

# A Bayesian Approach to Modeling Dynamical Systems in the Social Sciences

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**Abstract:** The paper presents a new modeling approach using longitudinal or panel data to study social phenomena and to make predictions of dynamic changes. While the most common way in social sciences to study the relations between variables is using regression, our modeling approach describes the changes in variables as a function of all included variables, using differential equations with polynomial terms that capture linear and/or nonlinear effects. The mathematical models represented by these differential equations are derived directly from data. The models can then be run forward to forecast future changes. A two-step model-fitting approach is applied to identify the best-fit models and included visualisation methods based on phase portraits help to illustrate modeling results. We show this approach on an example relating democracy to economic growth.

## 1 INTRODUCTION

Since the 1960s when James S. Coleman published his book on mathematical sociology (Coleman, 1964), sociologists and other social scientists have been working on mathematical modeling of social phenomena. However, it is only recently with the availability of increasing computational power and sophisticated modeling tools that the field of mathematical social sciences is beginning to flourish. Mathematical modeling can be used both to study macro-level phenomena (Saperstein, 2000; Ashimov et al., 2011; Weber, 2012) as well as interactions at the micro-level (Coleman, 1964; de Marchi, 2005).

A widely adopted way of mathematically modeling relations between two or more variables is the regression equation, with the dependent variable  $y$  and the independent variable or predictor  $x$  (in case of a multivariate regression  $x_1, x_2, \dots$  represent the different independent variables), intercept  $\beta_0$ , slope  $\beta_1$  (in case of a multivariate regression  $\beta_1, \beta_2, \dots$  represent the different slopes related to the different predictors) and error term  $\epsilon$  with  $i, \dots, n$  observations.

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i \quad (1)$$

$$y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \dots + \epsilon_i \quad (2)$$

Data is used to estimate these equations and the strength of the relations. This approach can be ex-

tended to quite complex and sophisticated statistical models (Wooldridge, 2010). Such approaches are necessary to get a better understanding of social processes, but they have two limitations in the way they relate to the reality of social processes.

The starting point to empirical modeling is usually a social science theory, which tells the researchers what variables are to be considered and how they are expected to relate to each other in terms of cause-effect relationships (Treiman, 2009; Ostrom, 1990; Lewis-Beck, 1995). The purpose of the empirical modeling is then primarily theory testing and revising theories. While these are important parts of doing social science research, theoretical models usually need to be continuously tuned to account for data patterns.

Secondly, empirical modeling in social sciences does not always sufficiently take into account the fact that social systems are complex and dynamic. The most common way to study the interaction between variables is to compute linear or logistic regressions (Ostrom, 1990; Menard, 2001; Andersen, 2007). But, irrespective of the specific models used, the interpretation of results is most often static.

We suggest a novel approach to empirically based mathematical modeling in social science. Our data-driven Bayesian modeling approach uses the data itself to inform model selection from a pool of feasible models. While traditionally a regression of one vari-

able on another is performed, we model the changes in one variable as a function of all included variables (explained in Methods section below). Differential equations represent the mathematical models of a variable's change in time. We define a set of polynomial terms that express various possibilities of how variables may interact, allowing non-linear effects and build a model using Ordinary Least Squares (OLS) regression with these polynomial terms. The differential equation and therefore the mathematical model consists then of one or more of those polynomial terms that best describe the change in the variable as a function of itself and/or included predictors. A two step model fitting, using the maximum likelihood approach and the Bayes factor, is then used to look at how closely any candidate model fits the available data.

Compared to the theory-testing approach, ours is an exploratory modeling approach. Such an exploratory approaches in social sciences may help to find new and unexpected patterns and explanations (Gelman, 2004; Stebbins, 2001; Tukey, 1977). This explorative approach is not completely a-theoretical, since theories still suggest which variables we investigate. But instead of defining how the variables should interact and then testing this pre-defined relation in the data, we allow the data to inform us about the mathematical linear or non-linear relationships between the variables.

Our methodology can be applied to any social system which has reasonable amounts of longitudinal or panel data, that is data with repeated measurement over time. On the macro-level the method can be used to study cross-national development dynamics, for instance, the relationship between a country's gross domestic product, child mortality and education levels. If regional or city district data is available it is possible to use the method to study for instance neighbourhood segregation processes. On a meso-level the researched entities could be organisations, companies or schools, to study, for instance, dynamic female employment patterns of companies. Finally, the approach is applicable to micro-level data like registered data or panel-data to study social phenomena on the individual level.

