

Similarity Function Learning with Data Uncertainty

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Keywords: Similarity Function, Uncertain Data, Missing Data, Face Recognition.

Abstract: Similarity functions are at the core of many pattern recognition applications. Standard approaches use feature vectors extracted from a pair of images to compute their degree of similarity. Often feature vectors are noisy and a direct application of standard similarity learning methods may result in unsatisfactory performance. However, information on statistical properties of the feature extraction process may be available, such as the covariance matrix of the observation noise. In this paper, we present a method which exploits this information to improve the process of learning a similarity function. Our approach is composed of an unsupervised dimensionality reduction stage and the similarity function itself. Uncertainty is taken into account throughout the whole processing pipeline during both training and testing. Our method is based on probabilistic models of the data and we propose EM algorithms to estimate their parameters. In experiments we show that the use of uncertainty significantly outperform other standard similarity function learning methods on challenging tasks.

1 INTRODUCTION

Many computer vision tasks like face verification or k -nearest neighbors classification include two steps: a feature extraction step which transforms the image into a feature vector and the computation of similarity scores between the feature vectors. The similarity score is the output of a parametric similarity function which is learned from training data.

The quality of extracted features has a strong influence on the system's overall performance and, in many applications, the uncertainty of a specific feature varies from one image to another. For example, the uncertainty of a local feature describing the top left corner of an image could depend on the signal to noise ratio in that area which can be different from one image to another and independent of the signal to noise ratio in, say, the bottom right corner. Nonetheless, this uncertainty information is ignored by most machine learning algorithms which simply treat each sample as a *point* in the feature space. To overcome this limitation, uncertainty-aware methods consider each sample as a *probability distribution* which is provided by the feature extraction process. Each sample has a specific distribution which reflects the uncertainty in the corresponding features.

In this paper, we design a method which takes advantage of uncertainty information to build a better

similarity function and we show that it helps to cope with images of different resolutions, pose variation or occlusion. Specifically, we extend the Joint Bayesian method (Chen et al., 2012) to deal with uncertainty information. The Joint Bayesian method is a similarity function learning algorithm which has been successfully applied to face verification. On the challenging LFW dataset (Huang et al., 2007) it is used in several of the best performing methods: (Cao et al., 2013), (Chen et al., 2012), (Sun et al., 2014a) and (Sun et al., 2014b).

This paper is organized as follows. In Section 2 we discuss the related work. To take into account uncertainty throughout the whole processing pipeline we propose an uncertainty-aware dimensionality reduction algorithm and a similarity function that we describe respectively in Section 3 and 4. Section 5 presents experiments which indicate the advantage of using uncertainty and finally we summarize our findings in Section 6.

2 RELATED WORK

Similarity learning has been a popular topic both in the machine learning and computer vision communities. Many methods have been developed in the recent years. Some are designed to improve near-

est neighbors classification like LMNN (Weinberger and Saul, 2009), whereas others such as ITML (Davis et al., 2007) or LDML (Guillaumin et al., 2009) are more generic. Several methods assume a statistical model of the data, often based on normal distributions, to build the similarity function. For example, the Linear Discriminant Analysis (LDA) or more recent methods like the Probabilistic LDA (Prince and Elder, 2007), KISSME (Köstinger et al., 2012) or the Joint Bayesian method (Chen et al., 2012) are all based on Gaussian models. As opposed to most similarity function methods, the Joint Bayesian method does not operate in the space formed by the difference of feature vectors but works on the joint distribution of feature vectors pair. To deal with uncertain data, this paper proposes to generalize the Joint Bayesian method by considering each sample as a probability distribution in the feature space instead of a simple point.

Whereas, up to our knowledge, this kind of approach has never been applied to similarity function learning, this idea has been explored for other machine learning tasks. Several classification algorithms have been extended to deal with uncertain data such as SVM (Bi and Zhang, 2004) and (Shivaswamy et al., 2006), decision trees (Tsang et al., 2011), or naive Bayes classifier (Ren et al., 2009). Clustering algorithms have also been adapted to uncertain data, see, for example, (Cormode and McGregor, 2008), (Kriegel and Pfeifle, 2005) and references therein.

The Probabilistic PCA (PPCA) (Tipping and Bishop, 1999) gives a probabilistic view point of the standard PCA. We have been inspired by it to design our dimensionality reduction algorithm presented in the next section.

3 DIMENSIONALITY REDUCTION

In computer vision and in face recognition in particular, raw features extracted from images (LBP, SIFT, Gabor jets, etc.) are often very high dimensional so, in order to limit the computational cost, most similarity function methods start with a dimensionality reduction step. PCA has been shown to be both simple and effective for this task but does not take into account any uncertainty information. In the next section, we propose a dimensionality reduction method which uses the uncertainty information to learn the low dimensional space and to project new feature vectors into it.

