A NEW APPROACH FOR DENOISING IMAGES BASED ON WEIGHTS OPTIMIZATION

Qiyu Jin\textsuperscript{1,2}, Ion Grama\textsuperscript{1,2} and Quansheng Liu\textsuperscript{1,2}

\textsuperscript{1}UMR 6205, Laboratoire de Mathématiques de Bretagne Atlantique, Université de Bretagne-Sud, Campus de Tohanic, BP 573, 56017 Vannes, France
\textsuperscript{2}Université Européenne de Bretagne, Rennes, France

Keywords: Non-local Means, Image Denoising, Optimization of Weights, Oracle, Statistical Estimation.

Abstract: We propose a new algorithm to restore an image contaminated by the Gaussian white noise. Our approach is based on the weighted average of the observations in a neighborhood as in the case of the Non-Local Means Filter. But in contrast to the Non-Local Means Filter, we choose the weights by minimizing a tight upper bound of the Mean Square Error. Our theoretical results show that some “oracle” weights defined by a triangular kernel are optimal. To construct a computable filter the “oracle” weights are replaced by some estimates. The implementation of the proposed algorithm is straightforward. The simulations show that our approach is very competitive.

1 INTRODUCTION

We deal with the additive Gaussian noise model:

\[ Y(x) = f(x) + \epsilon(x), \quad x \in \mathbf{I}, \]

where \( \mathbf{I} \) is the uniform \( N \times N \) grid of pixels on the unit square, \( Y = (Y(x))_{x\in\mathbf{I}} \) is the observed image brightness, \( f : [0,1]^2 \rightarrow \mathbb{R}_+ \) is an unknown target regression function and \( \epsilon = (\epsilon(x))_{x\in\mathbf{I}} \) are independent and identically distributed (i.i.d.) Gaussian random variables with mean 0 and standard deviation \( \sigma > 0 \).

Important denoising techniques for the model (1) have been developed in recent years. A very significant step in these developments was the introduction of the Non-Local Means Filter by (Buades et al., 2005). For closely related works, see for example (Polzehl and Spokoiny, 2006; Kervrann and Boulanger, 2008; Buades et al., 2009; Katkovnik et al., 2010; Lou et al., 2010).

The basic idea of the filters by weighted means is to estimate the unknown image \( f(x_0) \) by a weighted average of observations \( Y(x) \) of the form

\[ \tilde{w}(x_0) = \sum_{x \in U_{y, h}} w(x)Y(x), \]

where for each \( x_0 \) and \( h > 0 \), \( U_{y, h} \) denotes a square window with center \( x_0 \) and width \( 2h \), \( w(x) \) are some non-negative weights satisfying \( \sum_{x \in U_{y, h}} w(x) = 1 \).

The choice of the weights \( w(x) \) are usually based on two criteria: a spatial criterion so that \( w(x) \) is a decreasing function of the distance between \( x \) and \( x_0 \), and a similarity criterion so that \( w(x) \) is also a decreasing function of the brightness difference \( |Y(x) - Y(x_0)| \) (see e.g. (Yaroslavsky, 1985; Tomasi and Manduchi, 1998)).

In this paper we address the problem of choosing the weights \( w \) in (2) in some optimal way. Generally, the weights \( w(x) \) are calculated according to the similarity between data patches \( Y_{x, \eta} = (Y(y) : y \in U_{y, \eta}) \) (defined as a vector whose components are ordered lexicographically) and \( Y_{x_0, \eta} = (Y(y) : y \in U_{x_0, \eta}) \), instead of the similarity between just the pixels \( x \) and \( x_0 \). Here \( \eta > 0 \) is the size parameter of data patches.
bound of the Mean Square Error

\[ R \left( \hat{f}_w(x_0) \right) = \mathbb{E} \left( \hat{f}_w(x_0) - f(x_0) \right)^2 \]

in terms of the bias and variance and to minimize this upper bound in \( w \) under the constraints \( w \geq 0 \) and \( \sum_{x \in U_{x_0,h}} w(x) = 1 \). We first obtain an explicit formula for the optimal weights \( w^* \) in terms of the unknown function \( f \). In order to get a computable filter, we estimate \( w^* \) by some adaptive weights \( \hat{w} \) based on data patches from the observed image \( Y \). We thus obtain a new filter, which we call Optimal Weights Filter. Numerical results show that the new filter outperforms the typical Non-Local Means Filter, thus giving a practical justification that the optimal choice of the kernel improves the denoising quality.

