</td>
<td>18</td>
<td>Nothing to report</td>
<td>73.2</td>
<td>0.0</td>
</tr>
<tr>
<td>Profession Mother</td>
<td>7</td>
<td>Employee: white collar</td>
<td>35.4</td>
<td>0.0</td>
</tr>
<tr>
<td>Schooling Mother</td>
<td>6</td>
<td>Secondary school</td>
<td>42.2</td>
<td>0.0</td>
</tr>
<tr>
<td>Ethnicity Mother</td>
<td>10</td>
<td>Mediterranean Europe</td>
<td>68.9</td>
<td>0.0</td>
</tr>
</tbody>
</table>

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### 2.5 Predicting Postnatal Growth with Machine Learning

The features described in the previous section comprise of 245 features before encoding the categorical features. After creating dummies with one-hot encoding from the categorical features there are 411 features available. The large number of predictors calls for a feature selection, that is automated in the selected machine learning models.

Two specific machine learning models were taken into consideration: Random Forest (RF) and Least Absolute Shrinkage and Selection Operator (LASSO). A baseline model will also be provided, using a basic decision tree regressor. The intent is to show the performance of a basic model, and to compare the improvements of RF and LASSO over it.

Hyper parameter selection was automated with a grid search. For each model, the best performing parameters were selected, and performances were tested with 5-fold nested cross validation. The following hyper parameters were attempted for each model: (i) Random Forest: max depth ∈ \{5, 10, 20\} × number of estimators \{50, 100\}, (ii) LASSO: Alpha ∈ \{0.5, 1.0, 2.0\}, (iii) Decision Tree: max depth ∈ \{5, 10, 20, unlimited\}.

Models’ regression performances are reported with the Mean Absolute Error (MAE) and coefficient of determination \(R^2\) metrics.

The performance of the model is also compared with the performance on 30 test cases of a human being, an experienced doctor (and co-author of this article). As the doctor performance is based only on few numerical features made available in a CSV file, she is expected to have prognostic performance lower than the machine. She is also asked to report her confidence in the prediction in a 4-value scale, from 1 meaning ‘not confident at all’ to 4 ‘almost certain’.

We are aware that the better predictive performance of the model would not be enough to assert the usefulness of the proposed model, let alone general over-performance (as a human interpreter would never ground on tabular values only for their prognosis). That notwithstanding, demonstrating the model supremacy represents a necessary (but not sufficient) condition to consider the model reliable.

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### 2.6 Dichotomized Classification

We also show the results of a classification model. The target variable was dichotomized as follows: positive (EUGR) in the weight Intergrowth percentile at

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Table 2: Best performing regression models for each estimator. Performance assessed with $R^2$ and MAE score. Mean and 95% CI calculated over 5 fold nested cross validation. Table sorted by mean $R^2$.

<table>
<thead>
<tr>
<th>Estimator</th>
<th>Hyperparameters</th>
<th>$R^2$ (mean, CI)</th>
<th>MAE (mean, CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random Forest</td>
<td>Max Depth: 10, N Estimators: 50</td>
<td>0.74 [0.71, 0.77]</td>
<td>0.49 [0.47, 0.52]</td>
</tr>
<tr>
<td>LASSO</td>
<td>Alpha: 0.5</td>
<td>0.62 [0.58, 0.67]</td>
<td>0.60 [0.54, 0.65]</td>
</tr>
<tr>
<td>Human Doctor</td>
<td>None</td>
<td>0.32 [-0.11, 0.75]</td>
<td>0.59 [0.57, 0.62]</td>
</tr>
<tr>
<td>Decision Tree</td>
<td>Max Depth: 5</td>
<td>0.54 [0.47, 0.61]</td>
<td>0.64 [0.61, 0.67]</td>
</tr>
</tbody>
</table>

Table 3: Best performing classification models for each estimator. Performance assessed with Accuracy, ROC AUC and Matthews Correlation Coefficient (MCC) score. Mean and 95% CI calculated over 5 fold nested cross validation. Table sorted by mean AUC.

