

Quantifying Depth and Complexity of Thinking and Knowledge

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Abstract: Qualitative approaches to cognitive rigor and depth and complexity are broadly represented by Webb's Depth of Knowledge and Bloom's Taxonomy. Quantitative approaches have been relatively scant, and some have been based on ancillary measures such as the thinking time expended to answer test items. In competitive chess and other games amenable to incremental search and expert evaluation of options, we show how depth and complexity can be quantified naturally. We synthesize our depth and complexity metrics for chess into measures of difficulty and discrimination, and analyze thousands of games played by humans and computers by these metrics. We show the extent to which human players of various skill levels evince shallow versus deep thinking, and how they cope with 'difficult' versus 'easy' move decisions. The goal is to transfer these measures and results to application areas such as multiple-choice testing that enjoy a close correspondence in form and item values to the problem of finding good moves in chess positions.

1 INTRODUCTION

Difficulty, complexity, depth, and discrimination are important and related concepts in cognitive areas such as test design, but have been elusive to quantify. Qualitative approaches are legion: Bloom's taxonomy (Bloom, 1956; Krathwohl et al., 1973; Anderson and Krathwol, 2001), Webb's Depth of Knowledge Guide (Webb, 1997), Bransford et al.'s studies of learning (Bransford et al., 2000; Donovan and Bransford, 2005). Quantitative approaches have mainly either inferred values from performance data, such as results on large-scale tests (Morris et al., 2006; Hotiu, 2006), or have measured ancillary quantities, such as deliberation time in decision field theory (Busemeyer and Townsend, 1993) or estimations of risk (Tversky and Kahneman, 1992).

Our *position* is to approach these concepts by starting in a domain where they can be clearly formulated, cleanly quantified, and analyzed with large data. Then we aim to transfer the formulations, results, and conclusions to domains of wider interest. Our home domain is *competitive chess*, in which the *items* are thousands to millions of positions from recorded games between human players in various kinds of high-level tournaments. Work to date (Chabris and Hearst, 2003; Haworth, 2003; Guid and Bratko, 2006, 2011; Regan and Haworth, 2011) has established solid relationships between quality mea-

asures arising from direct analysis of players' move decisions and standard skill assessment metrics in chess, mainly grades of mastery and the Elo rating system. Some prior work (Chabris and Hearst, 2003; Moxley et al., 2012) has extended the correspondence to time available and/or taken for (move) decisions, but this is still short of isolating depth or difficulty as factors.

Our aims are helped by similarities between the tasks of finding an optimal move (or at least a good move) in a chess position and finding the best answer to a multiple-choice question (or at least a good answer in case there are partial credits). There are also mathematical correspondences between the Elo rating system (Elo, 1978; Glickman, 1999) and metrics in Rasch modeling (Rasch, 1961; Andersen, 1973; Andrich, 1978; Masters, 1982; Andrich, 1988; Linacre, 2006; Ostini and Nering, 2006), item-response theory (Baker, 2001; Morris et al., 2006; Thorpe and Favia, 2012), and other parts of psychometrics.

Elo ratings r_P of players P maintain a logistic-curve relationship between the expected score of P over an opponent Q and the rating difference $r_P - r_Q$. A difference of 200 points gives roughly 75% expectation, and this has produced a scale on which 2200 is recognized as "master;" the highest few players are over 2800, and many computer *chess engines* are rated well over 3000 even on inexpensive hardware. The engines can hence act as an objective and authoritative "answer key" for chess positions.

Essentially all engines give values in standard units of *centipawns* and use *iteratively deepened* search. That is, beginning with $d = 1$ (or some other floor value) they search to a basic depth of d plies (meaning moves by White or Black, also called half-moves), give values $v_{i,d}$ to each legal move m_i at that depth d , and then deepen the search to depth $d + 1$. This incremental search can be capped at some fixed maximum depth D . Based on depth-to-strength estimates by Ferreira Ferreira (2013) for the Houdini 1.5a engine and matches run by us between it and versions 2.3.1 and 3 of the Stockfish engine used for the results reported here, we estimate depth 19 of the latter (in so-called *Multi-PV* analysis mode) at 2650 ± 50 .