We would like to point out that related optimization problems for non parametric signal and density recovering have been proposed earlier in (Sacks and Ylvisaker, 1978; Nazin et al., 2008). In these papers the weights are optimized over a given class of regular functions and thus depend only on some parameters of the class. The novelty of our work is to deal with optimal weights depending on the image \( f \) at hand. Results of this type are related to the "oracle" concept developed in (Donoho and Johnstone, 1994).

## 2 OPTIMAL WEIGHTS FILTER

In this section, we present our new filter called Optimal Weights Filter, and explain the idea behind its construction.

We begin with some mathematical notations that will be used throughout the paper. For a vector \( x = (x_1, \ldots, x_d) \in \mathbb{R}^d \), we denote by \( \|x\|_2 = (\sum_{j=1}^{d} x_j^2)^{1/2} \) its Euclidean norm and by \( \|x\|_{\infty} = \max_{1 \leq j \leq d} |x_j| \) its supremum norm. The cardinality of a set \( A \) is denoted by \( \text{card} A \). For a positive integer \( N \) the uniform \( N \times N \) grid on the unit square is defined by

\[ \mathbf{I} = \left\{ \frac{1}{N}, \frac{2}{N}, \ldots, \frac{N-1}{N}, 1 \right\}^2. \]

Each element \( x \) of the grid \( \mathbf{I} \) will be called pixel. The number of pixels is \( n = N^2 \). For any pixel \( x_0 \in \mathbf{I} \) and a given \( h > 0 \), the square window of pixels \( U_{x_0,h} = \{ x \in \mathbf{I} : ||x - x_0||_{\infty} \leq h \} \) will be called search window at \( x_0 \). We naturally take \( h \) as a multiple of \( \frac{1}{N} \) ( \( h = \frac{k}{N} \) for some \( k \in \{1, 2, \ldots, N\} \)). The size of the square search window \( U_{x_0,h} \) is the positive integer number

\[ M = (2Nh + 1)^2 = \text{card} U_{x_0,h}. \]  

For any pixel \( x \in U_{x_0,h} \) and a given \( \eta > 0 \) a second square window of pixels \( U_{x,\eta} \) will be called patch at \( x \). Like \( h \), the parameter \( \eta \) is also taken as a multiple of \( \frac{1}{N} \). The size of the patch \( U_{x,\eta} \) is the positive integer

\[ m = (2N\eta + 1)^2 = \text{card} U_{x,\eta}. \]  

The vector \( Y_{x_0,\eta} = (Y(y))_{y \in U_{x_0,\eta}} \) formed by the values of the observed noisy image in the patch \( U_{x_0,\eta} \) in the lexicographical order will be called data patch (or similarity patch) at \( x \in U_{x_0,h} \). Finally, the positive part of a real number \( a \) is denoted by \( a^+ : a^+ = a \) if \( a \geq 0 \) and \( a^+ = 0 \) if \( a < 0 \).

Let \( h > 0 \) be fixed. For any pixel \( x_0 \in \mathbf{I} \) consider a family of weighted estimates \( \hat{f}_{h,w}(x_0) \) of the form

\[ \hat{f}_{h,w}(x_0) = \sum_{x \in U_{x_0,h}} w(x)Y(x), \]

where the unknown weights satisfy

\[ w(x) \geq 0 \quad \text{and} \quad \sum_{x \in U_{x_0,h}} w(x) = 1. \]

The usual bias plus variance decomposition of the Mean Square Error gives

\[ \mathbb{E} \left( \hat{f}_{h,w}(x_0) - f(x_0) \right)^2 = \text{Bias}^2 + \text{Var}, \]

with

\[ \text{Bias}^2 = \left( \sum_{x \in U_{x_0,h}} w(x) \left( f(x) - f(x_0) \right) \right)^2 \]

and

\[ \text{Var} = \sigma^2 \sum_{x \in U_{x_0,h}} w(x)^2. \]

The decomposition (7) is commonly used to construct asymptotically minimax estimators over some given classes of functions in the nonparametric function estimation. With our approach the bias term \( \text{Bias}^2 \) will be bounded in terms of the unknown function \( f \) itself. As a result we obtain some "oracle" weights \( w \) adapted to the unknown function \( f \) at hand, which will be estimated further using data patches of the image \( Y \).

