

# Combinatorial Identification of Broad Association Regions with ChIP-seq Data

Jieun Jeong<sup>1</sup>, Mudit Gupta<sup>2</sup>, Andrey Poleshko<sup>2</sup> and Jonathan A. Epstein<sup>2</sup>

<sup>1</sup>Gastrointestinal Unit and Center for Computational and Integrative Biology, Massachusetts General Hospital, Harvard Medical School, Boston, MA 02114, U.S.A.

<sup>2</sup>Department of Cell and Developmental Biology, Institute for Regenerative Medicine, and the Penn Cardiovascular Institute, Perelman School of Medicine at the University of Pennsylvania, Philadelphia, PA 19104, U.S.A.

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**Abstract:** Motivation: Differentiation of cells into different cell types involves many types of chromatin modifications, and mapping these modifications is a key computational task as researchers uncover different aspects of that process. Modifications associated with heterochromatin formation pose new challenges in this context because we must define very broad regions that have only a moderately stronger signal than the rest of the chromatin. Lamin-associated domains (LADs) are a prime example of such regions. Results: We present Combinatorial Identification of Broad Association Regions (CIBAR), a new method to identify these types of broad regions. CIBAR is based on an efficient solution to a natural combinatorial problem, which adapts to widely variable yields of reads from ChIP-seq data and the associated controls and performs competitively with previous methods, including DamID, which has been used in many publications on LADs but cannot be applied in most *in vivo* situations.

## 1 INTRODUCTION

It is now widely accepted that there are many types of epigenetic chromatin modifications that profoundly impact cellular phenotypes and thus all types of life processes (Berstein et al., 2010). A typical task in processing genome-wide data regarding such modifications is to simplify the complex distribution of these modifications into simple collections of regions where they are present.

There is no universal gold standard concerning *present/absent* decisions, but in general, we wish to see reproducible results (thus removing or decreasing the impact of data noise) that are associated with biological outcomes of interest, such as activating or repressing gene expression or the presence of other modifications.

Because different modifications are associated with diverse molecular mechanisms, they require different computational tools. The best-investigated “regions of presence” are so-called (narrow) peaks of sequence-specific transcription factors, which can have lengths around 100 bps. Broad peaks, with lengths ranging from hundreds to thousands of bps, are typical regions for the presence of chromatin

factors, e.g., chromatin-modifying enzymes that form protein complexes with transcription factors but have some mobility that allows them to modify longer chromosomal intervals. A prime example of regions with peaks and broad peaks are so-called promoters and enhancers (Ernst et al., 2011), and even longer broad peaks may be associated with regions with Polycomb repression complex activity (Ernst et al., 2011; Mikkelsen et al., 2007). In this note, we focus on *broad regions*, which sometimes cover many millions of base pairs, that in many cases are associated with modifications that repress gene expression in wide domains because they are involved in the establishment of heterochromatin.

The tools for finding broad peaks are not adequate for finding broad regions because they share the basic assumption with tools for narrow peaks: namely, they define locations that have a much higher normalized concentration of ChIP reads than so-called background locations. However, heterochromatin covers most of the genome, and the modifications associated with heterochromatin have a very wide distribution; therefore, assumptions of that kind are not applicable.

One type of chromatin modification that is attracting

increasing attention is binding to lamins proteins present in the nuclear lamina, as well as the lamina of the nucleolus (Padeken and Heun, 2014).

The lamina lines the membranes of the nucleus and nucleolus and forms a mesh-like structure that stabilizes the position of a portion of the chromatin in the nuclear periphery. Chromosomal regions where contact with lamins is detected are called lamin-associated domains (LADs); binding to lamins is mediated through several proteins, and it is still under investigation. The following proportions have been reported in the mouse genome: (1) ca. 65% of chromatin is heterochromatin (Ernst *et al.*, 2011), (2) ca. 40% of chromatin is in LADs (Guelen *et al.*, 2008), and (3) more than 90% of LADs are in heterochromatin (this estimate is based on the depletion of LADs with expressed genes [Table 1]). It is important to note that eukaryotic genomes vary widely in length and thus in gene density; hence, those proportions should be different in different species. The extent of heterochromatin changes during cell differentiation, and these changes are part of the processes that determine cell fate. In turn, a number of chromatin factors have been implicated in the formation of LADs.

