

Simulation of Stochastic Activity Networks

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Abstract: Stochastic Activity Networks (SANs) are used in modeling and managing projects that are characterized by uncertainty. SANs are primarily managed using Monte Carlo Sampling (MCS). The accuracy of the results obtained from MCS depends on the sample size. So far the required sample size has been determined arbitrarily and independent of the characteristics of the SAN such as the number of activities and their underlying distributions, number of paths, and the structure of the SAN. In this paper we show that the accuracy of the SANs simulation results would depend on the sample size. Contrary to existing practices, we show that such sample size must reflect the project size and structure, as well as the number of activities. We propose an optimization-based approach to determine the project variance, which in turn is used to determine the number of replications in SAN simulations.

1 INTRODUCTION

Activity networks (AN) are known to be useful models for managing many real world projects. In routine projects such as construction projects, the time and resources required by each activity are known with certainty. In non-routine projects, the time or resources requirements of an activity may be, at best, characterized by a random variable with a given probability distribution function. Such networks are known as stochastic activity networks (SANs).

Many of the measures required for managing the SAN projects are hard to calculate using analytical methods. To illustrate the difficulty in calculating activity or project completion times consider the problem of calculating the probability distribution function (pdf) of the completion time of a project represented by the AN of Figure 1. The duration of activity i is represented by an independent random variable Y_i for $i = 1, 2, 3, 4, 5$, each with a specified pdf. The completion time of the project, represented by the random variable T , is the maximum of the duration of the three paths in SAN. Therefore, $T = \max \{T_1, T_2, T_3\}$ where T_j is the duration of path $j = 1, 2, 3$; or explicitly

$$T_1 = Y_1 + Y_4, T_2 = Y_1 + Y_3 + Y_5, \text{ and } T_3 = Y_2 + Y_5.$$

In general $T = \max \{T(k) \text{ over all } k \in P\}$ where P is the set of all paths in SAN. Consequently, T is the

maximum over many dependent random variables. While it is possible to calculate the exact pdf of T for small size SANs, such as the one in Figure 1, and for some underlying activity pdfs, it is not possible to calculate the exact pdf of T for larger size SANs and for various underlying activity distributions. This difficulty led to most of the research in SANs starting with the development of the well-known PERT procedure, Clark (1961). The research is focused on approximating or bounding the pdf of T or its statistics, Adlakha and Kulkarni (1989); and Herroelen and Leus (2005).

The difficulty in calculating pdf of T or any of the other measures stated above has led to the use of Monte Carlo sampling (MCS). It is assumed that as the number of simulation runs, known as sample size n , increases the resulting pdf of T and all of its statistics improve in their accuracy, and eventually converge to the exact values as $n \rightarrow \infty$. How large n should be to guarantee a certain level of accuracy in the estimated measures? The Central Limit Theorem (CLT) has been used to answer this question. For instance, in case of the mean value for the T , denoted by μ_T , if μ_s denotes the corresponding simulated value, then the level of accuracy is measured by the absolute difference $\varepsilon = |\mu_s - \mu_T|$, where it is desired to have the difference to be $\leq \varepsilon$ with a very high probability. Let $\alpha = 1 - \Pr(|\mu_s - \mu_T| \leq \varepsilon)$, then from the CLT it is concluded that

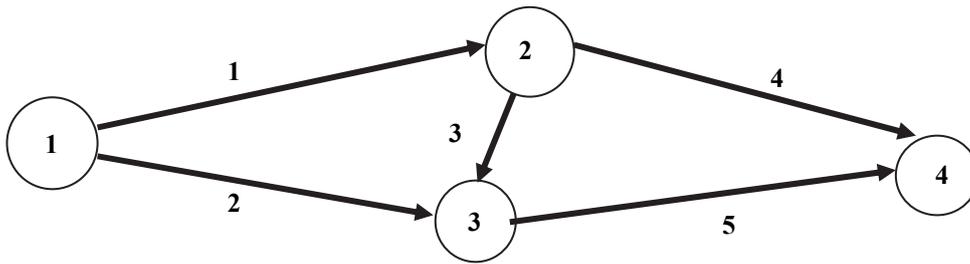


Figure 1: Example of a stochastic activity network.

$$n \geq Z_{\alpha/2}^2 \frac{\sigma^2}{\epsilon^2} \text{ where } Z_{\alpha/2} = \text{standard deviate for } \frac{\alpha}{2}$$

confidence level and $\sigma = \text{standard deviation of } T$.