3.1 Uncertainty-aware Probabilistic PCA

Our dimensionality reduction method, Uncertainty-Aware Probabilistic PCA (UA-PPCA), uses a generative model similar to that used in Probabilistic PCA (Tipping and Bishop, 1999) or Factor Analysis. This latent variable model explains the observation \tilde{x} as the sum of a linear transformation of a low dimensional latent variable x and some noise. x is assumed to follow the standard multivariate normal distribution $\mathcal{N}(0, I)$. Specifically, our model can be written as

$$\tilde{x} = \mu + Wx + \tilde{\epsilon}_x \quad (1)$$

where $\tilde{x} \in \mathbb{R}^n$, $\mu \in \mathbb{R}^n$ is the center of the observation space, $W \in \mathbb{R}^{n \times m}$ relates the observation and the latent space, $x \in \mathbb{R}^m$ and $\tilde{\epsilon}_x \in \mathbb{R}^n$ is a Gaussian noise of distribution $\mathcal{N}(0, \tilde{S}_x)$. The uncertainty associated with the feature vector \tilde{x} is represented by the covariance matrix \tilde{S}_x .

The difference between PPCA or Factor Analysis and our method is that we make a different assumption on the noise distribution. In PPCA and Factor Analysis, a single covariance matrix for the noise is common to all samples. This makes possible to learn this matrix from the data. In contrast, in UA-PPCA, each vector $\tilde{\epsilon}_x$ has its own covariance matrix \tilde{S}_x which reflects the uncertainty in each component of the specific feature vector \tilde{x} . The matrices \tilde{S}_x being all different, they cannot be learned and therefore have to be provided by the feature extractor. They are regarded as fixed during the learning process.

Considering that two features are uncorrelated is very different from saying that the noises which affect them are uncorrelated. In a picture of a face, the appearance of the two eye are obviously correlated. However, the noises affecting them on a given image can very well be different if, let say, there is a cast shadow on one side of the face. In this paper, we assume that the noise is uncorrelated and therefore consider that the covariance matrices \tilde{S}_x are diagonal.

Usually, dimensionality reduction consists in finding low dimensional projections corresponding to high dimensional data. In the context of uncertainty-aware similarity function, the whole probability distribution of \tilde{x} needs to be transferred into the low dimensional space. Following our generative model, the low dimensional projection x and its associated uncertainty are respectively the mean and the covariance matrix of the conditional probability distribution $P(x|\tilde{x}, \tilde{S}_x, W, \mu)$. Using Bayes theorem and the Gaussian product rule we obtain the closed-form formula:

$$P(x|\tilde{x}, \tilde{S}_x, \mu, W) = \mathcal{N}(x|\mu_x, S_x) \quad (2)$$

$$\text{where } S_x = (W^\top \tilde{S}_x^{-1} W + I)^{-1} \quad (3)$$

$$\text{and } \mu_x = S_x W^\top \tilde{S}_x^{-1} (\tilde{x} - \mu). \quad (4)$$

3.2 Learning μ and W

In this section we present an Expectation-Maximization algorithm (EM) to learn the parameters of the model $\Theta = \{\mu, W\}$ from an unlabeled training dataset composed of feature vectors $\tilde{x}_i \in \mathbb{R}^n$ and their associated diagonal covariance matrices $\tilde{S}_i \in \mathbb{R}^{n \times n}$.

The EM algorithm is composed of two steps performed alternatively. The Expectation step (E-step) consists in estimating the parameters of the distribution of the latent variables x_i given the previous estimate of the parameters $\bar{\Theta}$. During the Maximization step (M-step), we maximize $Q(\Theta, \bar{\Theta})$, the expectation over the latent variables of the log-likelihood of the complete data, with respect to Θ . It is equal to

$$-\frac{1}{2} \sum_i \int P(x_i|\tilde{x}_i, \tilde{S}_i, \bar{\Theta}) (\tilde{x}_i - \mu - Wx_i)^\top \tilde{S}_i^{-1} (\tilde{x}_i - \mu - Wx_i) dx_i + \text{const} \quad (5)$$

where *const* is a term which does not depend on Θ and can therefore be ignored.

During the E-step we estimate the parameters of the distributions of the latent variables $P(x_i|\tilde{x}_i, \tilde{S}_i, \bar{\Theta})$ using equation (2). The M-step, namely the maximization of Q with respect to Θ , is achieved by solving the system of equations $\partial Q(\Theta, \bar{\Theta})/\partial \Theta = 0$. Specifically, $\partial Q(\Theta, \bar{\Theta})/\partial \mu$ is equal to

$$\sum_i \tilde{S}_i^{-1} (\tilde{x}_i - \mu - W\mu_{x_i}) \quad (6)$$

and $\partial Q(\Theta, \bar{\Theta})/\partial W$ is given by

$$\sum_i \tilde{S}_i^{-1} \left((\tilde{x}_i - \mu) \mu_{x_i}^\top - W (S_{x_i} + \mu_{x_i} \mu_{x_i}^\top) \right). \quad (7)$$