Table 1: Joint length, the number of gene transcription start sites (TSS) and the number of TSSs of expressed genes (E-TSS, genes with an e9.5 expression level above the median) for LADs identified by DamID and by CIBAR. We define the consensus of the four e9.5 samples (three e12.5 samples) as windows that are selected to be in LADs for all of them, with one possible exception, and we call the resulting set "3 of 4" ("2 of 3" for e12.5 samples).

region type	Mbps %	TSSs %	E-TSSs %
whole genome	2500 <b>100</b>	18988 <b>100</b>	9538 <b>100</b>
DamID	Astro	1112 <b>44</b>	3601 <b>18</b>
	ESC	1067 <b>42</b>	2771 <b>14</b>
	MEF	1179 <b>47</b>	3106 <b>16</b>
	NPC	1070 <b>42</b>	3171 <b>16</b>
	A	864 <b>34</b>	2405 <b>12</b>
e9.5	B	847 <b>33</b>	2655 <b>13</b>
	C	860 <b>34</b>	3122 <b>16</b>
	D	866 <b>34</b>	3348 <b>17</b>
	K	841 <b>33</b>	2415 <b>12</b>
e12.5	L	870 <b>34</b>	2330 <b>12</b>
	M	849 <b>33</b>	2555 <b>13</b>
	e9.5 3 of 4	721 <b>28</b>	1957 <b>10</b>
e12.5 2 of 3	832 <b>33</b>	2122 <b>11</b>	416 <b>4</b>

Computation of LADs includes the following three basic stages, and for each stage we can choose a method: (1) we obtain the data from the genome (signal and control), (2) we convert the data to a "normalized signal", and (3) we use the normalized

signal to determine regions where the normalized signal values are mostly higher than in the remainder of the genome.

There are two methods of collecting the binding data for lamins and proteins that interact with lamins for the purpose of genome-wide mapping of LADs. The first is DamID (Guelen *et al.*, 2008). In DamID, the protein binding data are obtained by adding a DNA methylating domain to the investigated protein. Then, the loci with methylated DNA are identified by hybridization to microarrays (with ca. 2.5 million probes in the case of the mouse genome). These data are paired with data in which the protein is augmented with a "neutral" domain. The benefits are that no specific protein antibodies are required and that the control data match the signal data very well. However, such genomic manipulation is impractical *in vivo*; in particular, DNA methylation itself can affect the phenotype. The second method is to collect ChIP-seq using an antibody and input, which is the associated control formed from a portion of the starting material, i.e., the DNA from the respective cell sample. The signal data that are collected are quite uniformly distributed while methods of the broad peak-type perform well only when we have a high fold change between regions where we declare the *present* compared with the remainder of the chromatin.

The standard method for computing the "normalized signal" from the data is to use ratios, such as the "the number of ChIP reads" over "the number of input reads." However, we present our alternative method in Section 2.

Several methods have been previously used to identify broad regions from binding data.

The first method is a 2-state Hidden Markov Model (HMM). The use of HMM for defining genomic regions is well established (e.g., Ernst *et al.*, 2011) and seems indispensable when we integrate the signals of multiple types. HMM has been used by Xu *et al.* (2010) to compute the broad peaks from ChIP-seq/input data, and is also a part of the DamID method. However, when we have only one type of signal, combinatorial algorithms are faster and offer a better understanding of how the regions are defined, and consequently, of how to set the parameters to achieve the desired outcomes.

Lund *et al.* (2014) have tested several existing programs for computing "long peaks" and have found that all of them are inadequate for LADs. Another combinatorial approach used to identify LADs is the "sliding windows" method (Shah *et al.*, 2013), in which a long window consists of  $k$  short windows (called steps). We accept a long window if it contains

some threshold number of short windows with the positive score, and we declare LADs as the union of the accepted long windows. An undesirable aspect of this method is that it ignores the signal values that can have considerable amplitudes and are not captured by comparison with a threshold. Moreover, the user must arbitrarily choose many parameters, including the threshold on the values, the length of the long windows and the minimum number of “positive” short windows.

Those undesirable aspects are not present in a method that uses a combinatorial problem defined in terms of the sums of signal values (that can be positive and negative) in the selected intervals. The EDD method by Lund *et al.* (2014) incorporates the following combinatorial problem: (1) we are given a number of arrays with real values and (2) we find a contiguous fragment with the maximum sum of values. Then, we select that fragment and remove it from its array, which may increase the number of remaining arrays. We repeat this selection a desired number of times, and finally we reject the selected regional candidates that do not pass a permutation test for significance.