First, we address the problem of determining the "oracle" weights. With this aim denote

\[ \rho_{f,x_0}(x) = \|f(x) - f(x_0)\|. \]

Note that the value \( \rho_{f,x_0}(x) \) characterizes the variation of the image brightness of the pixel \( x \) with respect to the pixel \( x_0 \). From the decomposition (7), we easily obtain a tight upper bound in terms of \( \rho_{f,x_0} \):

\[ \mathbb{E} \left( \hat{f}_h(x_0) - f(x_0) \right)^2 \leq g_{f,x_0}(w), \]

113
where
\[ g_{\rho_{\varphi}, \omega}(w) = \left( \sum_{x \in U_{\varphi}, h} w(x) \rho_{\varphi, \omega}(x) \right)^2 + \sigma^2 \sum_{x \in U_{\varphi}, h} w(x)^2. \]
\[ (10) \]

From the following theorem we can obtain the form of the weights \( w \) which minimize the function \( g_{\rho_{\varphi}, \omega}(w) \) under the constraints (6) in terms of \( \rho_{\varphi, \omega}(x) \). Introduce the strictly increasing function
\[ M_{\rho_{\varphi}, \omega}(t) = \sum_{x \in U_{\varphi}, h} \rho_{\varphi, \omega}(x) (t - \rho_{\varphi, \omega}(x))^+, \quad t \geq 0. \]

Let \( K_{ir} \) be the usual triangular kernel:
\[ K_{ir}(t) = (1 - |t|^+) \quad t \in \mathbb{R}^1. \]
\[ (11) \]

**Theorem 1.** Assume that \( \rho_{\varphi, \omega}(x), x \in U_{\varphi}, h \), is a non-negative function. Then the unique weights which minimize \( g_{\rho_{\varphi}, \omega}(w) \) subject to (6) are given by
\[ w_{\rho_{\varphi}, \omega}(x) = \frac{K_{ir}(\rho_{\varphi, \omega}(x)/a)}{\sum_{y \in U_{\varphi}, h} K_{ir}(\rho_{\varphi, \omega}(y)/a)}, \quad x \in U_{\varphi}, h. \]
\[ (12) \]
where the bandwidth \( a > 0 \) is the unique solution in \((0, \infty)\) of the equation
\[ M_{\rho_{\varphi}, \omega}(a) = \sigma^2. \]
\[ (13) \]

**Remark 1.** The value of \( a > 0 \) can be calculated as follows. We sort the set \( \{ \rho_{\varphi, \omega}(x) \mid x \in U_{\varphi}, h \} \) in the ascending order \( 0 = \rho_1 \leq \rho_2 \leq \cdots \leq \rho_M < \rho_{M+1} = +\infty \), where \( M = \text{Card} U_{\varphi}, h \). Let
\[ a_k = \frac{\sigma^2 + \sum_{i=1}^{k} \rho_i^2}{\sum_{i=1}^{k} \rho_i}, \quad 1 \leq k \leq M, \]
\[ (14) \]
and
\[ k^* = \max \{1 \leq k \leq M \mid a_k \geq \rho_k\} = \min \{1 \leq k \leq M \mid a_k \geq \rho_k\} - 1, \]
\[ (15) \]
with the convention that \( a_k = \infty \) if \( \rho_k = 0 \) and that \( \min \emptyset = M + 1 \). Then the solution \( a > 0 \) of (13) can be expressed as \( a = a_{k^*} \); moreover, \( k^* \) is the unique integer \( k \in \{1, \cdots, M\} \) such that \( a_k \geq \rho_k \) and \( a_{k+1} < \rho_{k+1} \) if \( k < M \).

Let \( x_0 \in I \). Using the optimal weights given by Theorem 1, we first introduce the non-computable approximation of the true image, called "oracle":
\[ f^o_x(x_0) = \frac{\sum_{y \in U_{\varphi}, h} K_{ir}(\rho_{\varphi, \omega}(y)/a)Y(x)}{\sum_{y \in U_{\varphi}, h} K_{ir}(\rho_{\varphi, \omega}(y)/a)}, \]
\[ (16) \]
where the bandwidth \( a \) is the solution of the equation \( M_{\rho_{\varphi}, \omega}(a) = \sigma^2 \). A computable filter can be obtained by estimating the unknown function \( \rho_{\varphi, \omega}(x) \) and the bandwidth \( a \) from data paths.