There is no closed-form solution for this system of equations in the general case. However, in our model we constraint the uncertainty covariance matrices \tilde{S}_i to be diagonal. In this case, we obtain a closed-form solution for each component of μ and each row of W , namely

$$\mu^{(j)} = \frac{\left(\sum_i \frac{\tilde{x}_i^{(j)}}{\tilde{S}_i^{(j,j)}} \mu_{x_i} \right)^\top A_j a_j - \sum_i \frac{\tilde{x}_i^{(j)}}{\tilde{S}_i^{(j,j)}}}{a_j^\top A_j a_j - \sum_i \frac{1}{\tilde{S}_i^{(j,j)}}} \quad (8)$$

$$W^{(j,\cdot)} = \left(\sum_i \frac{\tilde{x}_i^{(j)}}{\tilde{S}_i^{(j,j)}} \mu_{x_i} - \mu^{(j)} a_j \right)^\top A_j \quad (9)$$

$$\text{where } A_j = \left(\sum_i \frac{S_{x_i} + \mu_{x_i} \mu_{x_i}^\top}{\tilde{S}_i^{(j,j)}} \right)^{-1}, \quad (10)$$

$$a_j = \sum_i \frac{1}{\tilde{S}_i^{(j,j)}} \mu_{x_i}, \quad (11)$$

$(\cdot)^{(j,j)}$ denotes the j th element of the diagonal of a matrix, $(\cdot)^{(j,\cdot)}$ its j th row and $(\cdot)^{(\cdot,j)}$ the j th component of a vector. The parameters μ and W have to be initialized before the first iteration of the EM algorithm. We simply initialize μ to the empirical mean of the data and W to the m first leading eigenvectors of the empirical covariance matrix of the training set multiplied by the square-root of their respective eigenvalue. The computational complexity of each EM iteration is $O(D(d^3 + Nd^2))$ where D and d are respectively the dimensionality of the original and low dimensional feature vectors and N is the number of training samples.

4 UNCERTAINTY-AWARE JOINT BAYESIAN

In this section, we present our similarity function: Uncertainty-Aware Joint Bayesian (UA-JB). The feature vectors and their associated uncertainty covariance matrices used in this section are usually the outputs of the dimensionality reduction method presented in the previous section. However, when the dimensionality of the original feature space is not too large, we can bypass the dimensionality reduction stage and directly apply the similarity function. We start by describing the uncertainty generative model. The associated similarity function is presented in Section 4.2. Finally in Section 4.3 we propose an EM-based algorithm to learn the model parameters.

4.1 Generative Model

Gaussian generative models are very popular because they are both relatively simple and effective. Many face recognition algorithms rely on Gaussian assumptions such as FisherFaces (Belhumeur et al., 1997), KISSME (Köstinger et al., 2012), Joint Bayesian Faces (Chen et al., 2012), and PLDA (Prince and Elder, 2007). Those approaches model the data as the sum of two terms, namely, $x = \mu_c + \delta$, where μ_c is the center of the class to which x belongs to and δ is the deviation relative to its class center. We propose to split δ into two further terms, leading to the following model:

$$x = \mu_c + w + \epsilon_x \quad (12)$$

where w is the intrinsic variation of the sample from its class center μ_c and ϵ_x is an observation noise. As

opposed to the previous methods, this model explicitly takes into account the uncertainty information by considering that it affects the distribution of ϵ_x . All those variables follow zero mean multivariate normal distributions: $\mu_c \sim \mathcal{N}(0, S_\mu)$, $w \sim \mathcal{N}(0, S_w)$ and $\epsilon_x \sim \mathcal{N}(0, S_x)$. In the remaining of this paper, S_μ is called between-class covariance matrix, S_w within-class covariance matrix and S_x uncertainty covariance matrix.

S_μ and S_w are common to all samples and are unknown. We propose a EM algorithm to estimate them in Section 4.3. On the contrary, S_x is specific to each feature vector and is either computed by the Uncertainty-Aware Probabilistic PCA described in the previous section from the original feature vectors \tilde{x} and their uncertainty covariance matrix \tilde{S}_x or, directly provided by the feature extractor when dimensionality reduction is not needed. The uncertainty matrix of the original input features \tilde{S}_x is always diagonal but, after dimensionality reduction, the matrix S_x computed with (4) is a full covariance matrix.

4.2 Similarity Function

In Bayesian decision theory, decisions based on thresholding the likelihood ratio are known to achieve minimum error rate (Neyman-Pearson lemma). In this method we use the log-likelihood ratio associated with the above generative model as our similarity function.

Two feature vectors belonging to the same class (similar pair hypothesis: H_{sim}) share the same value for μ_c and only differ in their respective intrinsic variation w and observation noise ϵ_x . In contrast, two vectors from different classes (dissimilar pair hypothesis: H_{dis}) are totally independent.