Taking care to begin with an empty hash table for each position in each game, we use Stockfish's values $v_{i,d}$ for $1 \leq d \leq D = 19$ to quantify our key concepts. Our measures are weighted so that values of poor moves have little effect, so we could effectively bound the number of legal moves at $\ell = 50$. We consider moves ordered so that $v_{1,D} \geq v_{2,D} \geq \dots \geq v_{\ell,D}$ at the highest depth, but of course the highest value v_d^* for $d < D$ might equal $v_{i,d}$ where $i > 1$. We actually work in terms of the differences $v_d^* - v_{i,d}$, and in order to reflect that differences matter less when one side has a large advantage, we further *scale* them by defining

$$\delta_{i,d} = \int_{x=v_{i,d}}^{x=v_d^*} \frac{1}{1+a|x|} dx.$$

Here the constant a might be engine-dependent but we fix $a = 1$ since we used only two closely-related Stockfish versions. Cases where v_d^* is positive but $v_{i,d}$ is negative (meaning that move m_i is an error leading from advantage to disadvantage) are handled by doing the integral in two pieces. All $\delta_{i,d}$ values are nonnegative, and are 0 for the optimal move at each depth and any other moves of equal value. The key idea of *swing* is exemplified by these two cases:

- A move m_i *swings up* if $v_{i,d} < v_{j,d}$ for some other moves m_j at low depths d , but $v_{i,D} \geq v_{j,D}$ for (almost) all other m_j for depths d at or near the maximum analyzed depth D .
- The move *swings down*—and intuitively is a “trap” to avoid—if it has one of the highest values at low depths, but is markedly inferior to the best move m_1 at the highest depth: $v_{i,D} \ll v_{1,D} = v_D^*$.

It is expected in the former case that $v_{i,D} > v_{i,d}$ for lower depths d , and in the latter that $v_{i,D} \ll v_{i,d}$, so that a swinging move changes its absolute value, but it is its value *relative to other moves* that is primarily assessed.

2 METRICS AND RATINGS

At each depth d , the chess program produces an ordered list L_d of moves and their values. Comparing these lists L_d for different d involves standard problems in preference and voting theory, with the twist that high values from poor moves have diminished weight. We speak of *rating aggregation* rather than *rank aggregation* because the values of each move, not just the ordinal ranks, are important.

We postulate that *swing* should be a signed quantity in centipawn units that pertains to an individual move option, while *complexity* should be nonnegative and dimensionless and pertain to a position overall. Swing should reflect a bulk comparison of L_d for low d versus high d , while complexity can be based on how L_d changes to L_{d+1} in each round of search. Thus for complexity we may employ some divergence measure between ordered sequences $X = (x_i)$, $Y = (y_i)$ and sum it up over all d . Whereas common voting and preference applications give equal weight to all choices, we wish to minimize the effects of appreciably sub-optimal moves.

Any anti-symmetric difference function $\mu(x_i, x_j)$ gives rise to the *generalized Kendall tau coefficient*

$$\tau_{X,Y} = \frac{\sum_{i,j} \mu(x_i, x_j) \mu(y_i, y_j)}{\|\mu_X\| \cdot \|\mu_Y\|}, \quad (1)$$

where $\|\mu_X\| = \sqrt{\sum_{i,j} \mu(x_i, x_j)^2}$ and $\|\mu_Y\|$ is defined similarly. Then always $-1 \leq \tau_{X,Y} \leq +1$, with $+1$ achieved when $Y = X$ and -1 when $Y = -X$. If μ is homogeneous, so that $\mu(cx_i, cx_j) = c' \mu(x_i, x_j)$ where c' depends only on c , then $\tau_{X,Y}$ becomes scale-invariant in either argument: $\tau_{X,cY} = \tau_{cX,Y} = \tau_{X,Y}$.