Our CIBAR method has two major differences from EDD. First, we use a different formula for computing the signal values that more accurately follows the dependency between the distribution of the ChIP and input reads. Second, instead of selecting the fragments one at a time, we apply a framework of global optimization such that for a given array and  $k$ , we select  $k$  array fragments in such a way that the sum of all of the values in the selected fragments is maximized. We also show a very efficient solution to that problem.

## 2 METHODS - DESCRIPTION OF CIBAR

### 2.1 Windows

We partitioned the genome into 1000 bp windows (also known as bins) and each window has a position, e.g., chr2:3001 for bps 3,001,001 to 3,002,000. Each window  $w$  has  $a_w$  input reads and  $b_w$  ChIP reads. Because the windows are merged into much longer regions, the precise selection of the window length has moderate importance. However, in our experience, the average value of the  $a_w$ , number of input reads in a window should be in the range of 50.

We collected reads from several ChIP and input sets, and we process only windows with at least one

read mapped in one of the sets. The other windows are ignored, and we call them “null windows.” As a result, we do consider windows with  $a_w = b_w = 0$  if they have reads in some other data sets.

### 2.2 Scores of Windows – The Motivation

The “significance” of an interval of windows is related to the p-value that we see for a particular number of signal reads in that interval, but we normalize that number on the basis of the number of input (control) reads.

The question that we address in a novel way is how to normalize the signal. ChIP reads are not uniformly collected through the genome because of the differences in the accessibility of different parts of chromatin to the processes that generate the reads (Kharchenko *et al.* 2008). In particular, heterochromatin tends to be less accessible, and the frequency of input reads in heterochromatin tends to be lower. This motivates normalization approaches, such as the ratio between the normalized ChIP and input read counts.

The ratio approach is valid if the relationship between the ChIP and input counts is linear. However, this is not the case.

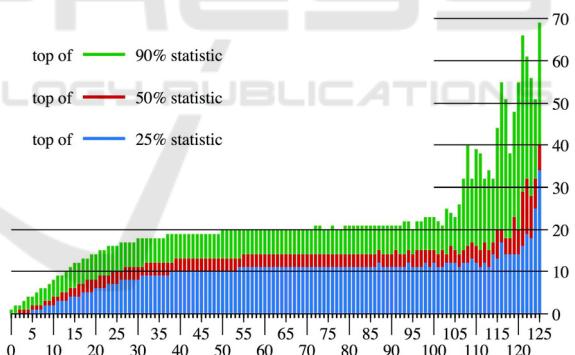


Figure 1: The X-axis is the  $a_w$ , which is the number of input reads in a 1 kbp window, and each bar represents the set of all windows with  $a_w$  indicated by the X-coordinate. The heights of the colors for the Y-coordinate are the values of  $b_w$ , which are the numbers of ChIP reads that are order statistics for this window set. The top of blue is 25% (the highest value in the bottom quadrant), the top of red is 50% (the median), and the top of green is 90% (the lowest value in the top decile). Forty-four is the median level for the input reads, less than 0.1% of the windows have more than 100 input reads, and because the sets for  $a_w > 100$  are very small, their statistics change quite irregularly.

Figure 1 illustrates the distribution for a typical input-ChIP pair. In that example, the input-ChIP relationship is well approximated as linear between 0

and 20 input reads, and as constant between 30 and 100 input reads; i.e., the ChIP read counts are almost independent of the input counts. For rare high counts of input reads, the relationship is once again roughly linear.

Normalization is important because 20-25% of windows have low input counts and proportionally smaller ChIP counts. Moreover, these windows are primarily in the heterochromatin where we expect LADs to be. However, for a wide range of input values, linear normalization would be a significant distortion. Our solution is to normalize the signals separately for each possible number of input reads.

To our knowledge, previous work has made no account of this problem. The use of ChIP-input ratio did not produce very distorted results because the effect is to exaggerate the link between LADs and heterochromatin, although in actuality, this link is very strong. However, in studies of embryonic differentiation, etc., we are very much interested in “marginal” LADs, where the dynamic extent of LADs has regulatory impact on gene expression without total heterochromatin silencing characteristics. Thus, misclassification of LADs, over 1-5% of the genome can have a high impact on the lists of genes whose regulation can be attributed to LADs. Moreover, the positive bias for heterochromatin can be detrimental when the method is applied to regions associated with PRC activity or gene elongation.

