ENTROPIC ANALYSIS AND SYNTHESIS OF BIOSIGNAL COMPLEXITY

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Abstract: Analysis of complexity of biological time-series data is investigated to gain knowledge about the biosignal predictability. Using modern biological data such as mass spectral, this complexity information can be utilized to identify novel biomarkers for drug discovery, early disease detection and therapeutic treatment. To enhance the complexity analysis, a probabilistic fusion scheme, which is an alternative to the assumption of the independence of probabilistic models, is applied to synthesize the information given by different entropy methods.

1 INTRODUCTION

The notion of complexity can be defined as a scientific study of systems which change irregularly over time or space (Havel, 1995). Thus, understanding the behaviors of dynamical systems in terms of predictability is a key purpose of the study of complexity. There are several new perspectives developed on the study of complexity in the physical and natural sciences over last few decades. Theories such as nonlinear dynamic systems, self-organization, catastrophe, self-organized criticality, antichaos, and chaos appear to offer novel perspectives on the long-standing problems of developing scientific measures of information in the specific domain from which they emerge. Depending on a particular discipline, these methods for studying complexity have been characterized as constituting everything from a major paradigm shift which challenges established scientific beliefs to the refinement of current methodology (Sprott, 2003).

An entropy-based measure of systems complexity known as approximate entropy (ApEn) (Pincus, 1991) and its extended family - sample entropy (SampEn) (Richman and Moorman, 2000) and multiscale entropy (MSE) (Costa et al, 2002) have been recently proposed to quantify the complexity of physiological and biological data. A low value of the approximate entropy indicates the time series is deterministic (low complexity); whereas a high value indicates the data is subject to randomness (high complexity) and therefore difficult to predict. In other words, lower entropy values indicate more regular time series; whereas higher entropy values indicate more irregular time series. Both ApEn and SampEn estimate the probability that the sequences in a dataset which are initially closely related remain closely related, within a given tolerance, on the next incremental comparison. ApEn differs from SampEn in that its calculation involves counting a self-match for each sequence of a pattern, which leads to bias in ApEn (Pincus and Goldberger, 1994). SampEn is precisely the negative natural logarithm of the conditional probability that two sequences similar for m points remain similar at the next point, where self-matches are not included in calculating the probability. Thus a lower value of SampEn also indicates more self-similarity in the time series. In addition to eliminating self-matches, the SampEn algorithm is simpler than the ApEn algorithm, requiring approximately one-half as much time to calculate. SampEn is largely independent of record length and displays relative consistency under circumstances where ApEn does not (Richman and Moorman, 2000).

It has been pointed that ApEn suffers from two major drawbacks (Lewis and Short, 2007): (1) because it is a function of the length of the sequence under study, it yields entropy values lower than expected for short sequences because its calculation involves counting a self-match for each sequence and this leads to bias (Pincus and Goldberger, 1994) ; (2) it can be inconsistent with different testing conditions using different parameters of the entropy index. Sam-

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pEn does not count self-matches and therefore can reduce bias. It has been found that SampEn can provide better relative consistency than ApEn because it is largely independent of sequence length (Richman and Moorman, 2000). MSE measures complexity of time series data by taking into account multiple time scales, and uses SampEn to quantify the regularity of the data. All of these three methods depend on the selection of the two parameters known as m and r: parameter m is used to determine the sequence length, whereas parameter r is the tolerance threshold for computing pattern similarity. Results are sensitive to the selections of these two parameters and it has recently been reported that good estimates of these parameters for different types of signals are not easy to obtain (Lu et al, 2008). In this paper we introduce a new entropy method called GeoEntropy (GeoEn) which can provide an analytical procedure for estimating the conrtol parameter r. We then apply various entropy methods to study the complexity or predictability of cancer using mass spectrometry data, which are complex and large datasets. To improve the entropy analysis, we use a novel probabilistic fusion framework based on the engineering hypothesis of permanence of ratio to combine the results from different entropy algorithms.