<table>
<thead>
<tr>
<th>Estimator</th>
<th>Hyperparameters</th>
<th>Accuracy (mean, CI)</th>
<th>AUC (mean, CI)</th>
<th>MCC (mean, CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Random Forest</td>
<td>Max Depth: 10, N Estimators: 50</td>
<td>0.84 [0.83, 0.86]</td>
<td>0.94 [0.92, 0.95]</td>
<td>0.71 [0.69, 0.72]</td>
</tr>
<tr>
<td>SVC</td>
<td>Degree: 3, Kernel: poly</td>
<td>0.83 [0.82, 0.85]</td>
<td>0.92 [0.90, 0.93]</td>
<td>0.66 [0.63, 0.69]</td>
</tr>
<tr>
<td>Logistic</td>
<td>Penalty: 11</td>
<td>0.80 [0.79, 0.80]</td>
<td>0.86 [0.85, 0.87]</td>
<td>0.60 [0.53, 0.67]</td>
</tr>
<tr>
<td>NB</td>
<td>None</td>
<td>0.73 [0.71, 0.74]</td>
<td>0.77 [0.75, 0.78]</td>
<td>0.44 [0.36, 0.51]</td>
</tr>
</tbody>
</table>

The model performs better than the clinician, as it was expected for this particular experiment. However, the average error is not significantly different. The human performance (MAE 0.52 [0.41, 0.64], $R^2$ 0.32 [-0.11, 0.75]) and the model’s MAE 95% confidence interval (mean 0.49, CI 95% [0.47, 0.52]) overlap. However, the clinician’s confidence in her predictions was quite low: a median of 2 (on a scale from 1 to 4, where 1 means “not confident at all” and 4 “almost certain”). Out of 30 manual test cases, the “not confident at all” option was selected 5 times (17%), “not very confident” 19 times (63%), “confident” 6 times (20%) and “almost certain” was never selected.

3 RESULTS

3.1 Patients’ Descriptive Statistics

Figure 1 and table 1 describe the predictors for 964 eligible infants. Nutritional intakes and 64 drugs numerical features are not included for brevity.

3.2 Regression Performance

Performances are calculated on the test set over 5 folds. The train set for each fold comprises 771 infants, while the test set 193. A baseline is provided by the means of a simple decision tree.

Regression performances are shown in table 2. Only the best hyperparameters combination is shown for each model. The overall best performing model is Random Forest with a max depth of 10 and 50 estimators.

3.3 Classification Performance

After dichotomization, 527 (55%) infants are in the ‘EUGR’ class, 437 (45%) in the ‘not EUGR’ class. The class distribution was balanced, so accuracy is a reliable estimate of the model performance. Classification performances are shown in table 3. Only the best hyperparameters combination is shown for each model. The overall best performing model is Random Forest with a max depth of 10 and 50 estimators, and its ROC curve is shown in figure 3.

3.4 Model Explanation

The most important features for the regressor model are shown in figure 4.

The birth weight Z score and centile reveal the fetus health status. These growth parameters measured at birth highly reflect the intrauterine growth. Accord-
According to the model, the prenatal conditions deeply impact on growth outcome stressing the importance of a careful obstetrical management of pregnant women.

Among the features that impact on outcomes, the growth parameters are the most important. The centile at day 6, that is the expression of the weight increase after the physiological weight loss of the first 3-5 days of life, seems to be determinant to ensure a proper growth also later on in life. There is a strict correlation between growth parameters and fluid (expressed as water) intake. It is very difficult to administer the proper amount of fluid to preterm infants. If on one hand an adequate fluid supply may reduce the entity of weight loss and therefore the entity of extrauterine growth restriction, on the other hand it has been demonstrated a correlation among high fluid intakes and morbidities (respiratory distress syndrome, arterial ductus patency…).