Let \( h > 0 \) and \( \eta > 0 \) be fixed numbers. For any \( x_0 \in I \) and any \( x \in U_{\varphi}, h \) consider the distance between the data patches \( Y_{x, \eta} = (Y(y))_{y \in U_{\varphi}, h} \) and \( Y_{y, \eta} = (Y(y))_{y \in U_{\varphi}, h} \) defined by
\[ d^2(Y_{x, \eta}, Y_{y, \eta}) = \frac{1}{m} \| Y_{x, \eta} - Y_{y, \eta} \|_2^2, \]
where \( m = \text{card} U_{\varphi}, h \), and \( |Y_{x, \eta} - Y_{y, \eta}|_2^2 = \sum_{|z| \leq h} (Y(x+z) - Y(x_0 + z))^2 \) which measures the similarity between the data patches \( Y_{r, x} \) and \( Y_{r, y} \). Our simulations show that a convenient approximation of \( \rho_{\varphi, \omega}(x) \) is given by
\[ \hat{\rho}_{\omega}(x) = \frac{d(Y_{r, x}, Y_{r, y})}{\sqrt{2\sigma^2}}. \]
\[ (17) \]
A theoretical justification for this choice is given in a convergence theorem that is not presented here.

Thus our **Optimal Weights Filter** is defined by
\[ \hat{f}(x_0) = \hat{f}_{r, \omega}(x_0) = \frac{\sum_{y \in U_{\varphi}, h} K_{ir}(\hat{\rho}_{\varphi, \omega}(y)/a)Y(x)}{\sum_{y \in U_{\varphi}, h} K_{ir}(\hat{\rho}_{\varphi, \omega}(y)/a)}, \]
\[ (18) \]
where the bandwidth \( \hat{a} > 0 \) is the solution of the equa-

---

**Algorithm 1: Optimal weights filter.**

Repeat for each \( x_0 \in I \):

1. give an initial value of \( \hat{a} \): \( \hat{a} = 1 \) (it can be an arbitrary positive number).
2. compute \( \{\hat{\rho}_{\varphi, \omega}(x) \mid x \in U_{\varphi}, h\} \) by (17)
3. compute the bandwidth \( \hat{a} \) at \( x_0 \)
4. reorder \( \{\hat{\rho}_{\varphi, \omega}(x) \mid x \in U_{\varphi}, h\} \) as increasing sequence, say
5. loop from \( k = 1 \) to \( M \)
6. if \( \sum_{i=1}^{k} \hat{\rho}_{\varphi, \omega}(x_i) > 0 \)
7. if \( \sigma^2 + \sum_{i=1}^{k} \hat{\rho}_{\varphi, \omega}(x_i) \geq \hat{\rho}(x_0) \)
8. then \( \hat{a} = \frac{\sigma^2 + \sum_{i=1}^{k} \hat{\rho}_{\varphi, \omega}(x_i)}{\sum_{i=1}^{k} \hat{\rho}_{\varphi, \omega}(x_i)} \)
9. else quit loop
10. else continue loop
11. end loop
12. compute the estimated weights \( \hat{\omega} \) at \( x_0 \)
13. compute \( \hat{\omega}(x) = \frac{K_{ir}(1 - \hat{\rho}_{\varphi, \omega}(x)/a^+)}{\sum_{y \in U_{\varphi}, h} K_{ir}(1 - \hat{\rho}_{\varphi, \omega}(y)/a^+)} \)
14. compute the filter \( \hat{f} \) at \( x_0 \)
15. compute \( \hat{f}(x_0) = \sum_{y \in U_{\varphi}, h} \hat{\omega}(x) Y(x) \).
tion \( M^*_{p,q}(\tilde{a}) = \sigma^2 \), which can be calculated as in Remark 1 with \( \rho_{f_{x_0}}(x) \) and \( a \) replaced by \( \tilde{\rho}_{x_0}(x) \) and \( \tilde{a} \) respectively. We end this section by giving an algorithm for computing the filter (18). The input values of the algorithm are the image \( Y(x), x \in I \), the standard derivation \( \sigma \) of the Gaussian noise and two numbers \( m \) and \( M \) representing the sizes of data patches and search windows respectively (cf. (3) and (4)).