We present first a discussion of the statistical method used in our approach in the next section and then give a simple example, applying the method to study the interaction of GDP per capita and democracy for a set of 189 countries from 1980 to 2006. Along with the paper, we present an R package (Bayesian Dynamical System Model, *bds*m, to be found on CRAN <http://cran.r-project.org>) that implements this novel mathematical modeling approach

and that will allow researchers to apply our method to their specific research field. Future predictions and policy suggestions, which are important components for the study of social phenomena, can also be generated using this method and therefore using our R package.

## 2 METHODS

Suppose that we are studying a social system with  $N$  indicator variables  $x_i, i = 1, \dots, N$ . Let us assume that we have longitudinal or panel data for the  $N$  variables for  $M$  entities (such as individuals, countries, organisations etc.) over a length of time  $T$ . Let us denote the data as  $x_i^j(t)$  and the changes in the variables over a time period as  $dx_i^j(t) = x_i^j(t+1) - x_i^j(t)$ , where  $j = 1, \dots, M$  and  $t = 1, \dots, T$ . We use this data to construct what is called a phase portrait of the system.

### 2.1 Phase Portrait

In dynamical systems theory, a phase portrait refers to a plot of the evolution of two variables with respect to each other (Strogatz, 2000). For example, in a system with only two variables  $x_1$  and  $x_2$  (and for only 1 individual, say), the phase portrait would refer to a plot of  $x_2(t)$  against corresponding  $x_1(t)$ . This plot shows the co-evolution of the two variables, and the phase portrait itself can be represented mathematically using Ordinary Differential Equations (ODE) as  $\frac{dx_1}{dt} = f_1(x_1, x_2)$ ,  $\frac{dx_2}{dt} = f_2(x_1, x_2)$  for some appropriate functions  $f_1$  and  $f_2$ . Note that when we have discrete data, we need to use difference equations instead of differential equations. Here we assume that the discrete data are the result of sampling from continuous functions and hence the ODEs represent the same process from which the corresponding discrete data can be obtained by suitable sampling.

Since the differential equations hold for any value of  $x_1$  and  $x_2$ , we could look at all the possible trajectories of the two variables starting at any point. Thus we can think of the available panel data with many individuals as corresponding to the different trajectories obtained in the same system but with *different initial conditions*. Thus in our modeling approach, we look at the data phase portrait, where we look at the changes in the indicator variables  $dx_i(t)$  as a function of the values of all the variables  $\{x_i(t)\}$  (or the current 'state' of the system).

We abstract individual entities as different initial conditions in the system trajectory. In other words, we assume that any individual entity on reaching a certain

‘state’ (represented by a unique set  $\{x_i(t)\}$ ) will experience the same effect, albeit with some additional noise. This approach may of course be problematic in studies that emphasise the differences between different entities and hence their different development trajectories (for example, the economic model in the communist Soviet Union was fundamentally different from that in the United States during a large part of the twentieth century). However the current approach provides a ‘mean-field’ approximation to the basic underlying process in all cases.

## 2.2 Model Selection

In general, if we take  $f_i$  to be polynomial of sufficiently high degree (including products of variables), so that we can model any general non-linearity in the system. For most applications we assume that the functions are polynomial in the indicator variables with each term being of degree  $-1, 0, 1$  in the variables or a product of such terms. We also allow for terms that are quadratic in the variables. This keeps the number of models to evaluate sufficiently small for computational purposes. The terms comprising products of variables capture non-linearities in the system, which can occur due to interactions. These higher order terms can typically result in multi-stable states, which are characteristic of realistic social systems. Moreover, because we include both degree  $-1$  and degree  $2$  terms the resulting models are cubic. In this current study, we assume that any further non-linearities due to degree  $3$  or higher order terms are relatively negligible, but this has to be tested on a case-by-case basis depending on the particular system being modeled. Our R-package provides the option to include order  $3$  polynomial terms.