1.1 GeoEntropy

Let $z(X)$ be a regionalized variable which has characteristics in a given region $\mathcal{D}$ of a spatial or time continuum (Matheron, 1989). In the setting of a probabilistic model, a regionalized variable $z(X)$ is considered to be a realization of a random function $Z(X)$. In such a setting, the data values are samples from a particular realization $z(X)$ of $Z(X)$. We now consider n observation: $z(X_\alpha)$, $\alpha = 1,\ldots,I$; taken at locations or times α. If the objects are points in time or space, the possibility of infinite observations of the same kind of data is introduced by relaxing the index α. The regionalized variable is therefore defined as $z(X)$ for all $X \in \mathcal{D}$, and $\{z(X_\alpha), \alpha = 1,\ldots,I\}$ is viewed as a collection of a few values of the regionalized variable.

We now consider that each measured value in the dataset has a geometrical or time point in the respective domain $\mathcal{D}$, which is called a regionalized value. The family of random variables $\{Z(X), X \in \mathcal{D}\}$, is called the random function. The variability of a regionalized variable $z(X)$ at different scales can be measured by calculating the dissimilarity between pairs of data values, denoted by $z(X_\alpha)$ and $z(X_\beta)$, located at geometrical or time points α and β in a spatial or time domain $\mathcal{D}$, respectively (from now on we address point/domain to imply either geometrical or time point/domain). The measure of this semi-dissimilarity, denoted by $\gamma_{\alpha\beta}$, is computed by taking half of the squared difference between the pairs of sample values (the term *semi* is used to indicate the half difference) as

$$\gamma_{\alpha\beta} = \frac{1}{2}(X_\alpha - X_\beta)^2 \tag{1}$$

The two points x_α and x_β in space or time can be linked by a space or time lag $h = X_\alpha - X_\beta$ (we use h here as a scalar but its generalized form is a vector to indicate various spatial orientations). Now let the semi-dissimilarity depend on the lag h of the point pair, we have

$$\gamma_\alpha(h) = \frac{1}{2}[(z(X_\alpha + h) - z(X_\alpha)]^2 \tag{2}$$

Using all samples pairs in a dataset, a plot of the $\gamma(h)$ against the separation h is called the semi-variogram. The function $\gamma(h)$ is referred to as the semi-variance and defined as

$$\gamma(h) = \frac{1}{2N(h)} \sum_{(\alpha,\beta)|h_{\alpha\beta}=h} [z(X_\alpha) - z(X_\beta)]^2 \tag{3}$$

where $N(h)$ is the number of pairs of data points whose locations are separated by lag h.

The semi-variance defined in (3) is known as the experimental semi-variance and its plot against h is called the experimental semi-variogram, to distinguish it from the theoretical semi-variogram that characterizes the underlying population. The theoretical semi-variogram is thought of a smooth function represented by a model equation; whereas the experimental semi-variogram estimates its form. The behavior of the semi-variogram can be graphically illustrated by the theoretical semi-variogram using the spherical or the Matheron model which is defined as (Isaaks and Srivastava, 1989)

$$\gamma(h) = \begin{cases} s\left[1.5\frac{h}{g} - 0.5(\frac{h}{g})^3\right] & : \quad h \le g \\ s & : \quad h > g \end{cases} \tag{4}$$

where g and s are called the *range* and the *sill* of the theoretical semi-variogram, respectively.

The concept of regionalized variables and its modeling of variability in space continuum by means of the semi-variogram have been described. What can be observed is that the range g of the semi-variogram presents an idea for capturing the auto-relationship of the time-series data: within the range g, the data points are related; when $h > g$, information about relationship between the data points becomes saturated and not useful. Based on this principle of the

semi-variogram, the length of the sub-sequences of the time-series data X can be appropriately chosen to be the range g, which ensures an optimal study of self-similarity of the signal, that is $m = g(X)$. To address the criterion of similarity/dissimilarity between the sub-sequences, we can establish its lower bound as the absolute difference between two consecutive interval of the semi-variance or the absolute one-step semi-variance difference: $r = |\gamma(h) - \gamma(h+1)|$, or multi-step semi-variance difference: $r = |\gamma(h) - \gamma(h+c)|$ where c is a positive constant. Having defined the subsequence length m and the similarity criterion r, determination of GeoEn can be obtained using the principle of either ApEn or SampEn.

GeoEn algorithm for calculating the complexity of time-series data is outlined as follows (Pham, 2009).