To avoid the undesirable border effects in simulations, we mirror the image outside the image limits, that is, we extend the image outside the image limits symmetrically with respect to the border. At the corners, the image is extended symmetrically with respect to the corner pixels.

The implementation of the proposed algorithm is straightforward. Notice that an important issue in the Non-Local Means Filter is the choice of the bandwidth parameter in the Gaussian kernel; our algorithm has the advantage that it automatically calculates the bandwidth.

A detailed analysis of the performance of our filter is given in Section 3 where the numerical simulations show that our filter outperforms the classical Non-Local Means Filter.

![Figure 1](image1.png)

Figure 1: The evolution of PSNR value as a function of the size of data patches.

3 SIMULATIONS

In this section we show the numerical performance of the Optimal Weights Filter by simulation results.

The performance of the Optimal Weights Filter \( \hat{f}_{h,\eta}(x_0) \) is measured by the usual Peak Signal-to-Noise Ratio (PSNR) in decibels (db) defined as

\[
PSNR = 10 \log_{10} \frac{255^2}{MSE}
\]

\[
MSE = \frac{1}{|x|} \sum_{x \in I} (f(x) - \hat{f}_{h,\eta}(x))^2,
\]

where \( f \) is the original image, and \( \hat{f}_{h,\eta} \) the estimated one.

![Figure 2](image2.png)

Figure 2: Results of denoising "Lena" 512 × 512 image. Comparing (d) and (f) we see that the Optimal Weights Filter (OWF) captures more details than the Non-Local Means Filter (NLMF).

In the simulations, we sometimes use the smoothed version of the estimate of brightness variation \( d_K(Y_{x,\eta}, Y_{s_0,\eta}) \) instead of the non smoothed one \( d(Y_{x,\eta}, Y_{s_0,\eta}) \), defined by

\[
d_K(Y_{x,\eta}, Y_{s_0,\eta}) = \frac{\|K(y) \cdot (Y_{x,\eta} - Y_{s_0,\eta})\|_2}{\sqrt{\sum_{y' \in U_{s_0,\eta}} K(y')}}
\]

where \( K(y) \) are some weights defined on \( U_{s_0,\eta} \). The corresponding estimate of brightness variation \( \tilde{\rho}_{f_{x_0}}(x) \) is given by

\[
\tilde{\rho}_{K_{s_0}}(x) = \left( d_K(Y_{x,\eta}, Y_{s_0,\eta}) - \sqrt{\sigma^2} \right)^+.
\]

With the rectangular kernel

\[
K_r(y) = \begin{cases} 
1, & y \in U_{s_0,\eta}, \\
0, & \text{otherwise,}
\end{cases}
\]

115
The PSNR values show that our approach is as good as simple as the Non-Local Means Filter and, with \( K(y) = K_0(y) \), has only two parameters \( M \) and \( m \) which are the sizes of data patches and search windows. The proposed approach gives a denoising quality which is competitive with that of the recent method BM3D (Dabov et al., 2007). The behavior of the PSNR in function of the size \( m \) of data patches is displayed in Figure 1 for "Lena" image. We fix \( M = 13 \times 13 \). For \( \sigma = 20 \), Figure 1 illustrates that the PSNR value increases as \( m \) varies between 3 \( \times \) 3 and 41 \( \times \) 41 (for which PSNR = 32.71 db), and that it just changes slightly when \( m \) is sufficiently large (e.g. \( \text{PSNR} = 32.68 \text{db} \) when \( m = 27 \times 27 \)). In our experimental results (cf. Table 1) we prefer \( m = 27 \times 27 \) as the choice \( m = 41 \times 41 \) is computationally expensive.