In case of SANs, the above determination of the sample size n is problematic as it assumes that:

- Standard deviation of T , σ , is known
- The pdf of T converges to a normal distribution
- Sample size n is independent of SAN size or number of paths in the AN
- n is also independent of the underlying pdf of the activity durations
- n is also independent of the structure or complexity of the AN.

To illustrate some of the above problems, is it possible that for a given confidence level α that the required n to simulate the SAN of Figure 1 (with five activities, four nodes, and three paths) is equal to the n required to simulate a much larger SAN (such as a SAN with 60 nodes, 500 activities, and thousands of paths)? Second, it has been proven that in most SANs the pdf of T is not normally distributed, Dodin and Sirvanci (1990). This paper focuses on exploring the relationships that may exist between the sample size and the above factors.

2 SAN CHARACTERIZATION

The following notation is used throughout the paper:

A: Set of activities/arcs in the project

$M = |A|$; Number of activities in the project/SAN

N: Number of nodes in the SAN.

Y_{ij} : A random variable denoting the duration of arc $(i,j) \in A$ which starts in node i and ends in node j .

P: The set of all paths in the network.

$K = |P|$; number of paths in the SAN.

$T(k)$: The duration of path $k \in P$; $T(k) = \sum_{(i,j) \in P(k)} Y_{ij}$.

T: The duration of the longest path in SAN designating node N realization time as well as the

project completion time. Hence $T = \max_{k \in P} \{T(k)\}$.

The shape of the AN also may reflect the degree of dependency among the paths. A rectangular shaped AN tends to have many paths in parallel; hence, less dependencies among paths than a triangular one. It is well known that the higher the precedence relations between activities the greater the dependency among paths and the harder the task of managing. One of the ways to measure some of the above attributes is to use the the following AN complexity index, as suggested by Kolisch, Sprecher, and Drexl (1995). $ANCI = [\text{sum of all precedence relations}]/N$. In some studies N in the ANCI expression is replaced by M . ANCI is often used in designing ANs to test the efficiency of algorithms/heuristics that are used to manage the corresponding projects.

In case of SANs, ANCI can be used to characterize such networks in addition to other characterizations. The above expression of T indicates that the following factors can play a role in characterizing SANs:

1. Underlying pdf of the activities
2. Number of paths in the set P
3. Dependency among the paths
4. The gap between the duration (expressed in terms of the mean and variance) of the longest path in SAN and the next longest.

From the definition of T we notice that the longest path is not unique. In fact each path $k \in P$ can be the longest path but with certain probability. This probability is known in the literature as the criticality index of the path, and it is used in SANs to rank the paths and the activities, Dodin and Elmaghraby (1985). The pdf of T , denoted by $F(t)$, is given by $F(t) = \text{PR}(T \leq t) = \text{Pr}(T(k) \leq t \text{ for all } k \in P)$. Hence, $F(t) \leq \text{Pr}(T(k))$ for any one path $k \in P$. This shows that approximating $F(t)$ by the pdf of the duration of only one path forms an upper bound on $F(t)$. In general, SANs examples can be given to show that such an approximation is grossly optimistic. It can be shown that the joint pdf of any paths combination continues to form an upper bound

on $F(t)$. Therefore, the mean value of the duration of any individual path, or a surrogate SAN consisting of any combination of the P paths of the original SAN continue to form a lower bound on $E(T)$, Feller (1968), and Dodin (1985).