Let x_i and x_j be two feature vectors and S_i and S_j their associated uncertainty covariance matrices. Following the same methodology as in (Chen et al., 2012), we derive the probability distributions $P(x_i, x_j | H_{\text{sim}}, S_i, S_j)$ and $P(x_i, x_j | H_{\text{dis}}, S_i, S_j)$ from the generative model (12) and compute the formula of the log-likelihood ratio $LR(x_i, x_j | S_i, S_j) = \log(P(x_i, x_j | H_{\text{sim}}, S_i, S_j) / P(x_i, x_j | H_{\text{dis}}, S_i, S_j))$. Specifically, a direct computation gives

$$\begin{aligned} LR(x_i, x_j | S_i, S_j) = & \\ & x_i^\top \left(M_1 - (S_\mu + S_w + S_i)^{-1} \right) x_i + \\ & x_j^\top \left(M_3 - (S_\mu + S_w + S_j)^{-1} \right) x_j + \\ & 2x_i^\top M_2 x_j - \log |S_\mu + S_w + S_i| - \log |M_1| + \\ & \text{const} \end{aligned} \quad (13)$$

where

$$M_1 = \left(S_\mu + S_w + S_i - S_\mu (S_\mu + S_w + S_j)^{-1} S_\mu \right)^{-1}, \quad (14)$$

$$M_2 = -M_1 S_\mu (S_\mu + S_w + S_j)^{-1}, \quad (15)$$

$$M_3 = (S_\mu + S_w + S_j)^{-1} (I - S_\mu M_2) \quad (16)$$

and const is a constant term which does not depend on neither x_i , x_j , S_i nor S_j and can therefore be ignored.

The similarity function is a quadratic form of the feature vectors x_i and x_j . The contribution of a specific component of the feature vectors to the similarity score depends on two factors: its discriminative power which is function of S_μ and S_w , and its reliability which is measured by S_i and S_j . The Uncertainty-Aware Joint Bayesian presented in this section combines those different types of information to compute a meaningful similarity.

4.3 Parameters Estimation

The parameters of our model are the covariance matrices S_μ and S_w and we propose an EM algorithm to estimate them.

We consider a training set with C different classes. Any class c contains m_c feature vectors, $x_{c,1}, \dots, x_{c,m_c}$. We denote by X_c the concatenation of those feature vectors and by $S_{x_{c,1}}, \dots, S_{x_{c,m_c}}$ their respective uncertainty covariance matrices. We define the latent variables $Z_c = \{\mu_c, w_{c,1}, \dots, w_{c,m_c}\}$ and the parameters to estimate $\Psi = \{S_\mu, S_w\}$. The graphical representation of the generative model of the dataset is depicted in Figure 1. The EM algorithm consists in iteratively maximizing $Q'(\Psi, \tilde{\Psi})$, the expectation of the log-likelihood of the complete data over the latent variables Z_c given the previous estimate of the parameter $\tilde{\Psi}$. Specifically, $Q'(\Psi, \tilde{\Psi})$ is given by

$$\sum_{c=1}^C \int P(Z_c | X_c, \tilde{\Psi}) \log P(X_c, Z_c | \Psi) dZ_c. \quad (17)$$

The standard E-step would consist in estimating the parameters of the distribution $P(Z_c | X_c, \tilde{\Psi})$. But Z_c might have a very high dimensionality especially for classes containing a large number of samples and therefore manipulating the parameters of $P(Z_c | X_c, \tilde{\Psi})$ could be a heavy computational burden. In order to make the optimization computationally tractable, we take advantage of the structure of the problem. Namely, we observe that the latent variables $w_{c,i}$ are conditionally independent among themselves given μ_c (see Figure 1). Therefore $P(Z_c | X_c, \tilde{\Psi})$ can be factorized as:

$$P(\mu_c | X_c, \tilde{\Psi}) \prod_{i=1}^{m_c} P(w_{c,i} | x_{c,i}, \mu_c, \tilde{\Psi}). \quad (18)$$

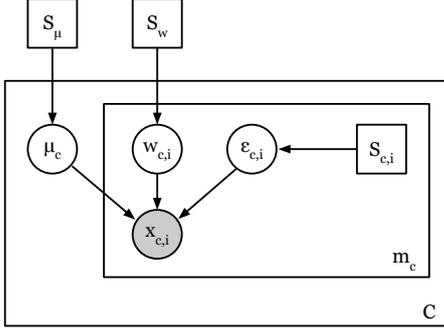


Figure 1: Graphical representation of the generation of the training set using plate notation. All the covariance matrices S_μ , S_w and $S_{c,i}$ are considered fixed in the generative model. However, while the matrices $S_{c,i}$ are provided by UA-PPCA or the feature extractor, the matrices S_μ and S_w are estimated by the EM algorithm.

To maximize $Q'(\Psi, \bar{\Psi})$ with respect to Ψ , we solve the equation $\partial Q'(\Psi, \bar{\Psi})/\partial \Psi = 0$. The optimal value for S_μ and S_w can be computed separately and we explicit the update formulas in the next two sections.