Other than separating the windows according to their input counts, our method for computing the signal is quite similar to the one by Lund *et al.* (2014).

### 2.3 Window Scores – The Formula

Given that a window  $w$  has  $a_w$  input reads and  $b_w$  ChIP reads, we define the following statistics:

- (1)  $n_i$  is the number of all of the windows with  $a_w = i$ ,
- (2)  $s_{ik}$  is the number of all of the windows with  $a_w = i$  and  $b_w = k$ , and
- (3)  $A_{ik}$  is the average rank of windows with  $a_w = i$  and  $b_w = k$  among all of the windows with  $a_w = i$ , i.e.,  $A_{ik} = s_{i0} + \dots + s_{i(k-1)} + \frac{1}{2}s_{ik}$ .

The score( $i, k$ ) =  $\kappa(2A_{ik}/n_i)$ , where  $\kappa(x) = \log x$  if  $x \leq 1$  and  $\kappa(x) = -\log(2-x)$  if  $x \geq 1$ .

Note that the score( $i, k$ ) is negative/zero/positive if  $A_{ik}$  is smaller/equal/greater than  $\frac{1}{2}n_i$ , respectively.

The idea of the scoring is that we prefer to select windows with positive scores and to avoid selecting windows with negative scores. However, with the

described method for assigning scores, about half of the windows have positive scores, and about half have negative scores, even though the target regions occupy less than half of the genome. Therefore, we use

$$\text{score}(w) = \text{score}(a_w, b_w) - \alpha,$$

where  $\alpha$  is the score adjustment, which is the same for all of the non-null windows. The larger the score adjustment we use, the fewer the windows belonging to the selected regions. The score adjustment is one of the two parameters supplied by the user.

This scoring formula is subjected to a high statistical fluctuation when  $n_i$  is small. If  $n_i$  is smaller than 1000, we enlarge the counts by forming a set of windows of size 1000 or more, which have between  $i-d$  and  $i+d$  input reads for the smallest possible  $d$ .

### 2.4 From Signal Values to Regions/ Domains

LAD signals computed with our formula exhibit large random-like variability, in part because the collection of reads, both input and ChIP, is essentially a random process of extracting the reads from a library. In the case of narrow and broad peaks, we observe concentrations of ChIP reads that are many times larger than background, although this is not the case with the much less-concentrated signal that we see in very broad regions. Instead, we can see a rather small advantage of the positive signal that persists over a long interval.

We identify our broad regions by solving the 1-dimensional fragment selection (*IDFS*): given an array with real values (scores of windows in chromosomes) and a target number  $k$ , find  $k$  disjointed fragments that are contiguous sub-arrays and have the maximum sum of scores. The *IDFS* problem describes the optimization goals more directly than the iterative formulation by Lund *et al.* (2014). Importantly, exact solution of *IDFS* can be found efficiently in time proportional to  $n \log n$ , where  $n$  is the number of windows. This is slightly slower than the speed of the algorithm in Lund *et al.* (2014), but the probabilistic validation of the computed regions takes approximately 10 times longer and our entire computation takes about one minute on a 1.7 Ghz Macintosh computer. Our algorithm for *IDFS* uses a greedy method and is easy to implement.

We give the details of the *IDFS* algorithm and the proof of its correctness in the Appendix. To our knowledge, this is a new algorithm that may have more applications in genomic region studies.

### 3 RESULTS

We applied our method to laminB1 ChIP data collected for the following two biological conditions (ChIP was performed as described in Shah *et al.* (2013)): (1) the mouse embryonic heart at e9.5 and (2) the mouse embryonic heart at e12.5 (days after conception), with four and three replicates, respectively. These samples are subsequently referred to as A-D and K-M, respectively. We found that using an input with less than 50 million reads leads to erratic results, so we used the same input for all of the ChIP samples; it was collected from the e9.5 sample and had 111 million reads. We set the target number of LADs at 1500 and experimented with different adjustment parameters. Adjustments that led to LADs having much more than 800 Mbps in joint length had a considerable proportion of regions with p-values above 0.001, so we decided not to further increase the size (see Appendix for the details of p-value test). Then, we checked whether the LADs have the properties previously reported in the literature. The first property is that LADs are severely depleted of genes and that the genes present in LADs are mostly silent (very low expression level). We tested this on 18,988 genes that have annotations in refGene.txt for the mm9 mouse genome build (available from UCSC Genome Browser) and for which we also had microarray expression data.