In our standard implementation of a two variable model, we look at functions containing one or more of the following terms:

$$f_1(x_1, x_2) = a_0 + \frac{a_1}{x_1} + \frac{a_2}{x_2} + a_3x_1 + a_4x_2 + \frac{a_5}{x_1x_2} + \frac{a_6x_2}{x_1} + \frac{a_7x_1}{x_2} + a_8x_1x_2 + a_9x_1^2 + a_{10}x_2^2 + \frac{a_{11}}{x_1^2} + \frac{a_{12}}{x_2^2}$$

There are  $13$  models with one term and, in general,  $\binom{13}{m}$ , models with  $m$  terms that describe the relations between the two variables included in the model.

In the first stage of our fitting process, we aim to rapidly narrow our search by finding the maximum-likelihood model for each possible number of terms,  $m$ . We fit the yearly samples of the yearly changes in the indicator variables using multiple linear regression over all  $8,192$  possible functions  $f_1(x_1, x_2)$  consisting of the polynomial terms shown above. For

each possible number of terms we find the model with the greatest likelihood (equivalently the model that minimises the sum of squared errors with the observed data). We repeat the same process to obtain the best possible  $f_2(x_1, x_2)$  and use the log-likelihood value to rank the different models.

In general, the log-likelihood of the best fit for  $dx_i$  models with  $m$  terms is

$$L_i(m) = \log P(dx_i|x_1, \dots, x_N, m, \phi_{i,m}^*) \quad (3)$$

where  $\phi_{i,m}^*$  is the set of unique parameter values obtained from the best fit regression out of all of the  $\binom{13}{m}$  models with  $m$  terms. Assuming that the actual observations are due to the underlying model with additional Gaussian noise,  $L_i(m)$  is the logarithm of the error sum of squares (ESS) scaled by the variance (Bishop, 2006). The log-likelihood value is also directly related to the coefficient of determination or the  $R^2$  value as  $R^2 = 1 - \frac{ESS}{N_{obs} * \text{Data variance}}$ .

## 2.3 Bayes Factor

An important question about the robustness of particular models is why we choose a particular number of terms. For polynomial function fitting  $L_i(m) \geq L_i(m-1)$ , that is the maximum likelihood is monotonically increasing with additional terms, since each term allows an extra degree of freedom on curve fitting. For a finite data set this extra degree of freedom can fit artifactual patterns due to noise. As a result, reliance on  $L_i(m)$  alone can lead to overfitting the data by selecting too many terms and thus accepting a model that accurately fits the existing data but that generalises poorly to unseen data and has little predictive power.

To address this problem and evaluate the fit of these models we adopt a Bayesian approach. We calculate the Bayesian marginal-likelihood or *evidence*  $B(m)$  for the set of models which have the largest log-likelihood within their respective number of terms. Note that ‘Bayes factor,’ which refers to a *ratio* of model likelihoods is used in Bayesian literature to compare pairs of models (Robert, 1994). We use the same term in this paper to refer to the Bayesian marginal likelihood as defined above, with the understanding that this quantity would have the same function as the Bayes factor when comparing between more than two models.

The Bayes factor compensates for the increase in the dimensions of the model search space by integrating over all parameter values, i.e.,

$$B_i(m) = \int_{\phi_{i,m}} P(dx_i|x_1, \dots, x_N, m, \phi_{i,m}^*) \pi(\phi_{i,m}) d\phi_{i,m} \quad (4)$$

The Bayes factor is thus the likelihood averaged over the parameter space with a prior distribution defined by  $\pi(\phi_{i,m})$ . We choose a non-informative prior distribution (Ley and Steel, 2009). For example,  $\pi(\phi_{i,m})$  can be chosen to be uniform over the range of parameter values. This range of values is chosen to include all feasible values but to be small enough for the integral to be computed using Monte Carlo methods.  $B_i(m)$  is computationally expensive to calculate, even for models with a small number of terms. Therefore we first identify the best fit model for each number of terms using maximum-likelihood, since models of equal complexity can be more fairly evaluated in terms of their maximum likelihood. We then compare those selected in terms of the Bayes factor to fairly compare models of varying complexity.

## 2.4 Correlated Errors

Calculating the best fit regressions for  $dx_i$  independently, as we do above, is equivalent to assuming that the errors in the differential equations are uncorrelated. In fact, there is a possibility that the errors are correlated due to any systematic reason causing the errors, for example the same omitted variable. In social systems this may be more likely the norm than the exception. In this case we have to include an error covariance matrix in our approach and use a generalised least squares approach to finding the regression coefficients. If the error covariance matrix is almost diagonal with off-diagonal elements negligible compared to the diagonal elements, this reduces to the ordinary least squares approach used here.