The potential of the estimation method is illustrated with the 512 \( \times \) 512 image "Lena" (Figure 2(a)) corrupted by an additive white Gaussian noise (Figure 2(b), PSNR \( = 22.10 \text{db}, \sigma = 20 \)). We used the kernel \( K_0(y) \) for computing the estimated brightness variation function \( p_{K_0} \), which corresponds to the Optimal Weights Filter as defined in Section 2. In Figure 2(c), we can see that the noise is reduced in a natural manner and significant geometric features, fine tex-

**Table 1:** Performance of denoising algorithms when applied to test noisy (WGN) images.

<table>
<thead>
<tr>
<th>Images Sizes</th>
<th>Lena 512 ( \times ) 512</th>
<th>Barbara 512 ( \times ) 512</th>
<th>Boat 512 ( \times ) 512</th>
<th>House 256 ( \times ) 256</th>
<th>Peppers 256 ( \times ) 256</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \sigma )</td>
<td>Method</td>
<td>PSNR</td>
<td>PSNR</td>
<td>PSNR</td>
<td>PSNR</td>
</tr>
<tr>
<td>15</td>
<td>Our method ( (M = 13 \times 13, m = 27 \times 27) )</td>
<td>33.93 db</td>
<td>32.31 db</td>
<td>31.64 db</td>
<td>34.09 db</td>
</tr>
<tr>
<td></td>
<td>(Buades et al., 2005)</td>
<td>32.72 db</td>
<td>31.67 db</td>
<td>30.39 db</td>
<td>33.82 db</td>
</tr>
<tr>
<td></td>
<td>(Foi et al., 2004)</td>
<td>32.72 db</td>
<td>29.61 db</td>
<td>30.93 db</td>
<td>33.18 db</td>
</tr>
<tr>
<td></td>
<td>(Roth and Black, 2009)</td>
<td>33.29 db</td>
<td>30.16 db</td>
<td>31.27 db</td>
<td>33.55 db</td>
</tr>
<tr>
<td></td>
<td>(Hirakawa and Parks, 2006)</td>
<td>33.97 db</td>
<td>32.55 db</td>
<td>31.59 db</td>
<td>33.82 db</td>
</tr>
<tr>
<td></td>
<td>(Kervrann and Boulanger, 2008)</td>
<td>33.70 db</td>
<td>31.80 db</td>
<td>31.44 db</td>
<td>34.08 db</td>
</tr>
<tr>
<td></td>
<td>(Hammond and Simoncelli, 2008)</td>
<td>34.04 db</td>
<td>32.25 db</td>
<td>31.72 db</td>
<td>33.72 db</td>
</tr>
<tr>
<td></td>
<td>(Aharon et al., 2006)</td>
<td>33.71 db</td>
<td>32.41 db</td>
<td>31.77 db</td>
<td>34.25 db</td>
</tr>
<tr>
<td></td>
<td>(Dabov et al., 2007)</td>
<td>34.27 db</td>
<td>33.00 db</td>
<td>32.14 db</td>
<td>34.94 db</td>
</tr>
<tr>
<td>20</td>
<td>Our method ( (M = 13 \times 13, m = 27 \times 27) )</td>
<td>32.68 db</td>
<td>31.04 db</td>
<td>30.30 db</td>
<td>32.83 db</td>
</tr>
<tr>
<td></td>
<td>(Buades et al., 2005)</td>
<td>31.51 db</td>
<td>30.38 db</td>
<td>29.32 db</td>
<td>32.51 db</td>
</tr>
<tr>
<td></td>
<td>(Foi et al., 2004)</td>
<td>31.43 db</td>
<td>27.90 db</td>
<td>39.61 db</td>
<td>31.84 db</td>
</tr>
<tr>
<td></td>
<td>(Roth and Black, 2009)</td>
<td>31.89 db</td>
<td>28.28 db</td>
<td>29.86 db</td>
<td>32.29 db</td>
</tr>
<tr>
<td></td>
<td>(Hirakawa and Parks, 2006)</td>
<td>32.69 db</td>
<td>31.06 db</td>
<td>30.25 db</td>
<td>32.58 db</td>
</tr>
<tr>
<td></td>
<td>(Kervrann and Boulanger, 2008)</td>
<td>32.64 db</td>
<td>30.37 db</td>
<td>30.12 db</td>
<td>32.90 db</td>
</tr>
<tr>
<td></td>
<td>(Hammond and Simoncelli, 2008)</td>
<td>32.81 db</td>
<td>30.76 db</td>
<td>30.41 db</td>
<td>32.82 db</td>
</tr>
<tr>
<td></td>
<td>(Aharon et al., 2006)</td>
<td>32.39 db</td>
<td>30.84 db</td>
<td>30.39 db</td>
<td>33.10 db</td>
</tr>
<tr>
<td></td>
<td>(Dabov et al., 2007)</td>
<td>33.05 db</td>
<td>31.78 db</td>
<td>30.88 db</td>
<td>33.77 db</td>
</tr>
<tr>
<td>25</td>
<td>Our method ( (M = 13 \times 13, m = 27 \times 27) )</td>
<td>31.59 db</td>
<td>29.92 db</td>
<td>29.16 db</td>
<td>31.95 db</td>
</tr>
<tr>
<td></td>
<td>(Buades et al., 2005)</td>
<td>30.36 db</td>
<td>29.19 db</td>
<td>28.38 db</td>
<td>31.66 db</td>
</tr>
<tr>
<td></td>
<td>(Foi et al., 2004)</td>
<td>30.43 db</td>
<td>26.62 db</td>
<td>28.60 db</td>
<td>30.75 db</td>
</tr>
<tr>
<td></td>
<td>(Roth and Black, 2009)</td>
<td>30.57 db</td>
<td>26.84 db</td>
<td>28.57 db</td>
<td>31.05 db</td>
</tr>
<tr>
<td></td>
<td>(Hirakawa and Parks, 2006)</td>
<td>31.69 db</td>
<td>29.89 db</td>
<td>29.21 db</td>
<td>31.60 db</td>
</tr>
<tr>
<td></td>
<td>(Kervrann and Boulanger, 2008)</td>
<td>31.73 db</td>
<td>29.24 db</td>
<td>29.20 db</td>
<td>32.22 db</td>
</tr>
<tr>
<td></td>
<td>(Hammond and Simoncelli, 2008)</td>
<td>31.83 db</td>
<td>29.58 db</td>
<td>29.40 db</td>
<td>31.54 db</td>
</tr>
<tr>
<td></td>
<td>(Aharon et al., 2006)</td>
<td>31.36 db</td>
<td>29.58 db</td>
<td>29.32 db</td>
<td>32.07 db</td>
</tr>
<tr>
<td></td>
<td>(Dabov et al., 2007)</td>
<td>32.08 db</td>
<td>30.72 db</td>
<td>29.91 db</td>
<td>32.86 db</td>
</tr>
</tbody>
</table>
ures, and original contrasts are visually well recovered with no undesirable artifacts (PSNR = 32.68 db for “Lena”). To better appreciate the accuracy of the restoration process, the square of the difference between the original image and the recovered image is shown in Figure 2(d), where the dark values correspond to a high-confidence estimate. As expected, pixels with a low level of confidence are located in the neighborhood of image discontinuities. For comparison, we show the image denoised by Non-Local Means Filter in Figures 2(e),(f). The overall visual impression and the numerical results are improved using our algorithm.