Table 1 shows that our LADs contain similar numbers of genes and expressed genes as LADs computed with the DamID method (see Peric-Hupkes *et al.*, 2010). DamID LADs cover 42-47% of the genome, 14-18% of genes and 4-6% of expressed genes. Our LADs (identified by CIBAR) cover 33-34% of the genome, representing proportionally fewer genes, and are also depleted expressed genes, though not as much.

Taking a consensus of regions computed for the same condition results in a similar depletion in expressed genes as in DamID; consensus is an effective method to eliminate false positives. We defined the consensus as the base pairs that occur in all or all but one of the LAD sets that were computed for that condition; this is shown in Tables 1-3.

We also measured the level of consistency for the LADs computed for the four cell types studied in Peric-Hupkes *et al.* (2010) and our LADs from the embryonic heart (see Table 2).

The degree of consistency between our embryonic LADs and the LADs computed in Peric-Hupkes *et al.* (2010) is similar to the consistency between those cell types (note that among those four types, neural

pluripotent cells (NPC) and astrocytes, a differentiated neural type, are the closest).

Table 2: Similarity of different LAD sets. For e9.5 and e12.5, we use consensus regions (called “3 of 4” and “2 of 3”). The percentage of the smaller set that belongs to both sets being compared is in red.

Set	Mbps	Common Mbps % of the smaller set					
		e9.5	e12.5	Astro	ESC	MEF	NPC
e9.5	721		565 78	584 81	581 81	600 83	594 82
e12.5	832	565 81		577 69	605 73	635 76	579 70
Astro	1112	582 81	577 69		772 72	881 78	931 87
ESC	1067	581 81	605 73	772 72		836 78	786 87
MEF	1179	600 83	635 76	881 78	836 78		880 82
NPC	1070	594 82	579 70	931 87	786 74	880 82	

In Table 3 we compare EDD with CIBAR.

Both EDD and CIBAR produce LADs with properties reported in Peric-Hupkes *et al.* (2010), namely that they cover approximately 40% of the genome (32-33% in the case of CIBAR and 34-36% in the case of EDD) and that LADs are depleted in genes, even more so in expressed genes, as we already discussed. The consistency with DamID LADs is also similar. CIBAR LADs have two advantages. First, they are computed with 1 kbp windows, while EDD automatically selects window sizes between 35 and 90 kbps, which means that the LAD boundaries are much less precise. Second, EDD produced 200-275 LADs, and CIBAR produced 880-1280 LADs, which is closer to what was previously reported.

Table 3: Comparison of results of CIBAR and EDD. The column labels correspond to row labels in Table 1.

	e9.5				e12.5			
					3 of 4	K	L	M
	A	B	C	D				2 of 3
EDD								
Mbps	898	941	898	849	871	875	911	907
TSS	2241	2378	2413	2169	2117	2162	2170	2291
E-TSS	425	481	547	316	389	459	481	475
CIBAR								
Mbps	829	814	824	820	708	812	828	805
TSS	2116	2193	2410	2465	1788	1927	1783	1980
E-TSS	298	442	618	371	213	412	372	387
								314

To a degree, this is a consequence of the parameters we chose for CIBAR, namely the selection of the

window size, target size (regulated by the adjustment) and the target number of regions.

## 4 CONCLUSIONS

We presented a new method of computing broad regions associated with chromatin modifications that is applicable even when ChIP data do not exhibit large fold changes between the affected regions and the rest of the genome. Although our method was conceived and implemented before the publication of EDD (Lund *et al.*, 2014), it shares the following three aspects: (a) using scores of windows rather than a selected cut-off between “good” and “bad” windows, (b) basing the score number on the ranks of the windows and (c) applying a natural combinatorial problem to group windows into regions.

Our scoring method models the distributions of ChIP and control reads more accurately; thus, we avoid the positive bias for selecting windows in the least accessible parts of the genome. It remains an open question whether this is a good model, and we expect further progress in this direction.

The combinatorial problem that we have applied, *IDFS*, is much more natural than the iterative selection of fragments with the highest sum of scores, which can excessively merge “positive” regions with the “negative” regions that separate them. Lund *et al.* (2014) introduced *gap penalty* (decreasing all negative scores by a constant) to reduce that tendency, but we suspect that this is one of the reasons why EDD works with such low granularity. Although *IDFS* is a global optimization problem, we have found a solution that is very efficient and easy to implement.

