TWO DOF CAMERA POSE ESTIMATION WITH A PLANAR STOCHASTIC REFERENCE GRID

Giovanni Gherdovich and Xavier Descombes
INRIA Sophia Antipolis, 2004 route des Lucioles, Sophia Antipolis, France

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Abstract: Determining the pose of the camera is a need to many higher level computer vision tasks. We assume a set of features to be distributed on a planar surface (the world plane) as a Poisson point process, and to know their positions in the image plane. Then we propose an algorithm to recover the pose of the camera, in the case of two degrees of freedom (slant angle and distance from the ground). The algorithm is based on the observation that cell areas of the Voronoi tessellation generated by the points in the image plane represent a reliable sampling of the Jacobian determinant of the perspective transformation up to a scaling factor, the density of points in the world plane, which we demand as input. In the process, we develop a transformation of our input data (areas of Voronoi cells) so that they show almost constant variances among the locations, and analytically find a correcting factor to considerably reduce the bias of our estimates. We perform intensive synthetic simulations and show that with few hundreds of random points our errors on angle and distance are not more than few percents.

1 INTRODUCTION

With “camera pose estimation” one refers to the problem of determining the position and orientation of a photo camera with respect to the coordinate frame of the imaged scene. When acquiring this information by external instruments is too expensive for the application of interests, or simply is impossible because the picture is already taken, one must resort to computer vision techniques and use at best the visual data at disposal. It is a task that needs to be performed in a wide range of different situations, subsumed by many higher level computer vision problems like object detection, object recognition, vision-based safety applications, augmented reality. Anytime one needs to measure metric or affine quantities in a 3D scene captured by a photo camera, the parameters of the perspective projection need to be recovered, i.e. information about the camera position have to be inferred from the image itself.

In the case the images containing artifacts like buildings or other non-natural structures, classical approaches consist in finding straight lines, known angles, orthogonalities or reference points and use them to invert partially or totally the perspective distortion (Hartley and Zisserman, 2004). Images of natural scenes don’t offer such references, and other approaches need to be investigated. In recent years, researchers have proposed approaches that leverage the statistical properties of the viewed scene. The main idea is to study how these properties are modified in the image by the perspective distortion, so that the desired information about the parameters of the perspective itself can be estimated. Shape from texture is a well established techniques that uses the perspective distortion of some homogeneous or isotropic pattern to get 3D clues about a surface shape (Permuter and Francos, 2000), (Malik and Rosenholtz, 1997), (Clerc and Mallat, 2002), (Gårding, 1992), (Kanatani and Chou, 1989).

Although the problem of estimating the position and orientation of the camera has received a considerable amount of attention, little has been done in estimating the camera pose using features uniformly distributed in a picture. Our approach leverages some of the intuitions foundational to shape from texture techniques; namely we use the concept of area gradient (Gårding, 1992) to determine the nature of the perspective transform. The main objective is to exploit the information content of the location of uniformly distributed features, which can actually come from preprocessing homogeneous image textures, but they can also correspond to the spatial locations of objects whose background is not a homogeneous tex-
ture. This differentiates with the *shape from texture* paradigm which relies on the presence of homogeneity or isotropy on a whole patch of the image. Therefore, the information embedded in the image required to estimate the camera pose is lesser, leading to a wider application area. Because of those reasons in the present work we suggest the decoupling, in the spirit of (Kanatani and Chou, 1989), between image processing and distortion analysis in the task of recovering the orientation of a planar surface subjected to perspective transformation.

In this paper we assume the viewed scene to be a planar surface on which a Poisson point process takes place, i.e. the points are uniformly distributed on such “world plane”; our technique first measures the perspective distortion induced on this pattern by the transformed size of small areas surrounding each point in the resulting image. Then this measure is linked to the parameters of the perspective transformation we’re interested in, i.e. the slant angle and the distance from the scene along the optical axis. In particular, we model the intuitive concept of “small areas” using the size of the cells in the Voronoi tessellation generated by the point pattern. We do that through the following observation: the size of the Voronoi cell centered at the point \( p \) in the image divided by the size of a typical cell of the Poisson-Voronoi tessellation in the world plane yields a reasonable approximation of the Jacobian determinant of the perspective transform computed in the back-image of the point \( p \). Our simulations confirm this intuitive claim: with this procedure, given the density \( \lambda \) of the Poisson process we get consistent estimation for the slant angle and the distance from the ground.