The electrolytes (as potassium and sodium) play a major role too. This is strictly related to the fluid balance. Administration of sodium and potassium in the first 3 days of life is a highly debated topic in the scientific literature. It’s quite impossible to avoid sodium administration after birth as sodium is contained in many parenteral nutrients (as phosphorus and amino acids compounds). Quite often the neonatologist has to decide if allowing to give sodium to infant to respect his phosphorus requirements or to allow the physiologic extracellular liquid reduction thus avoiding sodium (and therefore phosphorus) intake. Thanks to this model it’s now extremely clear that early sodium administration highly impacts on growth.

The model moreover considers all the therapies administered to the patients. The infants at more risk of growth restriction are those who receive antibiotic therapy (ceftazidime and gentamycin) expression of occurrence of severe sepsis. Surprisingly those factors impact on growth more than caloric and proteic supply.

Gestational age is a determinant factor; it is well known that the lower the GA, is the higher the risk of growth impairment is.

4 DISCUSSION AND CONCLUSIONS

The Machine Learning models predict the patient’s postnatal growth at discharge. The prediction can be made at one week after birth by providing the required features. We have also shown which parameters (features) are the most predictive, globally.

We compared our regression performance with those reported in the specialist literature (see section 2.3): notably, Lin et al. (Lin et al., 2015) report a simple multivariate linear regression model with a $R^2$ score of 0.23; conversely, ours achieved an $R^2$ 0.74 [0.72, 0.75]. To justify this impressive improvement, we could notice that our model encompasses features of postnatal care in the NICU, which their model did not consider.
In fact, the Random Forest model that achieved the best performance in our user study, encompassed hundreds of features. As discussed by Lin et al., adding too many factors introduces complexity from the data collection side, with relatively small returns on performance improvement. Thus, a limitation of our work is that many of the features considered are not easy to collect reliably, especially in less digitized NICUs; this might have negative effects on the real-world applicability of the model. Moreover, comparing our classification performance to (Radmacher et al., 2003), the authors report an AUC of 0.85; ours achieved a significantly higher performance: 0.94 [0.92, 0.95].

Comparing our regression performance to the clinician’s performance, we found that the average error is not significantly different, as seen in section 3.2. The confidence reported by the clinician was low; this might be because the prediction was made on the basis of tabular data instead of a live examination of the patient (i.e., the ordinary way to make a clinical and prognostic assessment for a doctor). However, the closeness of the two performances indicate that the proposed model has a potential for being useful in supporting a more objective clinician assessments. Future work could be aimed at comparing our model with the prediction of a clinician who knows and can visit the patient personally. In that case, we expect that the clinician will have higher accuracy.

The doctor was also asked to express what is in her opinion the maximum acceptable error (or equivalently minimum acceptable accuracy) on the part of a predictive model to consider this still useful for prognostic or therapeutic purposes. The error had to be expressed in terms of Z score, in a scale from zero (= only a perfect model that does not make any error is useful) to 3 (= a model that committed an error equal to practically half the tabulated weight). The doctor proposed a threshold at approximately two third of a standard deviation, an higher error than the mean MAE of our best performing model (Random Forest, MAE=0.49). This indicates that the predictive model’s error might be useful for clinicians, although more doctors’ opinions must be collected before making any definitive claim. We plan to spread this poll to other NICU practitioners and neonatologists to get an estimate of this minimum acceptable accuracy.

As for other future work, we would like to improve on the following aspects:

- To produce a more parsimonious model that uses less features with similar performances, in particular, to include only features that are easy to collect in order to make the model as accessible as possible;
- To test the performance of the above, less complex model on other NICUs, even in those that do not have access to electronic data recording;
- To evaluate the real-world use of the model (both usefulness and utility);
- To investigate the complex relationship between doctors and AI explanations, in both agreements and disagreements (Ribeiro et al., 2016).

In conclusion, the model provides solid performance and a substantial improvement over the literature (Lin et al., 2015, Radmacher et al., 2003), but still
Figure 4: Feature importance. The x axis marks the average decrease in variance over all the estimators (trees).

lacks a pragmatic validation in the field of work (Cabitza and Zeitoun, 2019).

REFERENCES