4.3.1 Update of S_μ

As shown in the next paragraph, the solution for S_μ depends on the parameters of the distribution $P(\mu_c|X_c, \bar{\Psi})$ which is a normal distribution $\mathcal{N}(\mu_c|b_{\mu_c}, T_{\mu_c})$ where

$$T_{\mu_c} = \left(\bar{S}_\mu^{-1} + \sum_{i=1}^{m_c} (\bar{S}_w + S_{c,i})^{-1} \right)^{-1} \quad \text{and} \quad (19)$$

$$b_{\mu_c} = T_{\mu_c} \sum_{i=1}^{m_c} (\bar{S}_w + S_{c,i})^{-1} x_{c,i}. \quad (20)$$

It is interesting to notice how the uncertainty impacts the probability distribution of μ_c . For samples with very large uncertainty, $(\bar{S}_w + S_{c,i})^{-1}$ becomes close to the null matrix and therefore these samples have little weight in the computation of T_{μ_c} and b_{μ_c} . This weighting operates at the feature level, meaning that a given sample can have a small weight for some features and a large one for others.

To find the matrix S_μ maximizing $Q'(\Psi, \bar{\Psi})$ we compute its gradient with respect to S_μ . It is given by

$$\sum_{c=1}^C \int P(Z_c|X_c, \bar{\Psi}) \frac{\partial \log P(\mu_c|S_\mu)}{\partial S_\mu} dZ_c \quad (21)$$

from which we obtain the closed-form update formula

$$S_\mu = \frac{1}{C} \sum_{c=1}^C (T_{\mu_c} + b_{\mu_c} b_{\mu_c}^\top). \quad (22)$$

4.3.2 Update of S_w

The optimization of $Q'(\Psi, \bar{\Psi})$ with respect to S_w requires the knowledge of the parameters of the distribution $P(w_{c,i}|X_c, \bar{\Psi})$. We can easily show that $P(w_{c,i}|X_c, \bar{\Psi}) = \mathcal{N}(w_{c,i}|b_{w_{c,i}}, T_{w_{c,i}})$ where

$$T_{w_{c,i}} = R_{c,i} S_{c,i}^{-1} T_{\mu_c} S_{c,i}^{-1} R_{c,i} + R_{c,i}, \quad (23)$$

$$b_{w_{c,i}} = R_{c,i} S_{c,i}^{-1} (x_{c,i} - b_{\mu_c}) \quad \text{and} \quad (24)$$

$$R_{c,i} = \left(S_{c,i}^{-1} + \bar{S}_w^{-1} \right)^{-1}. \quad (25)$$

The impact of $S_{c,i}$ on the parameters of the distribution is quite natural. If the uncertainty is large, the posterior probability $P(w_{c,i}|X_c, \bar{\Psi})$ converges to the prior $\mathcal{N}(w_{c,i}|0, \bar{S}_w)$. This is quite natural as in the absence of a reliable observation, the prior should be used. However, if the uncertainty is very small then $P(w_{c,i}|X_c, \bar{\Psi})$ converges to $\mathcal{N}(w_{c,i}|x_{c,i} - \mu_c, T_{\mu_c})$ which does not depend on the prior over $w_{c,i}$ anymore.

To maximize $Q'(\Psi, \bar{\Psi})$ with respect to S_w we compute its gradient which is given by

$$\sum_{c=1}^C \int P(Z_c|X_c, \bar{\Psi}) \sum_{i=1}^{m_c} \frac{\partial \log P(w_{c,i}|S_w)}{\partial S_w} dZ_c \quad (26)$$

and find the value of the matrix S_w which sets it to 0. The calculation uses the factorization (18) and leads to the closed-form update equation

$$S_w = \frac{1}{\sum_{c=1}^C m_c} \sum_{c=1}^C \sum_{i=1}^{m_c} (T_{w_{c,i}} + b_{w_{c,i}} b_{w_{c,i}}^\top). \quad (27)$$

4.3.3 Parameter Estimation Overview

EM algorithms need an initial estimate of the parameters to start with. We initialize S_μ and S_w with their respective empirical estimate. To this end, we compute the empirical mean of each class, set S_μ to the covariance matrix of the means and S_w to the covariance matrix of the difference of each sample with the mean of its class. After initialization, we alternate between the E-step: the computation of the parameters T_{μ_c} , b_{μ_c} , $T_{w_{c,i}}$ and $b_{w_{c,i}}$ using equations (19), (20), (23) and (24) and the M-Step: the update of S_μ and S_w using equations (22) and (27). This process is repeated until the Frobenius norms of the difference between two consecutive estimates of S_μ and S_w are both smaller than a predefined threshold. The complexity of each iteration of the EM algorithm is $O(Nd^3)$ where d is the feature vector dimensionality and N the number of training samples.