The main contribution of this work is the tailoring of the *area gradient* concept to the case of a discrete set of points through the use of Voronoi tessellations. In the process, we develop a transformation of our input data (areas of Voronoi polygons) so that they show almost constant variances among the different locations in the image plane, and analytically find a correcting factor for the linear model to which we fit such data in order to have an unbiased linear least squares estimation of the parameters of interest, slant angle and distance.

The remaining of the paper is organized as follows: in Section 2 we present the homography we will consider during the subsequent sections; in Section 3 we define formally the problem, then we introduce our algorithm in Section 4 and present the results of our experiments in Section 5. Our conclusions will be given in Section 6.

## 2 THE HOMOGRAPHY UNDER STUDY

In this work we consider homographies between planes, i.e. our perspective transformation, induced by an ideal pinhole camera, will be a function \( P: \mathbb{P}^2 \to \mathbb{P}^2 \) where \( \mathbb{P}^2 \) is the projective plane; \( P \) is represented in homogeneous coordinates by a \( 3 \times 3 \) matrix, defined up to a multiplicative factor. We name *world plane* the domain of \( P \), and *image plane* its codomain. The *principal ray* is the line from the camera center perpendicular to the image plane and the *principal point* is its intersection with the image plane. In a completely general settings, we have eight degrees of freedom (DOF) over \( P \); we restrict our attention to the case of two DOF, namely the distance \( d \) between the camera center and the inverse image of the principal point through \( P \), and the angle \( \theta \) between the principal ray and the world plane (see figure 1). The remaining parameters are reference frames for the two coordinate systems and orientation of the world plane with respect to the image plane, and they are chosen in order to send the vanishing line of the world plane to an horizontal line in the image; the focal length \( f \) is assumed to be known, the pixels are squared and there is no skew.

Let us consider the origins of the world and image coordinate systems are the inverse image of the principal point and the camera center respectively then expression of \( P \) is the following:

\[
P : \mathbb{P}^2 \to \mathbb{P}^2 \quad \begin{bmatrix} u \\ v \\ w \end{bmatrix} \mapsto \begin{bmatrix} f & 0 & 0 \\ 0 & f \sin \theta & 0 \\ 0 & \cos \theta & d \end{bmatrix} \begin{bmatrix} u \\ v \\ w \end{bmatrix}
\] (1)
For the sake of clarity we remark its behavior on the affine part:
\[
P_{a,b}(u,v) = \left( \frac{uvf}{v \cos \theta + d} \right) \left( \frac{v f \sin \theta}{v \cos \theta + d} \right)
\]  
(2)

3 PROBLEM STATEMENT

Suppose that a certain number of points are uniformly distributed on some portion of the world plane, and that we look at them through our pinhole camera, whose model is described in Section (2). Then, their spatial distribution is not uniform anymore in the image plane. We assume to have some good algorithm to detect their position in the picture, and our aim is to recover the parameters \( \theta \) and \( d \) of the planar homography studying such distortion. As stated in Section (2) the “horizon” is assumed to be parallel to the horizontal axes of the image, but it can be very well outside the picture. In our framework, the density of points \( \lambda \) in the world plane (i.e. the number of points per unit area) needs to be known — it’s a consequence of the formula (2), where \( P_{a,b}(u,v) = P_{a,b}(u,v) \) for every \( s \in \mathbb{R} \), which means that changing the distance \( d \) is equivalent to change scale in the world plane, and just looking at the picture we cannot distinguish between small distance with dense points and big distance with sparse points.

4 THE ALGORITHM

In order to study the perspective distortion of uniformly distributed points, one need to capture the following intuitive notion: points get closer to each other as approaching to the horizon. What is needed is a statistical quantity able to discriminate between different perspective projections; our suggestion is to measure “how much free space” \( S_p \) is present around each point \( p \) of our random configuration. Suppose that we were able to know \( S_p \) in the world plane for a given \( p \), and also its transformation, denoted by abuse of notation \( P(S_p) \); if \( S_p \) was small enough, i.e. if the points where sufficiently dense, the ratio of areas \( |P(S_p)|/|S_p| \) would have been a good estimation of the determinant of the Jacobian matrix of \( P \) at the point \( p \) — the Jacobian determinant measures the factor with which a function modifies volumes around a point. And doing this for all the points of the configuration, one can have many samples of the Jacobian, hopefully enough to do a regression and estimate the parameters of interest \( d \) and \( \theta \).