1. Compute the semi-variance of X_N and its range $g(X_N)$
2. Set vector length $m = g(X_N)$
3. Construct vectors of length m, X_1 to X_{N-m}, defined as
$$X_i = (x_i, x_{i+1}, \ldots, x_{i+m-1}),\ 1 \le i \le N-m$$
4. Set semi-variance lag $h = 1, \ldots, \min_i[g(X_i)]$
5. Compute distance between X_i and X_j as
$$d(X_i, X_j) = |\gamma_{X_i}(h) - \gamma_{X_j}(h)| \quad (5)$$
6. Set the criterion of similarity r as follows.
$$r = |\gamma_{X_i}(h) - \gamma_{X_i}(h+1)| \quad (6)$$
7. Calculate either ApEn or SampEn for each h to obtain multiscale $GeoEn(X_N, h)$.

2 PROBABILISTIC FUSION

Based on the engineering paradigm of the permanence of updating ratios, which asserts that the rates or ratios of increments are more stable than the increments themselves, as an alternative to the assumption of the full or conditional independence of probabilistic models; Journel introduced a scheme for information fusion of diverse sources (Journel, 2002). This scheme allows the combination of data events without having to assume their independence. This information fusion is described as follows.

Let $P(A)$ be the prior probability of the occurence of data event A; $P(A|B)$ and $P(A|C)$ be the probabilities of occurence of event A given the knowledge of events B and C, respectively; $P(B|A)$ and $P(C|A)$ the probabilities of observing events B and C given A, respectively. Using Bayes' law, the posterior probability of A given B and C is

$$P(A|B,C) = \frac{P(A,B,C)}{P(B,C)} = \frac{P(A)P(B|A)P(C|A,B)}{P(B,C)} \quad (7)$$

The simplest way for computing the two probabilistic models is to assume the model independence, giving

$$P(C|A,B) = P(C|A) \quad (8)$$

and

$$P(B,C) = P(B)P(C) \quad (9)$$

Thus, (7) can be rewritten as

$$\frac{P(A|B,C)}{P(A)} = \frac{P(A|B)}{P(A)}\frac{P(A|C)}{P(A)} \quad (10)$$

However, the assumption of conditional independence between the data events usually does not statistically perform well and leads to inconsistencies in many real applications (Journel, 2002). Therefore, an alternative to the hypothesis of conventional data event independence should be considered. The permanence of ratios based approach allows data events B and C to be incrementally conditionally dependent and its fusion scheme gives

$$P(A|B,C) = \frac{1}{1+x} = \frac{a}{a+bc} \in [0,1] \quad (11)$$

where

$$a = \frac{1-P(A)}{P(A)},\ b = \frac{1-P(A|B)}{P(A|B)},$$

$$c = \frac{1-P(A|C)}{P(A|C},\ x = \frac{1-P(A|B,C)}{P(A|B,C)}.$$

An interpretation of the fusion expressed in (11) is as follows. Let A is the target event which is to be updated by events B and C. The term a is considered as a measure of prior uncertainty about the target event A or a distance to the occurrence of A without any updated evidence. The term x is the distance to the target event A occurring after observing evidences given by both events B and C. The ratio c/a is then the incremental (increasing or decreasing) information of C to that distance starting from the prior distance a. Similarly, the ration x/b is the incremental information of C starting from the distance b. Thus, the permanence of ratios provides the following relation

$$\frac{x}{b} \approx \frac{c}{a} \quad (12)$$

which says that the incremental information about C to the knowledge of A is the same after or before

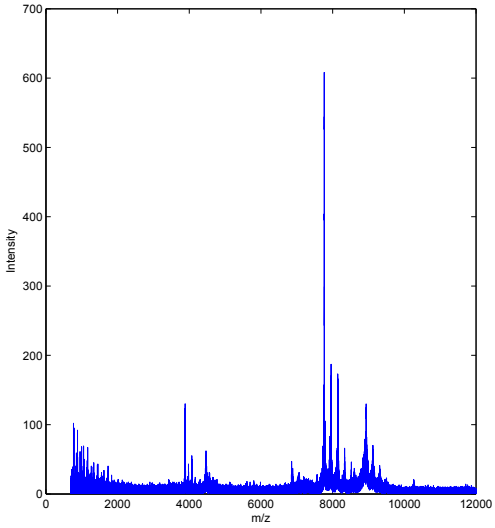

Figure 1: Ovarian mass spectrometry data: disease sample.

knowing B. In other words, the incremental contribution of information from C about A is independent of B. This expression relaxes the restriction of the assumption of full independence of B and C.