5 EXPERIMENTS

The set of experiments presented in this section demonstrates the performance of the Uncertainty-Aware PPCA and the Uncertainty-Aware Joint Bayesian. We present results on two datasets: MNIST to which we artificially add noise and FRGC to show how the use of uncertainty can contribute to tackle challenges in a real world application.

5.1 MNIST

MNIST dataset is composed of handwritten digit images of size 28×28 . We simply use the pixel values as feature vectors for this set of experiments. Performance on MNIST is usually measured by classification accuracy so similarity functions are commonly combined with a nearest neighbor classifier to perform the actual classification. Our aim is to investigate the impact of noise and uncertainty on the performance of similarity functions. To evaluate solely similarity functions, we have conducted a digit verification experiment (given a pair of images, do they contain the same digit?) and report the Equal Error Rate (EER). For information, we have observed that an EER of 10% usually leads to around 97% or 98% of classification accuracy.

On this dataset we artificially add noise to the images to create uncertain data. The data generation protocol takes two steps: first, for each image, for each pixel p , the noise standard deviation σ_p is drawn from a uniform law between 0 and t and second, we add to each pixel a noise drawn from a centered normal distribution with standard deviation σ_p . The uncertainty matrix of an image is simply the diagonal matrix containing the σ_p^2 of this image. By varying the value of t , we simulate different noise intensities. Figure 2 shows examples of an image affected by the three levels of noise we tested: none, medium and strong.

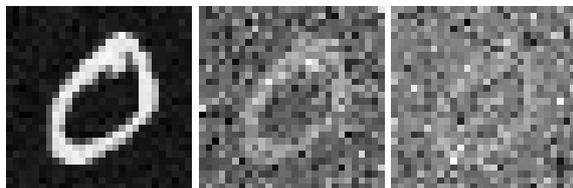


Figure 2: The three levels of additional noise: none (left), medium (middle) and strong (right).

We compare our method, Uncertainty-Aware Joint Bayesian (UA-JB), to three other methods: Joint Bayesian (JB) (Chen et al., 2012) to which our method is equivalent in the absence of noise, ITML (Davis et al., 2007) and LMLML (Bohné et al., 2014)

Table 1: EER on MNIST.

Noise Level	Methods			
	UA-JB	JB	ITML	LMLML
None	10.1%	10.1%	9.1%	8.7%
Medium	12.2%	13.5%	12.9%	12.5%
Strong	14.7%	20.6%	19.4%	18.8%

in single metric mode. We start by reducing the dimensionality to 100 using UA-PPCA for UA-JB and standard PCA for the three others as prescribed by the authors. As we can see in Table 1, the proposed method does not get the best results on noiseless data, however, thanks to the use of the uncertainty information, it outperforms the other methods on noisy data. Whereas error rates of other methods are more than doubled when a strong noise is added, UA-JB's EER relative increase is only of 46%.

In real applications the exact values of the uncertainty are unknown and only estimates can be provided to our algorithm. To evaluate its sensitivity to the accuracy of the uncertainty values, we propose to artificially perturb each σ_p by multiplying it by a factor uniformly drawn from $[0.7, 1.3]$ (for light perturbation) or $[0.4, 1.6]$ (for strong perturbation). Table 2 shows that our method is robust to this perturbation as the error rates increase of less than 11% even when a strong perturbation is applied.

Table 2: Sensitivity to the uncertainty accuracy.

Noise Level	Perturbation intensity		
	None	+/-30%	+/-60%
Medium	12.2%	12.3%	12.8%
Strong	14.7%	14.9%	16.3%

In Section 3 we have proposed a new dimensionality reduction method named UA-PPCA which takes uncertainty into account. We evaluate the performance of UA-JB if we use the standard PCA instead of the proposed method to compute the matrix W and μ and/or if we replace the projection described in Section 3.1 using $P(x|\tilde{x}, \tilde{S}_x, W, \mu)$ by the linear projections ($W\tilde{x}$ for feature vectors and $W^\top \tilde{S}_x W$ for uncertainty matrices). Ignoring uncertainty at the dimensionality reduction stage leads to higher error rates (see Table 3). UA-JB does not even bring any improvement over the Joint Bayesian method if standard PCA and linear projection are used because the highly uncertain features contaminate all the dimensions of the low dimensional space. Uncertainty needs to be taken into account throughout the whole processing pipeline to be effective.

Table 3: EER on MNIST with strong noise function of the dimensionality reduction method used for training (rows) and how the low dimensional projection is performed (columns).

Training	Projection	
	Linear	Probabilistic
PCA	20.2%	17.1%
UA-PPCA	18.8%	14.7%

5.2 Application to Face Verification

We have conducted experiments on different face recognition datasets to demonstrate that uncertainty can contribute to cope with challenges like image resolution changes, occlusion and pose variation. We used the FRGC, PUT and MUCT databases. On these biometric datasets, it is common to report performance by looking at the False Negative Rate (FNR) at a given False Positive Rate (FPR) which is typically quite low, such as 0.1%.