But what does “free space around a point” means? And what is \( |S_p| \)? We don’t have any clue about the world plane, we just have its perspective view. Again: what is \( P(S_p) \)? We don’t even know the function \( P \).

A reasonable answer to the first question is given by the 2-dimensional Voronoi diagram, a tessellation of the plane generated by a set of points \( \{p_i\} \) such that a point \( q \) belongs to the cell of \( p_i \) if it’s closer to \( p_i \) than to any other \( p_j \); small Voronoi cells mean that the generating points are “dense”. So for us the free space around a point is its Voronoi cell. The answer to the second question is \( 1/\lambda \), where \( \lambda \) is the density of points in the world plane. In fact, this is the expected value for \( |S_p| \), assuming a Poisson distribution for the points (Hayen and Quine, 2002); the key point is its independence from \( p \), which can be intuitively understood observing that if we take some region \( A \) in the world plane, the expected value of the number of points inside \( A \) is proportional to the area of \( A \), no matter of its location. In the third question we ask how to approximate the projection of the Voronoi cell \( S_p \), our answer is to compute the Voronoi diagram generated by the projected points.

Before stating precisely our algorithm we write the formula of the Jacobian determinant (from now on just “Jacobian”) of our homography, using the same notation as in eq. (2)
\[
J_p(u,v) = \begin{vmatrix} x & y \\ u & v \end{vmatrix} = \frac{f^2d\sin\theta}{(v \cos \theta + d)^2}
\]  
(3)

and point out that since all our measurements are done in the image plane, while the domain of the above Jacobian is the world plane, what we are going to sample is the composition
\[
(J_P \circ P^{-1})(x,y) = \frac{(f\sin\theta - v \cos \theta)^3}{(f \sin \theta)^2} = (a + by)^3 
\]  
(4)

At this point, the natural choice to recover the parameter \( a \) and \( b \), hence \( d \) and \( \theta \), would be to set the linear regression model
\[
\lambda|P(S_p)| = (a + by)^3 + \epsilon 
\]  
(5)

where the zero mean noise is taken into account by the random variable \( \epsilon \). Unfortunately, the variance of \( \epsilon \) varies with the location of the cell \( S_p \); furthermore, it is reasonable to assume that such variance is transformed under perspective likewise areas, i.e.
\[
\text{Var}(\epsilon) = (\lambda(a + by)^3)^2 \text{Var}(S_p) 
\]  
(6)

This means that knowing \( \text{Var}(\epsilon) \) is equivalent to know the parameters \( a \) and \( b \) that we’re about to estimate. Any heuristic we could use to estimate the variances of the errors \( \epsilon \) will result in a poor fitting to the 3rd
degree polynomial (4); also, in weighting the least squares summation with empirical variances we experimented numerical problems due to very small condition numbers of the matrices involved. For this reason we resort to a variance-stabilizing transformation; by the error propagation formula (G.Cowan, 1998) and given the assumption (6), the ideal candidate would be the logarithm, but the regression wouldn’t be linear any more. We found experimentally that the cube root of our data \( \{\lambda_i P(S_p)\}_p \) shows approximately constant variances.

Thus we don’t use model (5) for our regression, but
\[
\sqrt[3]{\lambda_i P(S_p)} = (a + by) + \eta \tag{7}
\]
With a second order Taylor expansion of the left-hand-side around \((a + by)^3\) one can show that \(E(\eta) = -\lambda^2(a + by)\text{Var}(S_p)/9\) — interestingly, the mean of the noise is proportional to \(a + by\). A closed form for the variance \(\text{Var}(S_p)\) of cell sizes for a Poisson tessellation is not known, but the six decimal digits approximation 0.280176/\(\lambda^2\) found in (Hayen and Quine, 2002) is more than enough for our purposes. Thus we can restate (7) in a suitable form for the ordinary least squares method, i.e. with an error term \(\hat{\eta}\) that has zero mean and constant variance:
\[
\sqrt[3]{\lambda_i P(S_p)} = (a + by) \left(1 - \frac{0.280176}{9}\right) + \hat{\eta} \tag{8}
\]

What follows is the detailed algorithm we use.