Our method uses only two parameters,  $k$  and  $\alpha$ , but the proper selection of parameters remains an open problem. In Section 2, we set  $k$  to obtain the number of LADs, which is close to the number reported in papers applying the DamID method (see Peric-Hupkes *et al.*, 2010). Parameter  $\alpha$  can be selected in different ways. As it increases, the proportion of windows with positive  $score(w)$  decreases, as does the sum of lengths of identified LADs. However, when we decrease  $\alpha$  too much, the p-values of the computed LADs tend to increase, and we cannot suggest a statistic that allows to optimize  $\alpha$ . In fact, we tested our program and EDD on six genes confirmed to be in LADs using ChIP-qPCR (data not included). We found that increasing  $\alpha$  may paradoxically exclude some of them, whereas choosing a consistent  $\alpha$  of 0.12 led to consistent inclusion of 5 of the genes in LADs computed for all

four e9.5 samples. LADs computed by EDD (which automatically adjusts parameters to optimize the p-values) consistently included exactly 3 of these genes. This small evidence suggests that at present there is no better way to select the parameters than using whatever knowledge we have, most preferably some genomic positions confirmed to be in LADs or outside LADs, and picking the parameters to be consistent with that knowledge. The situation with identifying short peaks of transcription factors is similar because the existing programs can produce “false positives,” i.e., statistically significant peaks that are too weak to have a biological impact. Therefore, these programs provide options to select the parameters, such as maximum p-value/FDR, minimum fold change.

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## APPENDIX – THE GREEDY ALGORITHM FOR THE *IDFS* PROBLEM

### 1.1 Problem Definition

An instance of *IDFS* consists of an array of real values  $A[0] \dots A[n-1]$  and a positive integer  $k$ . A fragment  $A[s, t]$  is specified by two distinct integers; if  $s < t$  then it consists of  $A[s] \dots A[t-1]$ , otherwise it is a “wrap around” fragment that consists of  $A[s] \dots A[n-1]$  and  $A[0] \dots A[t-1]$ . The *value* of a fragment is the sum of its entries. The goal is to find  $j \leq k$  disjoint fragments with the maximum possible sum of values.

Allowing the wrap-around fragments simplifies the problem such that every entry has two neighbors, left and right, that can join it in the same fragment. Moreover, it is equivalent to find  $j$  disjoint and non-adjacent fragments with the maximum sum of values and to find  $j$  fragments with the minimum sum of values because the complement of  $j$  disjoint and non-adjacent fragments also consists of  $j$  disjoint and non-adjacent fragments.

In our application, we have a number of arrays, one for each chromosome. However, even though *IDFS* has only one array and allows wrap-around fragments, it is in fact equivalent. We can add an entry with a very low negative value to the end of each of those arrays and combine all of them into one. Then, the “improper” fragments are formally allowed, but they cannot be present in an optimal solution because they have negative values.

### 1.2 Instances with Restricting Partitions and the Greedy Algorithm

We generalize problem instance by adding  $P$ , a *restricting partition* of array  $A$ , into disjoint fragments, and then we allow only fragments that are unions of fragments from  $P$  in the solution. The fragments from  $P$  can be numbered  $p_0 \dots p_{m-1}$ , and we use  $Ap[i]$  to denote the value of  $p_i$ .

Our algorithm starts with the restricting partition of  $A$  into 1-entry fragments and then applies the rules described in Figure 2 until it terminates. The algorithm is greedy in the sense that the rules are very simple, and it either terminates or applies a rule that reduces the size of the problem (the number of fragments in the restricting partition).

Use the first applicable rule from the following list:

- (1) **Terminating Rule.** If  $P$  contains  $k$  or fewer fragments with positive values, return the set of these fragments.
- (2) **Equal Sign Rule.** If  $p_i$  and  $p_{i+1}$  both have negative values or both have non-negative values, merge them together.
- (3) **Minimum Value Rule.** Select  $i$  with the minimum  $|Ap[i]|$ , merge together  $p_{i-1}, p_i$  and  $p_{i+1}$ .

Figure 2. The rules in the algorithm for the *IDFS* problem.