The usual difference function $\mu(x_i, x_j) = x_i - x_j$ is linear, and also invariant under adding a fixed quantity to each value. It is not, however, invariant under augmenting the lists with irrelevant alternatives having low ratings. We swap these properties by employing

$$\mu(x_i, x_j) = \frac{x_i - x_j}{x_i^2 + x_j^2}$$

instead. When either x_i or x_j is large, say of order K representing a poor move, then $\mu(x_i, x_j)$ will have order at most $1/K$. Assuming that the same move is poor in Y , the augmentation will add terms of order only $1/K^2$ to the numerator and denominator of (1), yielding little change. This naturally confines attention to *reasonable* moves at any juncture. We define the *complexity* $\kappa(\pi)$ of a position π , for d ranging from the minimum available depth d_0 to $D - 1$, by:

$$\kappa(\pi) = 1 - \frac{1}{D-1} \sum_{d=1}^{D-1} \tau_{L_d, L_{d+1}}.$$

Notice that high agreement (τ always near 1) flips around to give complexity κ near 0. The definition of complexity might be modified by weighting higher depths differently from lower depths.

To define the *swing* of a move m_i we use a simple sum of scaled differences in value between depth d and the highest depth D , rather than average or otherwise weight them over d :

$$sw(m_i) = \sum_{d=1}^D (\delta_{i,d} - \delta_{i,D}).$$

This is a signed quantity—if positive it means that the value of move m_i “swings up”, while negative means it “swings down”—in the manner of falling into a trap. The overall “swinginess” of a position π , however, is a non-negative quantity. It is convenient first to define it between any two depths d and e :

$$s_{d,e}(\pi) = \sum_{i=1}^{\ell} |\delta_{i,d} - \delta_{i,e}|.$$

For overall swing it is expedient to dampen the effect of moves for which $\delta_{i,d}$ is large. Unlike the case with Kendall tau, we want to dampen a difference $|\delta_{i,d} - \delta_{i,e}|$ only if *both* values are large. We also wish to divide by a dimensionless quantity, in order to preserve the centipawn units of swing. Hence we postulate a scaling factor c that might depend on the chess program, and divide by an exponential function of the harmonic mean of the deltas divided by c :

$$v(\delta, \delta') = \exp\left(\frac{-2\delta\delta'}{c(\delta + \delta')}\right).$$

Since this paper uses only one chess program, we again take $c = 1$. Thus we define the damped overall swing between depths d and e by:

$$s_{d,e}^*(\pi) = \sum_{i=1}^{\ell} v(\delta_{i,d}, \delta_{i,e}) |\delta_{i,d} - \delta_{i,e}|.$$

Then the swing *at* depth d is given by $s_{d,d+1}^*(\pi)$, while the aggregate swing to the highest depth is defined by

$$S(\pi) = \sum_{d=1}^{D-1} s_{d,D}^*.$$

We employ weighted versions of this to define our key concepts. We desire the measure of *difficulty* to be in units of *depth* rather than centipawns. Our idea is that a position is deeper, hence more difficult, if most of the swing occurs at higher depths. It is OK to multiply it by the complexity since that is dimensionless.

Accordingly, we first define the *relative depth* ρ to be the depth below which half of the swing has occurred. For this we add up the swing from each

depth to the next, rather than the swing relative to the highest depth. With respect to nonnegative weights $w(d)$ summing to 1, define

$$\Sigma(\pi) = \sum_{d=1}^{D-1} w(d) s_{d,d+1}^*(\pi).$$

We used $w(d) = d$ normalized by $\sum_{d=1}^{D-1} d$. Then, letting $\Sigma_e(\pi)$ be the sum up to e rather than $D - 1$, define

$$\rho(\pi) = \max\{e : \Sigma_e(\pi) \geq \frac{1}{2}\Sigma(\pi)\} - \psi,$$

where the adjustment ψ term for the indicated e is

$$\psi = \frac{\Sigma_e(\pi) - \frac{1}{2}\Sigma(\pi)}{w(e-1)s_{e-1,e}^*(\pi)}.$$