5.2.1 Resolution Change

The FRGC Experiment 1 dataset is composed of face images acquired in controlled conditions, there are variations in illumination and expression but the pose is always nearly frontal. In our experiments we train on 5000 images from 194 identities and test on 5000 images from 252 other identities.

We have aligned the images using eyes location. The native inter-eye distance is of approximately 80 pixels and during the alignment process the images are rescaled so that every image has an inter-eye distance of 64 pixels. Those images are called high resolution (HR) images in the remainder of this paper. Our feature vectors are composed of Gabor filter response magnitudes sampled on a regular grid (see (Li and Jain, 2011), Section 4.4 for more information). We use 4 scales and 8 orientations and the resolution of the grid is specific to each scale. The feature vectors we obtain are 14216-dimensional. For all the experiments with FRGC we have arbitrarily set the dimensionality of the space after reduction to 300. For other methods we compare ours to, standard PCA is used.

We created a low resolution (LR) version of each image by scaling it down by a factor 4 and then up by the same factor (using Lanczos resampling) so that they have the same size as the HR images. Figure 3 shows the two versions of an image.

The loss of resolution affects mostly the high frequency filters. It makes them more noisy but also shrink their distribution. To cope with this issue we post-process each feature vector depending on the res-



Figure 3: High resolution (left) and low resolution (right) versions of an FRGC image.

olution of the image. First, we subtract to each feature vector the mean of the feature vectors of its kind (HR or LR). Second, we multiply each component of LR feature vectors by a factor such that its variance after post-processing is equal to the sum of the variance of this component in HR feature vectors plus the variance of the noise. On a dataset including for each image the HR and LR versions, the noise variance is estimated by $\mathbb{E}[(x_{HR} - x_{LR})^2]$. The mean feature vectors and the factors have been computed once for all on a special training dataset, they are then used to post-process all the feature vectors involved in the training and the tests of the experiments presented in this section.

We now demonstrate the effectiveness of the proposed method to deal with scenarios where the training and the tests are performed on images of different resolutions. To this aim we have performed three experiments which differ by the images used for training. The training of the first experiment is performed with the HR images, that of the second with the LR images and for the last experiment a random mix of 50% of HR images and 50% of LR images is used. For each experiment we have evaluated the performance of all the methods on a test set of HR images and a test set of LR images. The results of the proposed method (UA-JB), Joint Bayesian (Chen et al., 2012), ITML (Davis et al., 2007) and LMLML (Bohné et al., 2014) are presented in Table 4. UA-JB performs well in all configurations and it worth noticing that, thanks to the use of the uncertainty, it is more robust than other methods. The benefit of using uncertainty is the most visible with the training on HR images because the other methods tend to learn that the high-frequency Gabor filters are the most discriminative whereas these features are very noisy when the test set is composed of LR images.

5.2.2 Occlusion

Occlusion is an issue in many applications of face recognition and uncertainty gives a framework to deal with it. In this experiment we use for training the non

Table 4: FNR at FPR=0.1% on FRGC depending on the training set and test set resolutions.

Train.	Test	Methods			
		UA-JB	JB	ITML	LMLML
HR	HR	2.5%	2.5%	4.1%	2.5%
	LR	4.1%	6.3%	8.4%	6.7%
LR	HR	3.0%	3.2%	5.3%	3.8%
	LR	3.0%	4.2%	6.6%	4.2%
Mix	HR	2.6%	2.7%	6.8%	2.7%
	LR	3.2%	4.6%	7.5%	4.2%

occluded HR images of FRGC described in the previous section. We have artificially created occluded test images by drawing random masks on the original images. The mask of each image is composed of two possibly overlapping rectangles which are symmetric with respect to vertical axis. We use symmetric masks because otherwise it would be too easy to recover the occluded part using the natural symmetry of faces. Figure 4 shows some examples of occluded faces. The masks on images are transformed into masks on feature vectors by considering that a feature is occluded if more than 5% of the energy of the corresponding filter is in an occluded area.

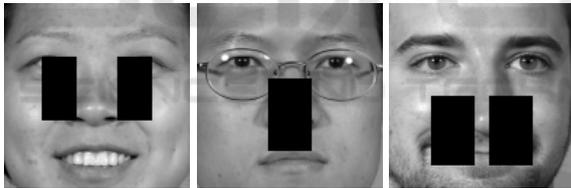


Figure 4: Examples of occluded faces.