The above shows that, in general, T cannot be approximated by one $T(k)$ for $k \in P$, or a combination of paths in the set P and simulation may become necessary to determine the realizations of T and its pdf. We show below how the above factors characterize SAN with respect to the ease of calculating the pdf of T . To pursue this consider first the following two cases:

Case 1: A SAN consisting of only one path where the M activities are in series. In this case the CLT is applicable, and as M increases, the pdf of T converges to a normal distribution regardless of the pdf of the activity distribution. The project completion time T , mean, and variance of T are given by:

$$T = \sum_{ij} Y_{ij}, \mu(T) = \sum_{ij} \mu(i,j), \text{ and } \sigma^2(T) = \sum_{ij} \sigma^2(i,j),$$

and $F(t)$ can be easily calculated for any value of $t \geq 0$. In this case there is no need for simulation. This is also the case if a path $j \in P$ in SAN dominates all other paths $k \in P$, where a path j dominates a path k if

$$\Pr(T(j) \leq t) \leq \Pr(T(k) \leq t) \text{ for all } t \geq 0.$$

The dominance test is easy to implement if the number of paths in SAN is small as, based on CLT, each path is normally distributed. Its mean, variance, and pdf are calculated as above. However if the number of paths is large, then the dominance test can be cumbersome. In this case the dominance test can be replaced by calculating the gap between the longest path, j , and the second longest path, k , in mean; then perform the above dominance test between these two paths, j and k , or the weaker form of dominance given by:

$$\Pr(T(j) \geq T(k)) \geq \Pr(T(k) \geq T(j)) = 1 - \Pr(T(j) \geq T(k))$$

If j dominates k then in all of these instances T can be approximated by $T(j)$ and no need for simulation. In this case PERT Estimates are excellent regardless of the activities underlying distributions. For this type of network structure simulation is not needed as illustrated by Table 1 below.

What if there is no one path in SAN that dominates all other paths? The next case responds to this question:

Case 2: SAN consists of K independent paths in parallel. In this case $F(t)$ is given by: $F(t) = \Pr(T(k) \leq t \text{ for all } k \in P) = \prod_{k \in P} F_k(t)$, and $F(t)$ is not

normal even if any or all of the underlying K distributions are normal, David (1981), Galambos

Table 1: Mean, μ , and Standard Deviation, σ , for One Path Network.

Number of Activities	Parameter	Normal		Uniform		Exponential		Mixed	
		PERT	Sim	PERT	Sim.	PERT	Sim	PERT	Sim
10	μ	104.4	104.4	80.9	80.9	7.6	7.6	52.9	53.1
	σ	6.82	6.82	5.48	5.61	2.5	2.46	11.76	11.88
20	μ	208.3	208.2	170.3	170.4	42.5	42.5	124.3	124.3
	σ	9.701	10.15	7.74	7.67	13.86	13.77	8.30	8.44
40	μ	414.2	414.2	281.1	280.7	56.8	56.8	292.9	293.1
	σ	13.49	13.34	10.95	11.00	11.00	11.23	11.92	11.82

Table 2: Contrast of μ & σ for T of a SAN with K parallel paths.

Number of Paths	Normal				Uniform				Exponential				Mixed			
	Longest Path		Simulated for T		Longest Path		Simulated for T		Longest Path		Simulated for T		Longest Path		Simulated for T	
	μ	σ	μ	σ												
10	10	2	13.09	1.20	8	1.73	10.46	.50	.67	.67	1.97	.88	10	2	20.32	11.05
20	10	2	13.76	1.05	8	1.73	10.69	.28	.67	.67	2.41	.86	10	2	22.09	11.21
40	10	2	14.36	0.99	8	1.73	10.86	.14	.67	.67	2.87	.82	10	2	33.83	13.64

(1978). If the K paths are also identically distributed, then $F(t) = [\Pr(T(k) \leq t)]^K$. In this case $F(t)$ converges to an Extrême Value (EV) distribution with long right hand tail. Table 2 shows that as K increases $\mu(T)$ increases where the longest path mean and variance stay constant. Table 2 also shows that the underlying distributions and SAN's size affect T and its statistics. The increase in $\mu(T)$ is large especially if some of the underlying distributions have long right tails such as the exponential distribution. In spite of the fact that all paths are iid, except for the mixed case, $\mu(T)$ continues to increase as the number of paths increases. Consequently, if SAN consists of K iid paths, then T has an EV distribution. Its pdf, mean and variance can be easily computed without the need for simulation, Dodin and Sirvanci (1990). Consider a large SAN with several paths. In Case 1 above, the CLT continues to apply for any network as long as it contains a single path that dominates all other paths. Similarly in Case 2, the EV theory can be applied to many SANs with certain characteristics, while there are not many SANs that consist of many iid paths. As explained by Case 1, the duration of each path is normally distributed and the paths may not be independent. However, many of the paths may become normally distributed with close means, and variances, and the degree of dependencies is very low. Consequently, the conditions of the EV theory also exists in these networks, and T can be approximated by an EV distribution. This is the case for large SANs with rectangular or diamond shapes. For these networks the distribution of T is affected by two convergences: convergence to a normal distribution for the individual paths as the number of activities on each path increases, and convergence to an EV as