Finally, we stipulate that the *analyzed difficulty* of a position π is given by

$$Diff(\pi) = \kappa(\pi) \cdot \rho(\pi).$$

For calculating the discrimination we use the relative depth of the position. We evaluate the mean α_l and standard deviation σ_l of $s_{d,D}^*$ values where $d \in (1, e - 1)$ ($e = \rho(\pi)$) and mean α_r and standard deviation σ_r of $s_{d,D}^*$ values where $d \in (e, D - 1)$.

The discrimination parameter Ψ of the position π can then be evaluated as:

$$\Psi(\pi) = (\alpha_l - \alpha_r) \left(\frac{\alpha_l}{\sigma_l} + \frac{\alpha_r}{\sigma_r} \right) \frac{\sum_{i,j} w_{i,j} (s_{i,D}^* - s_{j,D}^*)^2}{\sum_{i,j} w_{i,j}}.$$

The weights $w_{i,j} = 1/(j - i)$ where $i \in [l]$ and $j \in [r]$ ensures more emphasize to the depths near the difficulty of the position while calculating discrimination.

Our first of two main datasets comprised all recorded games in standard round-robin¹ tournaments in 2006–2009 between players each within 10 Elo of a “milepost” value. The mileposts used were Elo 2200, 2300, 2400, 2500, 2600, and 2700. The second comprised all 900 games of the 2013 World Blitz (WB) Championship, which was held in Khanty-Mansiysk, Russia, and distinguished by giving an accurate record of the moves of every game. This form of blitz, 3 minutes per game plus an increment of 2 seconds per move, is comparable to the historical “5-minute” form of blitz, and gives markedly less time than the minimum 90 minutes plus 30 seconds per move of the “milepost” games. Our idea was to test whether the blitz games were played at an identifiably lower level of depth. The average rating of the 60 WB players was 2611.

¹“Small Swiss” events with up to 64 players over 9 rounds were also included.

Table 1: Best move and number of total moves played various swing.

Level	Swing < 1		1 ≤ Swing < 2		2 ≤ Swing < 3		3 ≤ Swing < 4		4 ≤ Swing ≤ 5	
	#EMP	TM	#EMP	TM	#EMP	TM	#EMP	TM	#EMP	TM
WB	22,785	38,150	3,523	11,025	1,542	5,920	697	3,267	825	4,566
2200	4,954	7,967	812	2,331	364	1,308	184	781	220	1,025
2300	8,081	12,164	1,413	3,740	575	2,030	303	1,235	350	1,731
2400	8,878	13,536	1,575	4,127	754	2,296	340	1,301	493	2,049
2500	7,203	10,620	1,374	3,351	691	2,040	332	1,162	383	1,635
2600	3,252	4,689	701	1,619	323	918	165	507	213	823
2700	2,823	3,927	596	1,315	261	737	144	430	208	708

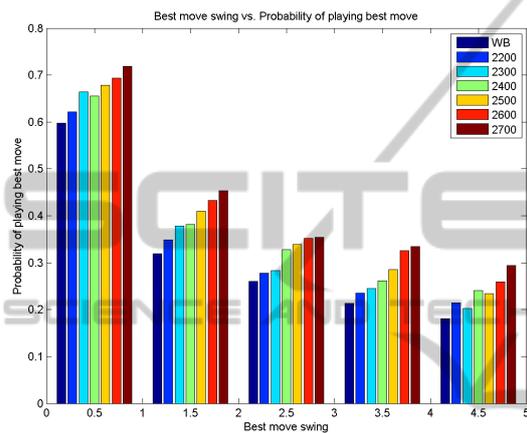


Figure 1: Frequency of playing engine moves with different swing values.