The Optimal Weights Filter seems to provide a feasible and rational method to detect automatically the details of images and take the proper weights for every possible geometric configuration of the image. The distribution of the weights inside the search window $U_{x,h}$ depends on the estimated brightness variation function $\tilde{p}_{K,h}(x), x \in U_{h, h}$. If the estimated brightness variation $\tilde{p}_{K,h}(x)$ is less than $\tilde{a}$ (see Theorem 1), the similarity between patches is measured by a linear decreasing function of $\tilde{p}_{K,h}(x)$, otherwise it is zero. Thus $\tilde{a}$ acts as an automatic threshold.

4 CONCLUSIONS

We have proposed a new filter to remove Gaussian noise, based on optimization of weights in the weighted means approach. Our analysis shows that a triangular kernel is preferred rather than the Gaussian kernel. The proposed filter improves the usual Non-Local Means Filter both numerically and visually in denoising performance; it also has the advantage to be adaptive in the sense that it calculates automatically the good bandwidth of the triangular kernel (while in the Non-Local Means Filter the choice of the bandwidth parameter in the Gaussian kernel is delicate). We hope that the optimal weights that we deduced can also bring similar improvements for recently developed algorithms where the basic idea of the Non-Local means filter is used.

ACKNOWLEDGEMENTS

The authors are grateful to the reviewers for helpful comments and remarks. The work has been partially supported by the National Natural Science Foundation of China, Grant No. 11101039 and Grant No. 11171044.

REFERENCES