Similarity functions can only compare feature vectors of a fixed specific size, therefore we need to provide a value for the occluded features too. We use a standard missing data imputation scheme based on the conditional probability of the hidden data given the visible ones for normally distributed data. Up to a feature reordering we can consider without loss of generality that all the occluded features are at the beginning of the feature vector. We use the formula of conditional multivariate normal random variables to compute the mean $o|_{v=a}$ and the covariance $S_o|_{v=a}$ of the filling pattern given the visible features v :

$$o|_{v=a} = \mu_o + C_{o,v}C_{v,v}^{-1}(\mu_v - a) \quad (28)$$

$$S_o|_{v=a} = C_o + C_{o,v}C_{v,v}^{-1}C_{v,o} \quad (29)$$

where μ_o and μ_v are respectively the mean of the occluded and visible features and C is the covariance matrix of the features which has the following struc-

Table 5: Impact of occlusion on the FNR at FPR=0.1% on FRGC.

	Methods			
	UA-JB	JB	ITML	LMLML
Standard	2.5%	2.5%	4.1%	2.5%
Occluded	8.0%	9.8%	12.5%	11.9%

ture

$$C = \begin{bmatrix} C_{o,o} & C_{o,v} \\ C_{v,o} & C_{v,v} \end{bmatrix}. \quad (30)$$

μ_o, μ_v and C are computed on the training set which is not occluded.

We provide to all methods the feature vectors where the occlusions have been filled with $o|_{v=a}$. $diag(S_o|_{v=a})$ is used by UA-JB as uncertainty matrix and is ignored by other methods.

As seen in the previous section, UA-JB exhibits similar performance to Joint Bayesian and LMLML on the original images but it outperforms them on the occluded images thanks to the use of uncertainty (see Table 5).

5.2.3 Pose Variations

Robustness to pose variations is a challenge for face recognition algorithms. A popular approach is to cancel most of the impact of pose variations with the help of a 3D morphable model. Synthetic frontal views are generated from non-frontal images and those synthetic images are used for comparison instead of the original ones. This process is called face frontalization. In our experiments, we use a method similar to that described in (Blanz et al., 2005) and use the Gabor-based feature vectors described in Section 5.2.1. Creating frontal views from non-frontal images is a difficult task and artifacts might appear on generated images, especially in portions of frontalized images which correspond to areas poorly visible in the original non-frontal views. In this section, we show that performance is improved if the most affected areas are not taken into account by the similarity function.

The pose of the face in a given image is estimated during the 3D morphable model fitting process. We propose to automatically choose a mask of pixels which should be ignored among a set of predefined masks function of the yaw angle estimated. Yaw angles are discretized into 5 bins: $yaw < -20^\circ$, $-20^\circ \leq yaw < -5^\circ$, $-5^\circ \leq yaw < +5^\circ$, $+5^\circ \leq yaw < +20^\circ$ and $+20^\circ \leq yaw$. Each bin is associated with a mask of pixels to ignore which has been empirically created. They are depicted in Figure 6. The discarded pixels are those which should be ignored during the



Figure 5: Original (left) and frontalized version (right) of an image from MUCT.

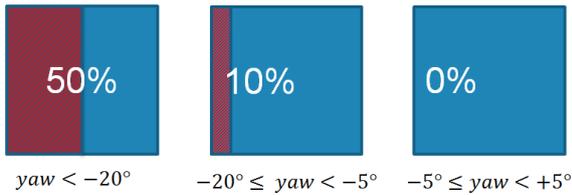


Figure 6: Masks associated with 3 of the 5 bins of yaw angle. The proportions of discarded pixels (hatched areas) are written in white.

comparison process because they are poorly visible on the original non-frontal image. These masks are transformed into uncertainty matrices on the feature vectors and are provided to our methods exactly as explained in Section 5.2.2 for random occlusions.

We use the FRGC images to learn the parameters of our model (μ , W , S_μ and S_w) and test our similarity function on two face datasets with large variations in pose: PUT (9971 images) and MUCT (3755 images). We compare our method to standard PCA + Joint Bayesian by looking at the FNR for a FPR of 0.1%. On PUT, our method obtains a FNR of 2.7% whereas the baseline achieves 3.1%. On MUCT, the FNR are respectively 3.4% and 3.6%. First, we remark that error rates on databases with pose variations are not much higher than those we report on FRGC in the previous section. This is due to the frontalization scheme used in this section, FNR are much higher if comparisons are performed on the original images. Second, we observe that using an uncertainty-aware similarity function leads to a notable improvement in performance on both databases despite the simple and coarse correspondence between yaw angles and pixel masks we use.

6 CONCLUSION

In this paper, we have introduced a novel similarity learning method which, unlike previous approaches, can take advantage of uncertainty information made available by the feature extraction process. The two

stages of our method are based on probabilistic models and we provided EM algorithms to estimate their parameters.

Our experimental results show the benefit of explicitly accounting for uncertainty information in similarity function learning. We demonstrate the effectiveness of our method on various challenging tasks such as dealing with images of various resolutions, pose variations or occlusion.

The main limitation of our work is that our method requires to be provided uncertainty information about the data. An interesting direction for future research is to automatize this task. This could be achieved by designing a method to make the link between some image quality measures (for example, local signal-to-noise ratio at the pixel level) and the data uncertainty matrices on extracted features.

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