For the generation of k data events E_j, $j = 1,\ldots,k$; the conditional probability provided by a succession of $(k-1)$ permanence of ratios is given as

$$P(A|E_j,\ j=1,\ldots,k) = \frac{1}{1+x} \in [0,1] \qquad (13)$$

where

$$x = \frac{\prod_{j=1}^{k} d_j}{a^{k-1}} \geq 0$$

$$a = \frac{1-P(A)}{P(A)}$$

$$d_j = \frac{1-P(A|E_j)}{P(A|E_j)},\ j = 1,\ldots,k$$

It is clear that expression (13) requires only the knowledge of the prior probability $P(A)$, and the k elementary single conditional probabilities $P(A|E_j)$, $j = 1,\ldots,k$, which can be independently computed.

3 EXPERIMENTAL ANALYSIS AND SYNTHESIS OF BIOSIGNAL COMPLEXITY

We used a public MS-based ovarian cancer dataset, the ovarian high-resolution SELDI-TOF, to carry out

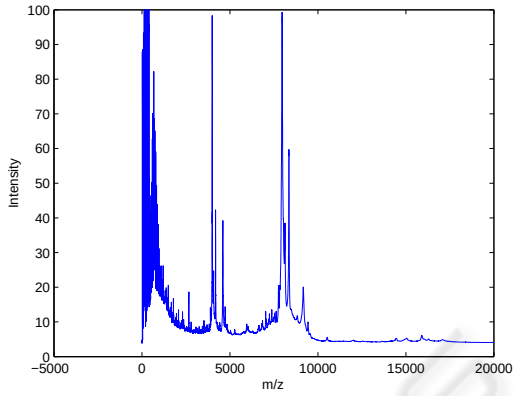

Figure 2: Ovarian mass spectrometry data: control sample.

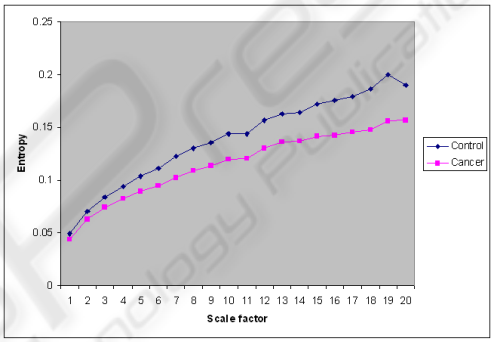

Figure 3: Mean entropy values of cancer and control groups using MSE.

the entropic analysis and synthesis. The dataset was obtained from the FDA-NCI Clinical Proteomics Program Databank. The ovarian cancer data consist of 100 control samples and 170 cancer samples. The length of each sample is 15,154 m/z values. Figures 1 and 2 show the plots of typical ovarian, and control samples, respectively.

Mass spectrometry (MS) in proteomics has been used to study the regulation, timing, and location of

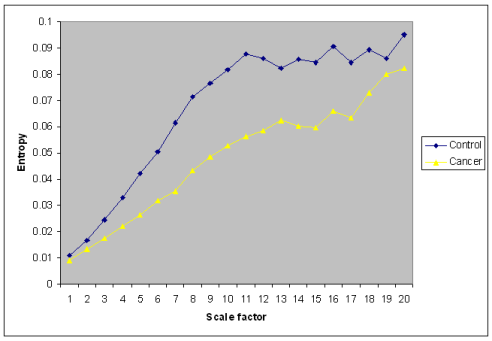

Figure 4: Mean entropy values of cancer and control groups using MSE with m values of GeoEn.

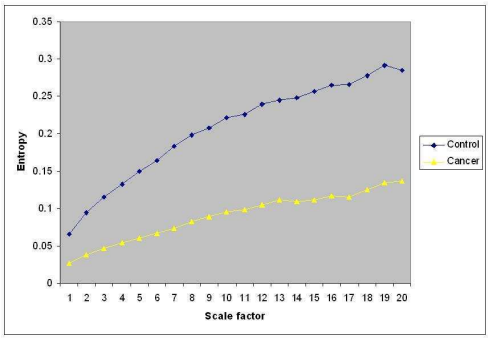

Figure 5: Probabilistic combination of entropy values of cancer and control groups.