#### 1.2.1 Correctness of the Terminating Rule

The sum of the values from the selected fragments cannot exceed the sum of the values from all of the positive fragments in  $P$ ; thus, the value of the returned solution is optimal, and because there are at most  $k$  such fragments, the solution is valid.

#### 1.2.2 Correctness of the Equal Sign Rule

Suppose that the solution  $S$  is consistent with the former restricting partition but not with the new one. Then, a fragment of  $S$  ( $F$ , for example) contains  $p_i$  but excludes  $p_{i+1}$  or vice versa. If the value of  $p_i$  is negative, we can modify  $F$  by excluding  $p_i$ , thus increasing the value of  $F$  and  $S$ ; this works because  $p_i$  is at the end of  $F$ . Similarly, if the value of  $p_i$  is non-negative, we can modify  $F$  by extending it with  $p_{i+1}$ , which before was fully excluded from the solution, and the value of the solution will increase or stay the same because the value of  $p_{i+1}$  is also non-negative. Thus we defined a solution  $S'$  that has the same value as  $S$  and is consistent with the new partition.

#### 1.2.3 Correctness of the Minimum Value Rule

Because the Termination Rule does not apply, the number of the positive restriction fragments is  $l > k$ . Because the Equal Sign Rule does not apply, the fragments of  $P$  with negative and non-negative values alternate, so the number of  $P$ -fragments is equal to  $2l > 2k$ .

Suppose that solution  $S$  is consistent with the former restricting partition,  $P$ , but not with the new one  $P'$ . This means that the fragments of  $S$  contain some but not all of  $p_{i-1}, p_i$  and  $p_{i+1}$ . We can assume that  $S$  cannot be improved by extending or shrinking one of its fragments by one  $P$ -fragment. Hence, in every fragment of  $S$ , the first and the last  $P$ -fragment have positive values, and in every fragment of the complement of  $S$ , the first and last  $P$ -fragment has negative values. Thus, each fragment of  $S$  and of the

complement of  $S$  consists of an odd number of restriction fragments.

**Case 1:** Fragments of  $S$  contain only  $p_{i-1}$ . We can extend  $S$  by incorporating  $p_i$  and  $p_{i+1}$ , and the value of  $S$  increases by  $|Ap[i+1]| - |Ap[i]| \geq 0$ .

**Case 2:** Fragments of  $S$  contain only  $p_i$ , so the restriction fragment  $p_i$  is also a fragment of  $S$ . We remove  $p_i$  from  $S$ , which decreases the value of  $S$  by  $x$ , which is the value of  $p_i$ . We have to modify  $S$  to increase its value by  $x$  or more. Because we have decreased the number of fragments in  $S$ , the modification is allowed to increase the number of fragments by 1. Because there are more than  $2k$  restriction fragments, one of the fragments of  $S$  or the complement of  $S$  ( $F$ , for example) contains at least three restriction fragments. If  $F$  is a fragment of  $S$ , it contains a restriction fragment  $p_j$  with a negative value  $-y$ . We remove  $p_j$  from  $F$ , which increases the number of fragments by one and increases the value of the solution by  $y \geq x$ . If  $F$  is a fragment of the complement of  $S$ , it contains a restriction fragment  $p_j$  with a positive value  $y$ , so we add  $p_j$  to  $S$ . Again, the number of fragments in  $S$  increases by 1 and the value by  $y \geq x$ . Importantly, as  $p_j$  neither starts nor ends  $F$ , it cannot be  $p_{i-1}$  or  $p_{i+1}$ .

Other cases are symmetric, so every solution that satisfies the restriction before we applied the merging can be modified to a solution that satisfies the new restriction with a value that is the same or larger.

## Implementation

We first compute a matrix with read counts for every data sample (input and ChIP) and every window, and we then use that matrix (tab separated file) as the input for our C-program. Null windows, those without any reads, are ignored as “unmappable”.

In the program, we use an array of doubly linked lists for fragments in the restriction partition, with one list for each chromosome. The fragments are also represented in a binary heap (priority queue), which allows rapid application of the Minimum Value Rule.

To assess the p-value of computed regions, we used the null model in which the scores of windows are randomly permuted. We found out that for regions with more than 200 windows, the p-value can be well approximated using the cumulative Gaussian distribution. For fewer windows, we use one million runs of 200 random window selections, which allowed us to tabulate empirical p-values for all of the region sizes up to 200 windows. As a result, this Monte Carlo test adds less than a minute to the computation time and uses very little additional memory.