To test if the errors are in fact significantly correlated, we use the “seemingly unrelated regressions” approach (Amemiya, 1985). For example, in the two variable case, the two regressions for  $dx_1$  and  $dx_2$  are first performed under the assumption that the errors are in fact uncorrelated. We then estimate an error covariance matrix from the model suggested by this first step and the data, and use it to estimate the parameters based on a generalised least squares approach. This process may be iterated until the true parameters are obtained. If the covariance matrix is “almost” diagonal, indicating that error terms are uncorrelated, the parameters estimated by the “seemingly unrelated regressions” approach will not differ significantly from the parameters obtained assuming uncorrelated errors. If not, we have to account for the difference in our calculation of Log-likelihood values and Bayes factor using an algorithm that uses the error covariance matrix in its calculations.

## 2.5 Model Complexity

When generating data-driven models, it is important to have a handle on model complexity. Specifically, in systems with many variables, model complexity is decided both by the number of terms used in the model and by the number of explanatory variables used in each differential equation. For example, in three variable models we would like to determine whether or not we require all of these variables to model the rates of change of each variable. To do this, we calculate Bayes factor for models including all three indicators and compare them to those including just pairs of indicators. For three indicators there are now  $\binom{33}{m}$  models with  $m$  terms and we generally restrict our analysis to those with up to  $m = 5$  terms. By plotting  $B_i(m)$  for three variable models as a function of  $m$  and comparing this to  $B_i(m)$  for two variable models we can assess the utility of adding a third explanatory variable to the model.

Similarly, our algorithm weights all possible models equally and evaluates their log-likelihood and Bayes factor values. But for systems with many variables, there are just too many models available even with the polynomial restriction. This makes the task computationally impossible. To resolve this problem, we can resort to a pruning algorithm which looks at models with increasing number of terms. In each stage, only the top  $M$  models survive, and in the next stage only the ‘descendants’ of these models - models which are the same as the  $M$  survivors from the previous stage except for an additional term - are evaluated. This keeps the number of feasible models evaluated in each step of the algorithm reasonable while a suitable value of  $M$ , say 10,000 will make sure that most feasible models are always tested.

## 3 APPLICATION

To give an example of an application of our approach, let us investigate a frequently studied macro-level phenomenon. Political scientists have been discussing the correlation between a country’s GDP and the level of democracy for 50 years while drawing a wide variety of conclusions (Lipset, 1959; Diamond and Marks, 1992; Barro, 1999; Boix and Stokes, 2003; Krieckhaus, 2003). The correlation between GDP per capita and democracy is usually represented in scatter plots like in Figure 1.

Having longitudinal data allows to estimate various sophisticated panel regression models (Wooldridge, 2010). The most common ones are fixed effect (Allison, 2005) and random effect