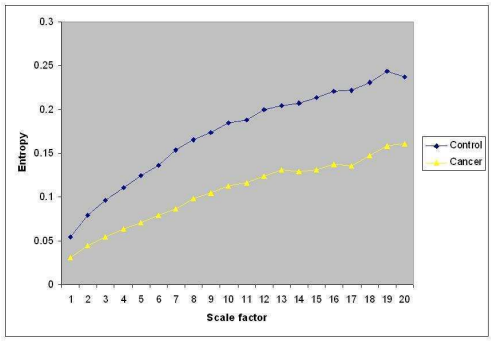

Figure 6: Average combination of entropy values of cancer and control groups.

protein expression. Interaction studies seek to understand how protein pair between themselves and other cellular components interact to constitute to more complex models of the molecular machines. In particular, protein expression profiles or expression proteomics can be used for large-scale protein characterization or differential expression analysis that has many applications such as disease classification and prediction, new drug treatment and development, virulence factors, and polymorphisms for genetic mapping, and species determinants (Adam et al, 2001). In comparison with transcriptional profiling in functional genomics, proteomics has some obvious advantages in that it provides a more direct approach to studying cellular functions because most gene functions are characterized by proteins. Current study on MS data concerns with peak detection for biomarker discovery and pattern classification for disease prediction. In this study, we examined the complexity of this type of MS data and applied the proposed classification scheme to classifying cancer and control samples and compared the performance of the proposed methods with other methods.

Applying the experimental semi-variogram, we obtained the range about 20 for both cancer and control populations and set this value to be the value m for the entropy calculation. For the entropy estimates using MSE, the constant values of m and r are 2 and 0.15, respectively. Figure 3 shows the plots of the mean entropy values of the cancer and control groups using MSE. Some difference of MSE values between the cancer and control groups can be observed; however, the entropy values of the two groups increase with increasing scales. Figures 4 shows the plots of the mean entropy values of the cancer and control groups using MSE procedure with the values of m obtained from GeoEn approach. The entropy profiles of the cancer and control populations obtained from MSE-based GeoEn can distinguish the complexities between the groups. The entropy values of the control group take higher values than those of the cancer. Both cancer and control populations tend to increase with increasing values of h.

To enhance the complexity analysis, we applied the permanence-of-ratio fusion to combine the two updating results obtained from MSE and GeoEn given the prior results estimated from SampEn. In this fusion, $P(A)$ is the prior probability of the complexity of the data given by SampEn, $P(A|B)$ and $P(A|B)$ are the probabilities of the complexity obtained from MSE and GeoEn, respectively, given the knowledge provided by SampEn. These three defined probabilities are ready to calculate the three parameters a, b and c to estimate the updated probability $P(A|B,C)$ defined in (11). Figure 5 shows the fused entropy values. It can be observed that the separation of the complexity profiles becomes better in comparisons with the profiles produced separately by MSE and GeoEn. Figure 6 shows the entropy profiles taken as the average of the two results given by MSE and GeoEn. It can be seen that combination by averaging does not yield better results than the probabilistic fusion scheme.

4 CONCLUSIONS

Approximate entropy was introduced for the analysis of short time series. Sample entropy was developed as a modified version of ApEn to offer the advantage of some independence on time series length. More recently, multiscale entropy has been introduced by averaging time series data with different intervals or scales which are then analyzed by sample entropy. All of these three entropy-based algorithms rely on the heuristic estimates of m and r. The GeoEn approach offers an analytical procedure for the estimation of these parameters. Data fusion has been widely used for improving results from multiple sources of information. This study has combined advantages of entropy-based methods by means of a data fusion

method. Furthermore, what we have reported is the application of a most recently developed data combination scheme which does not impose the strong independent assumption of probabilistic models. The approaches studied herein can be applied to many types of biosignals for early disease prediction and biomarker discovery where the entropy profiles can be used as novel features in pattern classification process.

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