the number of paths increases and more of them become more iid. It was shown in Dodin and Sirvanci (1990) that if SAN has $m \geq 4$ dominating paths ; i.e. m close to being iid, then pdf of T can be accurately approximated by:

$$F(t) = \Pr(T \leq t) = \exp[-e^{-(b_m^{-1} \cdot a_m)}]$$

$$\mu(T) = a_m + 0.57722 / b_m, \text{ and } \sigma(T) = \pi / (2.45 b_m)$$

where.

$$a_m = \mu + \sigma[(2 \ln m)^{1/2} - (1/2)(\ln \ln m + \ln 4\pi) / (2 \ln m)^{1/2}], \quad b_m = \frac{\sqrt{(2 \ln m)}}{\sigma}$$

and μ & σ are the mean and standard deviation of the first dominating paths.

Figure 2 and Table 3 show three estimates for the pdf of T for a large SAN (with $N=40$ nodes and $M=100$ activities) where the activities have varying exponential distributions. SAN does not have one dominating path, but it has many dominating paths. It is clear that the pdf based on one of the longest paths (normal distribution known also as PERT distribution) provides a weak upper bound on pdf of T ; where that EV estimates are very close to the simulated $F(t)$.

The above two cases prove that if SAN has one dominating path, then T is normally distributed; and if it has several dominating paths then T has an Extreme value distribution. In both cases simulation can be avoided. Determining if SAN has one or several dominating paths is relatively easy. This follows from repeating the dominance test stated in **Case 1**.

If one path is not **the dominating path**, then the test is repeated by contrasting one of the first two longest in mean to the third longest in mean, then the fourth, and so on. This leaves us with the **third case** where SAN does not have either of these two characterizations.

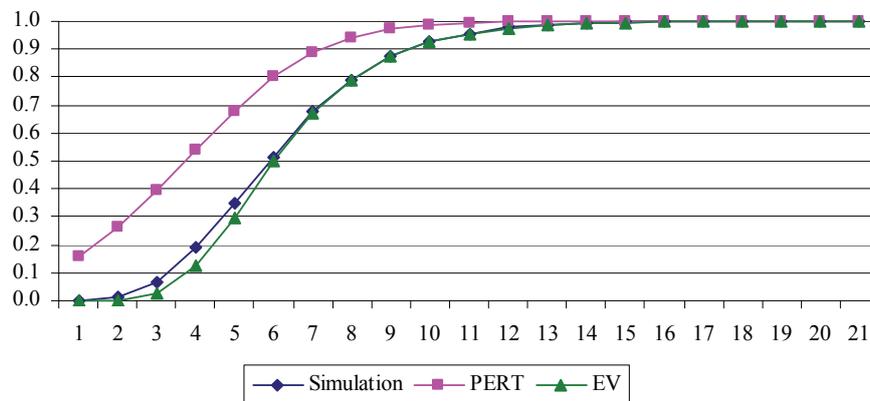


Figure 2: Cumulative distributions for the estimates of Normal, Simulation, and Extreme value.

Case 3: This includes SANs where the longest path/s in mean may not be the longest path/s in variance or vice-versa. The SAN is in between the above two extremes. Hence the pdf of T is not normally distributed and cannot be approximated by longest average (PERT), and cannot have the EV distribution. In this case simulation becomes necessary for obtaining an accurate approximation for T where determining the required sample size n is the main issue. The next section deals with the issue of determining the n required for achieving the desired level of accuracy.