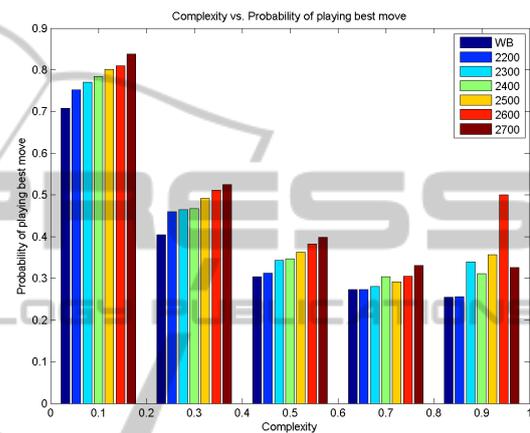


Figure 2: Frequency of playing engine moves for position with various complexity.

3 RESULTS

Our results show that the raw factor of *swing* makes a large impact on the ability of players at all levels to find the optimal move m_1 identified (at the highest depth) by the engine, and that this carries forward to our more-refined difficulty and discrimination measures. The WB games seemed to function as if they were a rating level below 2200, most often in the range 1800 to 2100.

Table 1 gives the total moves (TM) and times with the engine’s move played (EMP) for each of five intervals of swing values $sw(m_1)$, and Figure 1 graphs the frequencies of m_1 being played in each case. The plot clearly indicates that high-swing moves are “tricky” for players to find—the players more often chose inferior moves. The phenomenon is consistent with players of any Elo ratings, where higher rated players are slightly less tricked by the swing values. This feature is more prominent in the blitz tournament. Quick decision making often leads to pick inferior moves, or where the virtue of the engine move was not obvious at lower depths.

In our implementation, we rank the possible

moves at any particular position based on the order provided by the chess engines. Often the first move listed by the engine shows less swing, and make it attractive choice for the players from the beginning. Earlier studies show that players often chose the first move listed by the engine 58% of the time whereas the second move is chosen only 42% of the time. Table 2 shows that in fact the first listed move often has much lower swing with comparison to the other tied moves. This is true for players across any ability level.

Figure 2 represents the probability of playing the best move for positions of various complexity. The probability gets monotonically decreased. The random noise seen at positions with higher complexity is due to insufficient number of samples(see Table 3).

Figure 3 demonstrates difficulty and best-move probability for various positions. The figure clearly shows that players of all calibers could find the best move when the position is easy, but less than 50% of the time when the difficulty lies between 4 and 5. Table 4 shows the distribution of data across various difficulty levels. Figure 4 shows a similar but lesser effect for our measure of discrimination.

Table 2: Swing for Tied moves.

Level	Any # of Tied Moves			2 Tied Moves			3 Tied Moves			
	First	Second	#Moves	First	Second	#Moves	First	Second	Third	#Moves
WB	1.168	1.903	12,163	1.383	2.073	6,580	1.246	2.051	2.646	1,661
2200	1.263	1.912	2,447	1.430	2.056	1,391	1.141	2.005	2.265	329
2300	1.310	2.063	3,731	1.530	2.233	2,120	1.333	2.061	2.782	516
2400	1.330	1.964	4,270	1.555	2.178	2,514	1.211	1.841	2.449	569
2500	1.380	2.154	3,309	1.538	2.302	1,981	1.318	2.127	2.782	476
2600	1.408	2.216	1,607	1.533	2.349	974	1.244	2.338	3.083	242
2700	1.558	2.106	1,255	1.737	2.292	755	1.411	2.011	2.614	210
Overall	1.273	1.989	28,782	1.477	2.163	16,315	1.260	2.043	2.645	4,003

Table 3: Number of times best move played vs. number of total moves at positions of various complexity