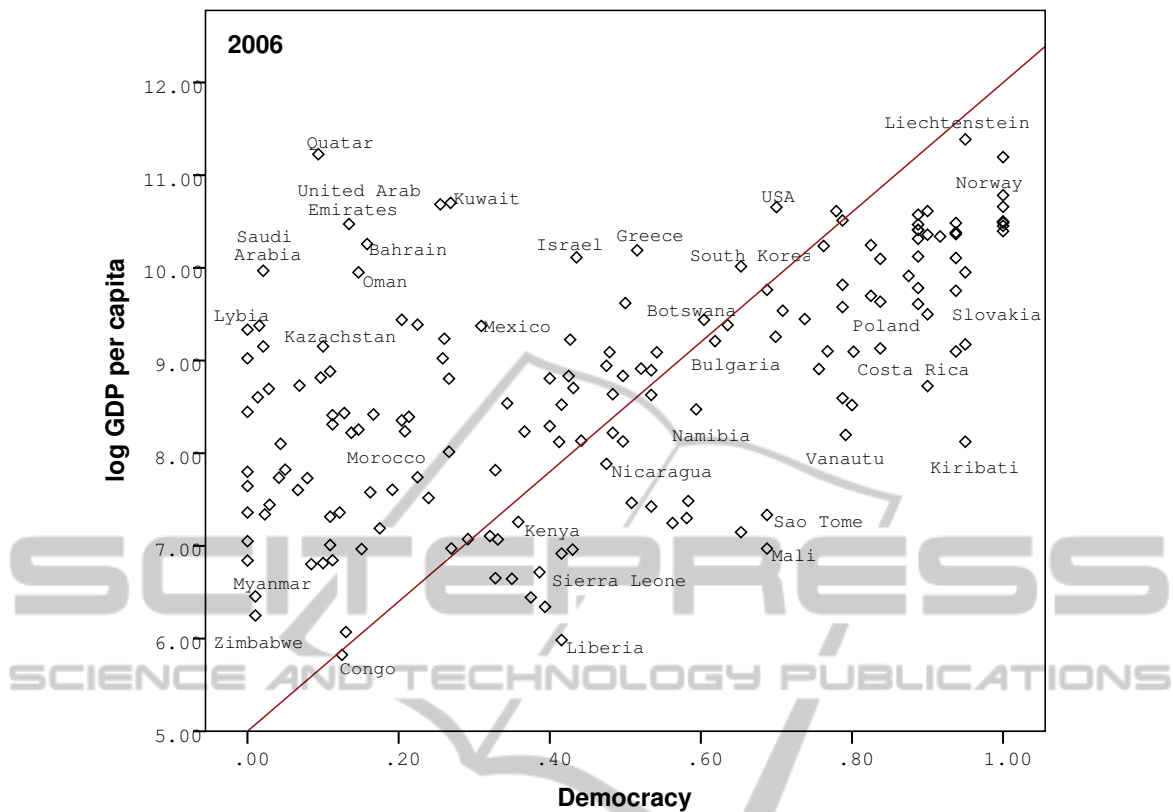


Figure 1: The figure shows a correlation scatter plot for GDP per capita and democracy in the year 2006 for 189 countries. The democracy index is based on Freedom House civil and political rights scores weighted for the actual human rights situation (based on Cingranelli/Richards Human Rights data project) in the respective countries (Welzel, 2013). Most of the outliers at the bottom right are oil-rich Middle East countries with high GDP but low democracy levels.

models (Laird and Ware, 1982). In these regression analyses lagged or difference variables are used as dependent variables to predict the value or the difference of the independent variable at some later point in time (Wooldridge, 2010). Autoregressive (ibid.), two-stage least-square (Garson, 2013) and simultaneous equation models (Wooldridge, 2010) are further elaborated model specifications. There are also non-linear versions panel regression models, like logit regression models (ibid.).

In the analysis of GDP per capita and democracy Barro (1996, 1999) used for instance a panel regression models with roughly 100 countries between 1960 and 1990 with GDP growth rates (difference variables) over three periods (1965-75), (1975-85) and (1985-90) as dependent variables in an instrumental variable estimation approach with amongst others democracy as predictor. He also computed regression models with average democracy levels (1965-74),(1975-84) and (1985-94) with amongst others lagged GDP levels as predictors. Performing these panel analysis, Barro concludes that while democracy has no significant direct effect on GDP per capita

growth, GDP per capita has a significant positive effect on democracy (Barro, 1996).

As suggested by Figure 2 there is a general linear growing tend for both GDP and democracy. However, from this general trend it is difficult to make any reasonable conclusions about the dynamical interaction of economy and democracy. More recent analysis (Boix and Stokes, 2003; Kriekhaus, 2003) indeed suggest that the relation between GDP and democracy might be rather a non-linear and dynamic one. When we create a phase portrait for GDP and democracy the non-linearity and dynamics of their interaction becomes clear (see Figure 3).

Our analysis approach with data from 1980 to 2006 would result in these two best-fit mathematical models for democracy's change as a function of GDP and democracy itself and for GDP's change as a function of democracy and GDP itself:

$$\frac{dD}{dt} = 0.0003G^2 - 0.4031\frac{D}{G} \tag{5}$$

$$\frac{dG}{dt} = 0.0246D + 0.0017G - 0.0002G^2 \tag{6}$$

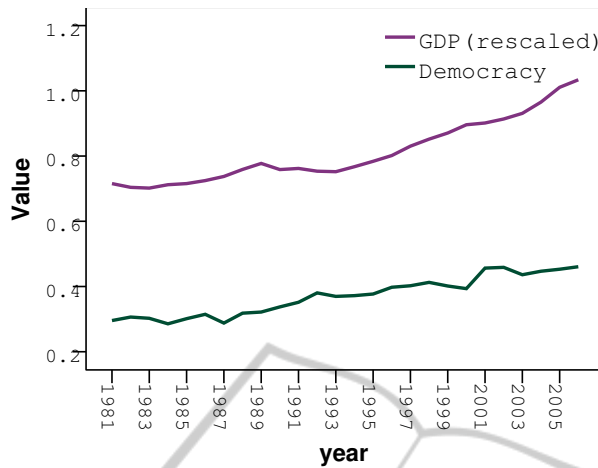


Figure 2: Sequence chart with average (all countries) rescaled GDP and average (all countries) democracy in a time line between 1981 and 2006.

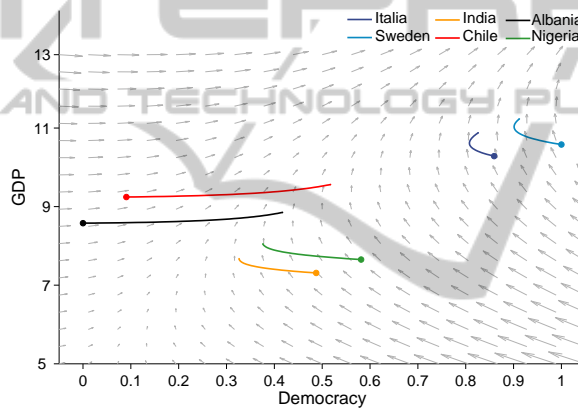


Figure 3: Visualisation of a phase portrait: changes in democracy values (x-axis) against log GDP (y-axis). The vector field shows average change according to the model, while the coloured lines give changes in representative countries as predicted by the model given initial conditions in 1980. Specifically, the ODE model is integrated forward to 2006 for each country starting with the actual initial condition for the corresponding country in 1980. The democracy index is based on Freedom House civil and political rights scores weighted for the actual human rights situation (based on Cingranelli/Richards Human Rights data project) in the respective countries.

The symbiotic interaction of these two variables, economy and democracy, produces an interesting development pattern in the phase-portrait figure (see Figure 3). It seems countries typically head towards, what in dynamical systems is called, a stable manifold. Countries begin either side of this manifold, some with high democracy and low GDP, others with low democracy and high GDP. Over time the countries move to a common trajectory moving from bottom left to top right of the phase plane. These results can explain the sometimes apparently contradictory patterns previously seen in relating GDP and democracy. If a country starts with high democratic levels but a GDP that is rather low, the democratic level is unstable and regresses to the point where it reaches

the stable manifold that then allows both democracy and GDP to grow again.

## 4 CONCLUSIONS

Our method provides social scientists with a tool to study complex and dynamic phenomena. Unlike classic panel analysis, where a decision is usually made to study a particular time frame, our methods takes in to account all of the available temporal data. In the application example, we are able to capture dynamical interplays of variables. We expect that the method will be able to detect more complex phenomena, such as amplification, growth limitation, glacial effects or

tipping effects. The analysis procedure results in best-fit models that explicitly depict precise and dynamic mechanisms. Equations such as 5 and 6 provide the researcher with rich information beyond correlation coefficients, since they express how variables change with respect to each other's state. In future studies, we will show how the same mechanisms can be used to look at three and more variables (Ranganathan et al., 2013).

A key feature of our approach is that no predefined model is imposed on the data. Instead the data itself is used to find the best model. The same approach of calculating Bayes factor can of course be used to test theoretically informed model specifications. Such testing can tell us how the best fit data-driven model compares in terms of statistical fit, to a model based on theoretical reasoning. There may well be strong grounds to accept a theoretically justified model with a slightly worse fit, over a purely data-driven model with the best fit. Indeed, we do not suggest that social scientists should forget about theories and always adopt the statistically best models. No doubt, theories are useful to interpret results and to evaluate models. But we think that social scientists should be equally open to finding meaningful patterns and mechanisms beyond established theories. If the detected patterns and models are plausible and help to understand social reality or give a new insight into a phenomenon, then even new theoretical mechanisms could be formulated or older theoretical mechanisms revised, based on these findings.

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