3 SIMULATION SAMPLE SIZE

This section shows that project variance depends on the network structure expressed by the number of activities and activity precedencies. In order to prove such result, we first represent the project network into a longest path formulation. We assume that all projects have one dummy starting activity and a final ending dummy activity, whose variances and durations are equal to zero. So, all paths in the network must start and end with these two starting and ending dummy activities. The formulation assumes activity on arc (AOA) representation. Let

v_{ij} = variance of activity (i, j) $i = 1, \dots, N$

A_i = set of nodes connecting to i by activity (i,j) $i = 1, \dots, N$
where $A_N = \emptyset$

B_i = set of nodes connecting to i by activity (j,i) $i = 1, \dots, N$
where $B_1 = \emptyset$

$V(P)$ = maximum variance in the project network P

$x_{ij} = 1$ if activity (i,j) is on the longest variance path
 $= 0$ otherwise

Then program (P) represents the project network.

Program (P):

$$\text{Max } V(P) = \sum_{i=1}^{N-1} \sum_{j \in A_i} v_{ij} x_{ij}$$

Subject to :

$$\sum_{j \in A_i} x_{ij} = 1$$

$$\sum_{i \in B_n} x_{in} = 1$$

$$\sum_{i \in B_j} x_{ij} - \sum_{k \in A_j} x_{jk} = 0 \quad j = 2, \dots, N-1$$

The above model contains N nodes (constraints) and

$D = \sum_{i=1}^N |A_i|$ precedence or arcs (variables). Program

(p) solution would provide the maximum variance path in the project network, $V(P)$, regardless of its completion time. The size of program (P) in terms

of number of variables and constraints reflects the network structure and accordingly its complexity. Projects with larger number of nodes would have more constraints and an increased number of precedence among these activities would be reflected into higher number of variables.

Proposition: consider two project networks:

i. The first is a project P_0 consisting of N_0 activities and D_0 precedencies among these activities.

ii. The second project, P_1 , is a sub-network of project P_0 and, consists of N_1 activities and D_1 precedencies among these activities, where $N_1 \leq N_0$ and $D_1 \leq D_0$, showing that the network complexity index of project P_0 , is greater than or equal to that of project P_1 , $NCI(P_1) \leq NCI(P_0)$

Then the variance of $V(P_1) \leq V(P_0)$

Proof: The number of paths in project P_0 is \geq than that of project P_1 since it contains $N_1 \leq N_0$ and $D_1 \leq D_0$. Therefore the feasible region for program $V(P_1)$ is a subset of program $V(P_0)$. Clearly maximizing over a larger feasible region (larger number of paths) for project P_0 would yield a higher objective function than that of P_1 . Therefore the optimum objective function value of program (P₁), $V(P_1) \leq$ the optimum objective function value of program (P₀), $V(P_0)$.

The above Proposition ensures that the network structure is represented in calculating the project variance, $V(P_0)$.

The following example illustrates that the project variance is dependent upon the network structure including the number of activities and the number of precedencies.

In order to maintain the same project structure, we will formulate program (P) using both AOA and AON. The AOA model would be used to illustrate the impact of the number of activities on the variance, while the AON would be applied to show the impact of the number of precedencies on the variance. Table 4 presents a sample eight-activity project.

Solving the longest variance route AOA model for the Eight-Activity project yields a variance = 44. Suppose we drop two activities (3,4) and (4,6) while maintaining the same project precedence among the remaining activities. The longest variance route for the resulting six-activity project provides a variance = 42, which illustrates that a project with a larger number of activities would have a project variance that is greater than or equal to a project with a subset of those activities. Table 5 summarizes the results of variances for the three models solved for the same project. The table also shows that as the NCI

Table 3: Three estimates of F(t): Normal, Simulation, and Extreme value.

t	Simulation	Normal (PERT)	EV
26.35	0.00	0.16	0.00
31.79	0.02	0.26	0.00
37.23	0.07	0.39	0.03
42.67	0.19	0.54	0.12
48.11	0.35	0.68	0.30
53.55	0.52	0.80	0.50
58.99	0.68	0.89	0.67
64.43	0.79	0.94	0.79
69.87	0.88	0.97	0.87
75.31	0.93	0.99	0.93
80.75	0.96	1.00	0.96
86.19	0.98	1.00	0.97
91.63	0.99	1.00	0.99
97.07	0.99	1.00	0.99
102.51	1.00	1.00	0.99
107.95	1.00	1.00	1.00
113.39	1.00	1.00	1.00
118.83	1.00	1.00	1.00
124.27	1.00	1.00	1.00
129.71	1.00	1.00	1.00
135.15	1.00	1.00	1.00

Table 4: Sample Eight-Activity Project.