1. INPUT: the points \(\{p_1, \ldots, p_n\}\) in the image plane.
2. Generate the Voronoi diagram from \(\{p_1, \ldots, p_n\}\).
3. Points close to the boundaries of the viewed scene will produce degenerate cells, i.e. unbounded or very oblong. Remove all the cells which have at least one vertex outside the viewed scene.
4. Compute the areas of the remaining cells \(\{C_1, \ldots, C_k\}\), where \(C_i\) is the cell generated by the point \(p_i\) (we reordered the cells so that the first \(k\) are the ones we keep). These are \(k\) noisy samples of the function \(J_P \circ P^{-1}\) up to the (known) scaling factor \(\lambda\), in the sense that \(\text{area}(C_i) \approx J_P \circ P^{-1}(p_i)/\lambda\).
5. Solve the linear least squares problem
\[
\min_{a,b} \sum_i \left\{ \sqrt{\lambda \text{area}(C_i)} - (a + y_ib) \left(1 - \frac{0.280176}{9}\right) \right\}^2 \tag{9}
\]
where \(y_i\) is the \(y\)-coordinate of the point \(p_i\).
6. OUTPUT: the estimates
\[
\hat{\theta} = \arctan \left( \frac{a}{\sqrt{b^2 + a^2}} \right) \quad \hat{d} = \frac{f}{a \sqrt{b^2 + a^2}} \tag{10}
\]

5 EXPERIMENTAL RESULTS

We evaluated our algorithm on a set of random point configurations synthetically generated; the parameters of interest that we vary are the slant angle \(\theta\) of the camera and the number of points in the configuration; we keep fixed the length \(d\) of the principal ray to 100 meters. We recall that at a given \(\theta\), adding points to the configuration while keeping \(d\) constant is equivalent to keep the point density constant and increase \(d\). We tested the algorithm with 12 angles \(\theta\) ranging uniformly from 2\(^\circ\) to 60\(^\circ\) and with the number of points ranging uniformly from 100 to 2000; for every of these combinations we generated one hundred random point configurations. During the simulation we kept the focal length \(f\) fixed to 50mm, and our pictures are squares of 25mm \(\times\) 25mm centered in the origin of the image plane.

Figure 2 reports our results concerning the estimation of the distance \(d\). It shows 4 plots, corresponding to some of the values of \(\theta\) for which we tested the algorithm. Each plot contains columns representing a boxplot and a 95% confidence interval for ten different quantities of points in the image, from 100 to 2000. For every of these quantities we simulated 108 different point patterns; the boxplots summarize those measures, with median, lower and upper quartile, maximum and minimum observed value, outliers. The confidence interval on the right of each boxplot refers the empirical mean \(E(d)\) and is based on the Student’ \(t\) distribution. Figure 3 is closely related to figure 2, since it shows the relative errors of the empirical mean \(|E(d) - d|/d\), where \(d\) is the true value and \(\hat{d}\) the estimator. The first information these two figures show is that as the number of points increases, our estimator gets more accurate: the variability in figure 2 decreases towards the right hand side, the confidence intervals get narrower and the relative error resembles loosely a multiple of the inverse of the square root of the number of points. We performed tests up to 5000 points, and the results show and improved convergence, but we consider unrealistic the demand for more than 2000 feature points in the scene. The same holds for the estimator of the slant angle \(\hat{\theta}\), figures 4 and 5. As a rule of thumb, we can say that with at least 1000 points in the image we can get reliable estimates of slant and distance (less than 5% of error).

6 CONCLUSIONS

In this work we propose a novel technique to study the distortion induced by perspective projection on a pla-
nar Poisson point process; we use this to recover the pose of a pinhole camera with two degrees of freedom, slant angle and distance from the ground along the optical axis. Our approach relies on the observation that the Voronoi tessellation generated by a Poisson point process under perspective is a faithful representation of the area transformation ratio, i.e. the Jacobian of the perspective, up to a scaling factor that we demand as input (the density of the Poisson process). We perform intensive simulations on synthetic data and do a careful error analysis, concluding that 1000 points on the scene are enough to get estimates of the parameters of interests with an error less than 5%. Our work borrows ideas from the shape from tex-
ture paradigm (the area gradient concept), but instead of assuming the presence of a whole patch of homogeneous or isotropic texture, we pursue a feature-based approach which considers only a discrete set of points with homogeneity properties; such different premises make a direct comparison with shape from texture algorithms non obvious. Our hypothesis make the proposed solution more suitable for camera pose estimation in general settings, specifically natural environments, where reference artifacts are missing and one must resort to stochastic modeling.

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REFERENCES