Level	$0 \leq \text{Cpx.} < 0.2$		$0.2 \leq \text{Cpx.} < 0.4$		$0.4 \leq \text{Cpx.} < 0.6$		$0.6 \leq \text{Cpx.} < 0.8$		$0.8 \leq \text{Cpx.} \leq 1$	
	#EMP	TM	#EMP	TM	#EMP	TM	#EMP	TM	#EMP	TM
WB	14851	20983	7804	19270	5216	17155	1395	5104	106	416
2200	3083	4098	1863	4051	1216	3894	340	1244	32	125
2300	5227	6779	2933	6308	1984	5790	519	1849	59	174
2400	5672	7236	3421	7315	2282	6574	619	2036	46	148
2500	4590	5731	2916	5923	1946	5369	474	1625	57	160
2600	2061	2547	1384	2706	954	2496	232	761	23	46
2700	1840	2197	1146	2182	830	2084	200	605	16	49

Table 4: Best move and number of total moves for positions with various difficulty.

Level	Diff. < 1		$1 \leq \text{Diff.} < 2$		$2 \leq \text{Diff.} < 3$		$3 \leq \text{Diff.} < 4$		$4 \leq \text{Diff.} \leq 5$	
	#EMP	TM	#EMP	TM	#EMP	TM	#EMP	TM	#EMP	TM
WB	10438	13038	3790	6979	3462	7497	3189	8049	8493	27365
2200	2178	2569	754	1269	853	1585	741	1680	2008	6309
2300	3666	4318	1333	2089	1278	2450	1211	2665	3234	9378
2400	3925	4492	1481	2396	1463	2743	1452	3133	3719	10545
2500	3140	3512	1205	1848	1262	2304	1189	2459	3187	8685
2600	1392	1558	572	843	558	996	605	1198	1527	3961
2700	1183	1300	538	761	534	855	460	888	1317	3313

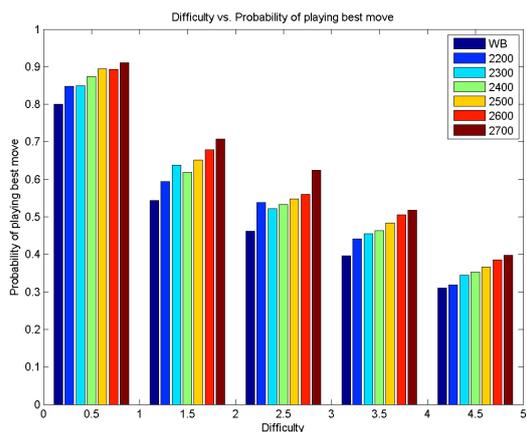


Figure 3: Frequency of playing engine moves for position with various difficulty.

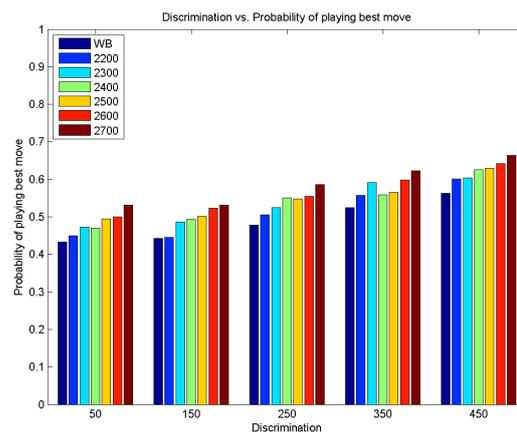


Figure 4: Frequency of playing engine moves for position with various discrimination.

4 CONCLUSION AND PROSPECTS

We have defined quantitative measures for qualitative concepts of depth, difficulty, complexity, and discrimination. The definitions are within a specific model of decision making at chess, but use no feature of chess apart from utility values of decision options, and are framed via mathematical tools that work across application areas. For the first three, we have shown a strong response effect on performance, though we have not distinguished the measures from each other. The effect shows across skill levels and persists when restricting to controlled cases such as moves of equal highest-depth value.

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