Activity (AoA)	(1,2)	(1,3)	(2,4)	(3,4)	(3,5)	(4,5)	(4,6)	(5,6)
Activity (AoN)	A	B	C	D	E	F	G	H
Precedence (AoA)	-	-	(1,2)	(1,3)	(1,3)	(2,4),(3,4)	(2,4),(3,4)	(3,5),(4,5)
Precedence (AoN)	-	-	A	B	B	C, D	C,D	E, F
Variance	9	6	10	15	12	10	8	13

Table 5: Maximum Variances for eight-activity project and its subset projects.

ProjectDescription	Number of			NCI	Maximum Variance
	Activities	Precedence	Nodes		
Original	8	9	6	1.5	44
Subset(activities D & G deleted)	6	5	6	0.83	42
Subset(No C-F & F-H precedence)	8	7	8	0.875	31

increases the project variance increases. Please note that in all cases, the variance provided by optimizing the longest variance route model would always be greater than or equal the variance of the longest path T.

Therefore, the variance given by the longest variance model would serve as the basis for computing the simulation sample size. Our proposed approach would ensure that the project network structure, expressed by the number of activities and precedence relationships, have a direct impact on the project variance and accordingly on the project simulation sample size. Let

σ_v^2 = the maximum variance V(P)

T_E = the critical path expected length

Then the proposed approach would be as follows:

4 PROPOSED SIMULATION APPROACH

- a. Formulate a longest variance model using program (P) with the variances as the cost coefficients of the (0, 1) variable corresponding to each activity.

- b. Optimize the model to get the maximum variance, σ_v^2 . This provides the value for the project variance. Please recall that the CLT would ensure that project duration distribution is a normal distribution with an expected project duration T_E and a standard deviation σ_v .
- c. For a specified upper bound on the error in T_E , given by ϵ , we can then compute a lower limit on the value of n , as follows: $n \geq Z_{\alpha/2}^2 \frac{\sigma_v^2}{\epsilon}$
- d. Replicate the simulation according to this n .

5 CONCLUDING REMARKS

In this position paper, we classify SANs based on the probability distributions of the activity durations as well as the network structure. We identify the situations where stochastic simulation is required. For those simulation experiments, we have also shown that the existing approaches for determining the number of replications based on the central limit theorem is inaccurate. The paper illustrates the need for determining the simulation sample size based on the project network structure and size. We proposed an optimization based approach to determine the number of replications based on the SAN structure and size. Future work would involve conducting extensive computational experiments on SANs with a variety of structures and size.

REFERENCES

- Adlakha, V. G. and Kulkarni, V. G., 1989, A Classified Bibliography of Research on Stochastic PERT Networks: 1966-1987, *INFOR*. 27(3): 272-296.
- Clark, C. E., 1961, The Greatest of a Finite Set of Random Variables. *Opns Res.* 9:146-162.
- David, H. A., 1981, *Order Statistics*, 2nd ed. Wiley, New York.
- Demeulemeester, E., Herroelen, W., 2002, *Project Scheduling: A Research Handbook*, Kluwer Academic Publishing.
- Dodin, B. M., 2006, "Practical & Accurate Alternative to PERT" Perspectives in Modern Project Scheduling, J. Weglarz, Springer International series; pp 3-23.
- Dodin, B. M. and Elmaghraby, S. E., 1985, Approximating the Criticality Indices of the Activities in PERT Networks, *Magt. Sci.* 31: 207-223.
- Dodin, B., and Sirvanci, M., 1990, Stochastic Networks and the Extreme Value Distribution, *Computers and Opns. Res.* 17(4): 397-409.
- Feller, W., 1968, *An Introduction to Probability Theory and its Applications*, Vol. I, 3rd ed., John Wiley & Sons, New York.
- Galambos, J., 1978, *The Asymptotic Theory of Extreme Order Statistics*. Wiley, New York.
- Glover, F., Klingman, D., and Phillips, N. V., 1992, *Network Models in Optimization and their Applications in Practice*, John Wiley & Sons, New York.
- Herroelen, W., and Leus, R., 2005, Project scheduling under uncertainty: Survey and research potentials, *European J. of Opnl. Res.*, 165(2): 289-306.